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Business continuity of energy systems: a quantitative framework for dynamic assessment and optimization

Jinduo Xing

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Business continuity of energy systems: a quantitative framework for dynamic assessment and optimization

Thèse de doctorat de l'Université Paris-Saclay
préparée à Centralesupélec

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Mme Jinduo XING

Composition du Jury:

Anne Barros Professeur, CentraleSupélec, Université Paris-Saclay	Président
Francesco Di Maio Professeur associé, Politecnico di Milano	Rapporteur
Emmanuel Remy Chercheur, EDF R&D	Examineur
Sébastien Travadel Professeur associé, Mines ParisTech	Rapporteur
Hong Xia Professeur, Harbin Engineering University	Examinatrice
Enrico Zio Professeur, Politecnico di Milano	Directeur de thèse

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Abstract

Concerns over disruptive events on the operation of energy systems have increased significantly during the past decades. This creates considerable demands on accurate business continuity assessment and effective management techniques for these systems. New opportunities for this come from using online-collected data and information to assess risk (dynamic risk assessment) and business continuity (dynamic business continuity assessment), and from using the assessment results for improving the optimal design of the system to achieve maximal business continuity.

With this perspective in the present thesis, first, a dynamic risk assessment (DRA) framework is developed to capture the time-dependent degradation behaviour of safety barriers by integrating both condition monitoring data and inspection data. Condition monitoring data are online-collected by sensors and assumed to indirectly relate to component degradation; inspection data are recorded in physical inspections that are assumed to directly measure the component degradation. A Hidden Markov Gaussian Mixture Model (HM-GMM) is developed for modelling the condition monitoring data and a Bayesian network (BN) is developed to integrate the two data sources for DRA. Risk updating and prediction are exemplified on an Event Tree (ET) risk assessment model. A numerical case study and a real-world application on a Nuclear power plant (NPP) are performed to demonstrate the application of the proposed work.

Then, a dynamic business continuity assessment (DBCA) framework is proposed to capture time-dependent behaviours and integrate the information on the conditions of components and system in the business continuity assessment (BCA). Specifically, a particle filtering (PF)-based method is developed to integrate condition monitoring data on the safety barriers installed for system protection and predict their reliability as their health states change due to ageing. An instalment model and a stochastic price model are also employed to quantify the time-dependent revenues and tolerable losses during the operation of the system. A simulation model is developed to evaluate dynamic business continuity metrics originally introduced. A case study regarding a NPP risk scenario is worked out to demonstrate the applicability of the proposed approach.

Finally, a joint optimization model is developed to optimally design safety barriers of different natures, including prevention, mitigation, emergency and recovery barriers to enhance the business continuity of the system. The joint

optimization is guided by a business continuity metrics called expected business continuity values (EBCV). A physics-of-failure model is developed to model the effectiveness of prevention safety barriers. An ET model is developed to describe the potential accident evolution process. A redundancy allocation model is, then, used to consider the efforts to enhance the mitigation and emergency barriers. Recovery measures are also considered by a widely used logarithmic function model. A mixed-integer genetic algorithm is employed to obtain optimal solutions of the joint optimisation model. The developed framework is applied on a case study of steam generator tube rupture accident in a NPP.

Overall, through the research of this thesis, we have established a framework that allows making BCA using online-collected information. We have also showed how to optimize the business continuity of a system through a joint optimization model. These findings demonstrate the prospects of applying BCM in accident prevention, mitigation, emergency, recovery, to better support the operation of energy systems by ensuring its business continuity.

Keywords: Dynamic risk assessment, Dynamic business continuity assessment, Condition monitoring data, Inspection data, Event tree, Hidden Markov-Gaussian Mixture model, Particle filtering, Stochastic electricity model, Joint optimization, Mixed integer genetic algorithm

Résumé

Les inquiétudes suscitées par des événements perturbateurs sur le fonctionnement des systèmes énergétiques ont considérablement augmenté au cours des dernières décennies. Cela crée des exigences considérables en matière d'évaluation de la continuité des opérations et de techniques de gestion efficaces pour ces systèmes. Les nouvelles opportunités à cet égard proviennent de l'utilisation des données et des informations collectées en ligne pour évaluer les risques (évaluation dynamique des risques) et la continuité de l'activité (évaluation dynamique de la continuité des activités), ainsi que de l'utilisation des résultats de l'évaluation pour améliorer la conception optimale du système et atteindre une continuité maximale des activités.

Dans cette perspective dans la présente thèse, un cadre d'évaluation dynamique des risques (DRA) est développé pour capturer le comportement de dégradation dépendant du temps des barrières de sécurité en intégrant à la fois des données de surveillance des conditions et des données d'inspection. Les données de surveillance des conditions sont collectées en ligne par des capteurs et supposées être indirectement liées à la dégradation des composants; les données d'inspection sont enregistrées lors d'inspections physiques censées mesurer directement la dégradation du composant. Un modèle de mélange gaussien caché de Markov (HM-GMM) est développé pour modéliser les données de surveillance de l'état et un réseau bayésien (BN) est développé pour intégrer les deux sources de données pour la DRA. La mise à jour et la prévision des risques sont illustrées dans un modèle d'évaluation des risques de l'arbre des événements. Une étude de cas numérique et une application réelle sur une centrale nucléaire (centrale nucléaire) sont réalisées pour démontrer l'application du travail proposé.

Ensuite, un cadre d'évaluation dynamique de la continuité des opérations (DBCA) est proposé pour capturer les comportements dépendant du temps et intégrer l'information sur les conditions des composants et du système dans l'évaluation de la continuité des opérations (BCA). Plus précisément, une méthode basée sur le filtrage de particules (PF) est développée pour intégrer les données de surveillance des conditions sur les barrières de sécurité installées pour la protection des systèmes et prévoir leur fiabilité lorsque leur état de santé évolue en raison du vieillissement. Un modèle de versement et un modèle de prix stochastique sont également utilisés pour quantifier les revenus et les pertes tolérables en fonction du temps pendant le fonctionnement du système. Un modèle de simulation est développé

pour évaluer les mesures de continuité d'activité dynamiques introduites à l'origine. Une étude de cas concernant un scénario de risque de centrale nucléaire est élaborée pour démontrer l'applicabilité de l'approche proposée.

Enfin, un modèle d'optimisation commun est élaboré pour concevoir de manière optimale des barrières de sécurité de différentes natures, notamment des barrières de prévention, d'atténuation, d'urgence et de reprise, afin d'améliorer la continuité des opérations du système. L'optimisation conjointe est guidée par une métrique de continuité d'activité appelée valeurs de continuité d'activité attendues (EBCV). Un modèle de physique de défaillance est développé pour modéliser l'efficacité des barrières de sécurité préventives. Un modèle ET est développé pour décrire le processus d'évolution des accidents potentiels. Un modèle d'allocation de redondance est donc utilisé pour prendre en compte les efforts visant à renforcer les barrières d'atténuation et d'urgence. Les mesures de récupération sont également prises en compte par un modèle de fonction logarithmique largement utilisé. Un algorithme génétique à nombres entiers mixtes est utilisé pour obtenir des solutions optimales du modèle d'optimisation conjointe. Le cadre développé est appliqué à une étude de cas d'accident de rupture de tube de générateur de vapeur dans une centrale nucléaire.

Globalement, à travers la recherche de cette thèse, nous avons établi un cadre qui permet de créer une BCA en utilisant des informations collectées en ligne. Nous avons également montré comment optimiser la continuité d'activité d'un système grâce à un modèle d'optimisation commun. Ces résultats démontrent les perspectives d'application de la BCM dans la prévention, l'atténuation, les urgences et la récupération des accidents, afin de mieux soutenir le fonctionnement des systèmes énergétiques en assurant la continuité de ses activités.

Mots-clés: Evaluation dynamique des risques, Evaluation dynamique de la continuité des opérations, Données de surveillance des conditions, Données de contrôle, Arbre des événements, Modèle de mélange caché markov-gaussien, Filtrage de particules, Modèle d'électricité stochastique, Optimisation d'articulation, Algorithme génétique d'entiers mixtes

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Acronyms

ATWS	Anticipated Transient without scram
BCA	Business continuity assessment
BCM	Business continuity management
BCV	Business continuity value
BN	Bayesian Network
DRA	Dynamic risk assessment
DBCA	Dynamic business continuity assessment
DBN	Dynamic Bayesian Network
EM	Expectation Maximization
EBCV	Expected business continuity value
ET	Event tree
FT	Fault tree
MIGA	Mixed-integer genetic algorithm
HM-GMM	Hidden Markov-Gaussian Mixture Model
IE	Initiating Event
NPP	Nuclear power plant
PF	Particle Filtering
PRA	Probabilistic risk assessment
PDF	Probability density function
QRA	Quantitative risk assessment
RA	Risk assessment
RUL	Remaining useful life
RCS	Reactor coolant system
RDS	Reactor depressurization system

RWST	Refuelling water storage tank
SGTR	Steam generator tube rupture

Notation

a	Crack length
π	Initial state distribution of the Markov degradation process
$b_i(\mathbf{x})$	Probability distribution of the degradation indicator \mathbf{x} when the degradation state is S_i
C_i	The i -th consequence in the ET
$c_i(t_k)$	Condition monitoring data from the i -th safety barrier at $t = t_k$
$\mathbf{c}_{Tr}^{(k)}(t)$	Condition monitoring data from the k -th training sample at t
$d(\cdot)$	Euclidean distance
$f_{ET}(\cdot)$	ET model
L_{in}	Indirect loss
L_d	Direct loss
K	Number of safety barriers with time-dependent failure probabilities
M	Number of safety barriers in a system
N	Number of consequences in the ET
N_S	Sample size of PF
$n_{feature}$	Number of features extracted from condition monitoring data
n_{Tr}	Number of samples in the training data set
P_{C_i}	Probability that consequence i occurs, given that the IE has occurred
$P_{CM,t_k}(S_{CM})$	Posterior distribution of the estimated degradation state from condition monitoring data, evaluated at t_k
$P_{INT,t_k}(S)$	Posterior distribution of the estimated degradation state by integrating condition monitoring data and inspection data, evaluated at t_k

Q	Number of health states
R_{IN}	Reliability of the inspection
R_{SB_M}	Reliability of the M – th safety barrier
S_{CM}	Estimated degradation state from condition monitoring data
$S_{CM,MAP}$	Most likely degradation state given the condition monitoring data
S_{IN}	Estimated degradation state from inspection data
S	True degradation state
t_{Tr}	Length of the observation period for the training samples
t_{recv}	Recovery time
W	Working set that contains all the working states
$\mathbf{x}_{Tr}^{(k)}(t)$	Health indicator of k - th training data at t
$\mathbf{x}(t)$	Health indicator of safety barrier at t
$\boldsymbol{\mu}$	Vector of the mean values of the multivariate Gaussian distribution
$\boldsymbol{\Sigma}$	Covariance matrices of the multivariate Gaussian distribution
$\alpha_i(S_i)$	Forward variable
$\beta_i(S_i)$	Backward variable

Appended papers

Paper I: J. Xing, Z. Zeng, E. Zio. A framework for dynamic risk assessment with condition monitoring data and inspection data. *Reliability Engineering and System Safety*, 2019, 191, 106552.

Paper II: J. Xing, Z. Zeng, E. Zio. Dynamic business continuity assessment using condition monitoring data. *International Journal of Disaster Risk Reduction*, 2019, 41, 101334.

Paper III: J. Xing, Z. Zeng, E. Zio. Joint optimization of safety barriers against steam generator tube rupture to enhance business continuity of nuclear power plants. *Reliability Engineering and System Safety*, 2019. (Under review).

Chapter 1 Introduction

Business operations of energy systems, such as nuclear power plants (NPPs), electricity transmission systems, are threatened by a number of hazards [1-4]. These should be properly managed [5]. Conventionally, risk assessment and management are employed to protect the business from disruptive events. In risk assessment, possible consequences and associated likelihoods are considered for accidents potentially developing from the identical hazards [6]. On the other hand, the process of recovering from an accident has a significant influence on business operations, as it directly affects downtime. Recently, a holistic method known as business continuity management (BCM) has been put forth, which integrates protection, mitigation, emergency and recovery to ensure the continuous operation of a business.

Many questions and challenges arise in the application of BCM to energy systems. For instance, as sensor technologies and computing resources advance, data and information can be collected online, as the system operates. How to use these data and information to support proactive and real-time quantitative risk assessment (QRA) and business continuity assessment (BCA) is an opportunity, and, at the same time, a challenging issue in BCM. Another challenging issue is how to optimize business continuity, considering the components and safety barriers of different nature that make up the system. Objective of this thesis is to address the aforementioned questions by providing a quantitative framework for the safe and continuous business operation of energy systems. The focus is on the quantitative assessment of business continuity for energy systems under disruptive events (e.g., steam generator tube rupture (SGTR) [7], anticipated transient without scram (ATWS) accidents [8]). An integrated framework for BCA is proposed first, where four stages named preventive stage, mitigation stage, emergency stage and recovery stage are comprehensively considered and integrated. Due to the vital role of risk assessment in BCM, a dynamic risk assessment (DRA) framework is proposed, capable of incorporating both inspection data and condition monitoring data. Finally, the optimization of business continuity is considered by developing a joint optimization model.

In the following of this chapter, we present a brief introduction of the context of the research and open issues in Section 1.1 and Section 1.2, respectively. The research objective and main contributions are discussed in Section 1.3. Finally, Section 1.4 shows the structure of the thesis.

1.1 Business continuity management

BCM is defined by the international organization of standards (ISO) as “the holistic management process that identifies potential threats to an organization and the impacts to business operations those threats, if realized might cause, and which provides a framework for building organizational resilience with the capability of an effective response that safeguards the interest of its key stakeholders, reputation, brand and value-creating activities”[9]. In a nutshell, BCM is a comprehensive method that integrates pre-event and post-event management together to ensure the resilience and continuous operation of system business. Compared to conventional risk analysis method, BCM not only focuses on the potential hazards and their impacts, but also considers how to mitigate the consequence and quickly recover from the disruption.

BCM aims at developing appropriate methods in order to prevent and resume system business to an acceptable predefined level [10]. Usually, pre-disruptive and post-disruptive measures are considered in a system with respect to system resilience and business continuity [11]. The former aims at identifying potential hazards and reducing their possibility. The latter is associated with resuming system business in the aftermath of disruptive event to reduce potential losses [12].

A conceptual model is presented in Figure 1-1 to illustrate the different processes involved in BCM. Business continuity measures the ability of an organization to resist, mitigate and recover to an acceptable state given a disruptive event.

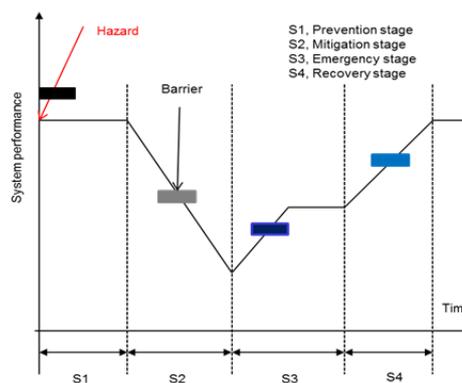


Figure 1-1. A conceptual scheme of the business continuity process [13].

For the pre-event stage, protection measures are installed in advance to resist to the potential event, i.e. by reducing the probability of occurrence of the accident event. Next comes the mitigation phase, where safety barriers are usually activated to mitigate the consequences of the disruptive event once occurred. The purpose of the mitigation

phase is to contain the evolution of the accident consequences [14, 15]. Emergency measures act to cope with the accident evolution and human intervention is often required [16]. Finally, recovery actions are taken to bring system business back to operation.

Most existing researches on BCM are, however, based on qualitative BCA [17-22]. This situation impedes the quantitative analysis of business continuity and its effective application. Thus, a quantitative framework for the assessment and optimization of the business continuity process is needed.

1.2 Open issues

Risk assessment plays a fundamental role in BCM. How to improve risk assessment accuracy with the help of different knowledge, information and data is the first research topic in this thesis. In Section 1.2.1, we review the related works on this topic. In Section 1.2.2, we review the researches related to a quantitative BCA considering available time-variant factors. Section 1.2.3 reviews the related research efforts on optimization models for enhancing business continuity.

1.2.1 Dynamic risk assessment

Traditional risk assessment methods, like event tree (ET) and fault tree (FT), mainly treat the failure probabilities of safety barriers as constant values, without explicitly modelling degradation and aging processes [23]. In practice, operational and environmental conditions of the system change with time, and this generally causes time-dependent behaviours of the safety barriers [24-26]. To account for the time-dependent characteristic of safety barriers, a number of DRA frameworks have been developed, which employ data and information collected during the system operation to update the estimated risk indexes [27]. The goal of DRA is to obtain an estimate of system's risk updated in real time with the accumulated information and data [28]. Bayesian theory has been used to update the probabilities of the events in an ET [29, 30]. A condition-based risk assessment has been performed in [24] for a spontaneous SGTR accident. A data-driven DRA model has been developed for offshore drilling operations, where real-time operational data have been employed to update the probability of kick events [31]. In [32], statistical failure data and condition monitoring data have been integrated in a hierarchical Bayesian model for DRA.

The data used by existing DRA methods can be broadly grouped into two categories: statistical data and condition monitoring data. Statistical failure data refer to counts of accidents, incidents or near misses collected in the field. Condition monitoring data are the online monitoring data collected by sensors that are installed in the system for monitoring the degradation process of the safety barriers. Uncertainty may exist in condition monitoring data due to possible noise during the monitoring process. Apart from these two data types, inspection data can also be collected by physical inspections performed by maintenance personnel [33], and might serve as another data source for online reliability assessment. In [34], a Bayesian method has been developed to merge experts' judgments with continuous and discontinuous inspection data for the reliability assessment of multi-state systems. A two-stage recursive Bayesian approach has been developed in [35], in order to update system reliability based on imperfect inspection data. Condition monitoring data and inspection data on wind turbine blades have been used separately for remaining useful life estimation in [36]. Inspection data directly measure the component degradation and provide valuable information complementary to condition monitoring data for DRA. In this thesis, we aim at developing new DRA methods that allow integrating condition monitoring data with inspection data, for real-time risk estimates update.

1.2.2 Dynamic business continuity assessment

Most of the existing methods for quantitative BCA focus on time-static problems [37], where the analysis is done before operation and is not updated to consider aging and degradation of components and systems. For instance, a statistical model integrating Cox's model and Bayesian networks has been proposed to model the BCM process [38]. In [12], the BCM outsourcing and insuring strategies have been compared based on the organization characteristics and the relevant data through a two-step fuzzy cost-benefit analysis. Two probabilistic programming models have been developed in [39] to determine appropriate business continuity plans given epistemic uncertainty in the input data. In [40], a new model for integrated business continuity and disaster recovery planning has been presented, considering multiple disruptive incidents that might occur simultaneously. An integrated framework has been developed for quantitative business continuity analysis, where four numerical metrics were proposed to quantify the business continuity level based on the potential loss caused by the disruptive event [14].

However, in practice, various time-dependent factors might affect the business continuity, e.g., the degradation of safety barriers, the dynamic behaviour of profits and losses. On the other hand, as sensor technologies and computing resources advance, it is possible to capture these dynamic factors even in real-time, based on online-

collected condition monitoring data [41, 42]. For example, a condition-based fault tree has been used for dynamic risk assessment (DRA) [43], where condition monitoring data are used to update the failure rates of specific components and predict the reliability. In [44], a Bayesian reliability updating method has been proposed for dependent components by using condition monitoring data. Therefore, in this thesis, we investigate how to use the online collected data and information to support dynamic business continuity assessment (DBCA), with time-dependent contributing factors.

1.2.3 Joint optimization

In general, the resources an organization can invest to safeguard the continuous operation of a system are limited. How to allocate and arrange the limited resources among the prevention, mitigation, emergency and recovery measures is an important topic to address. Some studies have developed methods to allocate resources to improve system resilience for a specific disaster. For instance, multi-systems' joint restoration processes modeling has been addressed and the effectiveness of five different restoration strategies has been compared in [45] regarding hurricane hazard. In [46], a two-stage mixed-integer programming resource allocation model for lifeline systems has been proposed to improve the efficiency of restoration. A multi-objective optimization model of emergency organization allocation for sustainable disaster supply chains has been developed to design optimized strategies of emergency organization allocation [47], with the objective of minimizing the expected outage duration of loads. A scenario-based two-states stochastic optimization for minimizing outage duration in distribution damage and road network damage has been exploited in [48]. In [49], a restoration resource allocation model has been proposed to enhance resilience of interdependent infrastructure systems. A resilience-based optimization methodology has been performed over the set of feasible restoration policies, information investments and human resource availability to determine optimal customer and system-wide monetary utility [50]. A stochastic optimization technique has been developed to allocate scarce national resources to cope with multiple simultaneous disasters occurring across the nation [51].

Most existing research, as reviewed above, considers the safety barriers separately. In this work, we aim to develop a joint optimization model that aims to assure an holistic optimal performance, considering all the safety barriers.

1.3 Research objectives and contributions

The focus of this thesis is to develop methods that support DBCA, based on the online-collected data and information. Besides, we also aim to develop a joint optimization model for maximizing system business continuity, through optimally allocating resources among prevention, mitigation, emergency and recovery measures.

The main contributions of the thesis can be summarized as follows:

- (1) A new DRA framework is developed, which allows integrating condition monitoring data and inspection data for online assessment;
- (2) An integrated DBCA model is proposed, which allows updating the business continuity in real time, using the online-collected data and information;
- (3) A joint optimization is developed to optimize the business continuity considering the prevention, mitigation, emergency and recovery phases.

1.4 Structure of the thesis

This thesis includes two parts. The first part contains five chapters, introducing the research context and describing the problems addressed, approaches proposed, and related results.

Chapter 2 begins with a state of art on DRA and continues with the roles of condition monitoring data and inspection data for risk and reliability analysis. A HM-GMM is developed for modelling the condition monitoring data and a Bayesian network (BN) is proposed to integrate the two data sources for DRA. A real-world application on a NPP [52] is conducted to demonstrate the use of the proposed framework.

Chapter 3 firstly reviews researches related to BCA, which are grouped into qualitative methods and quantitative methods. To capture the time-variant factors in BCA, a particle filtering (PF)-based method is developed to predict the reliability of the safety barriers in time. Moreover, an instalment model and a stochastic price model are also employed to model the time-dependent revenues and tolerable losses of the organization. Finally, a case study on a NPP is performed to demonstrate the applicability of the proposed approach.

Chapter 4 focuses on the joint optimization of business continuity. An optimization model is developed for resource allocation on system safety barriers to enhance business continuity, considering all the phases from pre-

disruption protection to post-disruption response and recovery. The optimal solution is obtained by a Mix-integer genetic algorithm (MIGA), which aims at maximizing system business continuity over a finite time horizon. To investigate the utility of the optimization model, a case study on a nuclear power plant (NPP) is performed to maximize expected business continuity value (EBCV) against threat of SGTR.

Chapter 5 draws conclusions of the thesis and points out the potential future works.

The second part contains a collection of three papers, describing the research work performed during the PhD, where readers can refer to for further technical details. In paper I, condition monitoring data and, inspection data are integrated to conduct DRA (corresponding to Chapter 2). In paper II, a dynamic BCA is proposed employing PF and the instalment model (corresponding to Chapter 3). In paper III, a joint optimization of the resources on safety barriers for enhancing system business continuity is proposed (corresponding to Chapter 4).

Chapter 2 Dynamic risk assessment using condition monitoring data and inspection data

The aim of this chapter is to present a simulation-based framework for DRA using condition monitoring data and inspection data. This chapter focuses on describing the condition monitoring data and inspection data influence on the system real-time risk index (here, the probability of different consequences). A model for integrating condition monitoring data and inspection data is proposed to update the safety barriers failure probabilities. The updated values are employed in a target ET to obtain the updated risk index.

Section 2.1 briefly reviews related works. Section 2.2 concretely describes the problem addressed. Section 2.3 provides a HM-GMM for reliability updating and prediction of the failure probability of safety barriers, based on condition monitoring data. A Bayesian network model is developed to integrate condition monitoring data and inspection data in Section 2.4. In Section 2.5, the developed method is used for the DRA of a real-world NPP. Finally, conclusions are discussed in Section 2.6.

2.1 State of the art

Dynamic risk assessment (DRA) attempts to use available data and new information collected during the system life to update the estimated risk index [27, 53], which may reshape the risk management framework. Many efforts on DRA have been conducted. For instance, in [54-57], near miss and incident data have been used to estimate the dynamic failure probability of accident. The basic theory under DRA using statistical data (near miss and incident data) is that using all available information and new data in the form of likelihood function, by means of Bayesian theorem. Afterwards, the updated probabilities are used in the re-estimation of risk index at the current moment [27, 28, 58, 59]. Due to possible component degradation, e.g. wear [60], fatigue [61], and crack growth [23], the failure of these component can lead to accident. Additionally, the degradation can be monitored by modern sensor technology. Therefore, condition monitoring data become the other type of data that has been emerging for DRA recent years, which refer to the online monitoring data and can capture the system real-time degradation state [23, 62]. For example, a condition-based fault tree has been used for DRA, where the condition monitoring data have

been used to update the failure rates of the specific components and predict the reliability [43, 63]. Particle filtering (PF) has been used for DRA based on condition monitoring data from a nonlinear non-Gaussian process [64]. In [23], condition monitoring data from a passive safety system have been used for DRA, without considering the uncertainty in the condition monitoring data.

2.2 Problem definition

In this chapter, we consider the DRA by integrating two data sources, i.e., condition monitoring data and inspection data. Condition monitoring data refer to the online monitoring data collected by sensors that are installed in the target system for monitoring the degradation process of the safety barrier [65]. Inspection data are collected by physical inspections performed by maintenance personnel. More specifically, the problem is formulated below.

Without loss of generality, we consider a generic Event Tree (ET) model for DRA, but the framework is applicable to other risk assessment models as well. Let IE represent the initiating event of the ET and assume that there are M safety barriers (SB) in the ET, denoted by $SB_i, i = 1, 2, \dots, M$, whose states can be working or failure. The sequences that emerge from the IE depend on the states of the SB s and lead to N possible consequences, denoted by C_1, C_2, \dots, C_N . The generic risk index considered in this chapter is the conditional probability that a specific consequence C_i occurs, given that the IE has occurred:

$$P_{C_i} = P\{C_i \text{ occurs} | IE \text{ has occurred}\}, i = 1, 2, \dots, N. \quad (2.1)$$

Conditioning on the occurrence of the IE , these probabilities are functions of the reliabilities $R_{SB_i}, i = 1, 2, \dots, M$ of the safety barriers along the specific sequences:

$$P_{C_i} = f_{ET}(R_{SB_1}, R_{SB_2}, \dots, R_{SB_M}), i = 1, 2, \dots, N. \quad (2.2)$$

where $f_{ET}(\cdot)$ is the ET model function. For example, in the ET in Figure 2-1, the risk index P_{C_2} of the consequence C_2 of the second accident sequence, in which the IE occurs with certainty, the first SB_1 functions successfully and the second SB_2 fails to provide its function, can be calculated as:

$$\begin{aligned} P_{C_2} &= f_{ET}(R_{SB_1}, R_{SB_2}) \\ &= R_{SB_1}(1 - R_{SB_2}). \end{aligned} \quad (2.3)$$

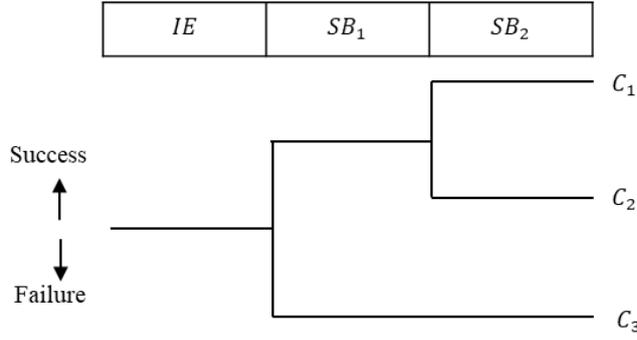


Figure 2-1. Illustrative Event Tree model.

Without loss of generality, we assume that in the ET:

- (1) Safety barriers SB_1, SB_2, \dots, SB_K are subject to degradation processes and, therefore, their reliability functions are time-dependent, whereas $SB_{K+1}, SB_{K+2}, \dots, SB_M$ do not degrade and have constant reliability values;
- (2) Condition monitoring data are collected for SB_1, SB_2, \dots, SB_K at predefined time instants $t = t_k, k = 1, 2, \dots, q$;
- (3) The collected condition monitoring data on the i -th safety barrier at $t = t_k$ are denoted by $c_i(t_k)$, where $i = 1, 2, \dots, K, k = 1, 2, \dots, q$ and $\mathbf{c}_i(t) = [c_i(t_1), c_i(t_2), \dots, c_i(t_q)]$ is a vector containing all the signals that are monitored, where q is the length of the time series;
- (4) At $t = t_m$, inspections are performed on the safety barriers $SB_i, i = 1, 2, \dots, K$. The inspection data are denoted by $S_{IN,i}, i = 1, 2, \dots, K$.

2.3 A HM-GMM for modelling condition monitoring data

In this section, we develop a HM-GMM to model condition monitoring data. In section 2.3.1, we formally define the HM-GMM. Then, in section 2.3.2, we show how to use the developed HM-GMM to estimate the degradation state of a safety barrier using condition monitoring data. The estimated degradation states are, then, used in section 2.4 for data integration in DRA.

2.3.1 Model formulation

Without loss of generality, we illustrate the HM-GMM using the i -th safety barrier in the ET. For simplicity of presentation, we drop the subscript i in the notations. An illustration of the model is given in Figure 2-2. It is assumed that the safety barrier degrades during its lifetime and the degradation process follows a discrete state discrete time Markov model $S(t)$ with a finite state space $S(t) \in \{S_1, S_2, \dots, S_Q\}$, where $S(t)$ represents the health state of the safety barrier, Q is the number of health states, and S_1, S_2, \dots, S_Q are in descending order of health (S_1 is the perfect functioning state, S_Q is the failure state). The evolution of the degradation process is characterized by the transition probability matrix of the Markov process, denoted by A , where $A = \{a_{ij}\}$ and $a_{ij} = P(S(t_{k+1}) = S_j | S(t_k) = S_i), k = 1, 2, \dots, q, 1 \leq i, j \leq Q$. The initial state distribution of the Markov process is denoted by $\pi = [\pi_1 \ \pi_2 \ \dots \ \pi_Q]$, where $\pi_i = P(S(t_0) = S_i), 1 \leq i \leq Q$. It should be noted that repairs are not considered in this chapter. Therefore, $S(t)$ can only transit to a worse state and cannot move backwards. Besides, the failure state S_Q is an absorbing state, such that $p(S(t_{k+1}) = i | S(t_k) = S_Q) = 1$ if and only if $i = S_Q$ and $p(S(t_{k+1}) = i | S(t_k) = S_Q) = 0$ for other values of i .

The discrete time discrete state Markov process model is chosen because it is widely applied for quantitatively describing discrete state degradation processes in many practical applications [66]. For example, a discrete state Markov model has been used to model the bearing degradation process in [67]. The degradation process of a safety instrumented system is modelled by a Markov model for availability analysis [68, 69]. Although only Markov process-based degradation models are discussed in this chapter, the developed methods for data integration into DRA can be easily extended to other degradation models.

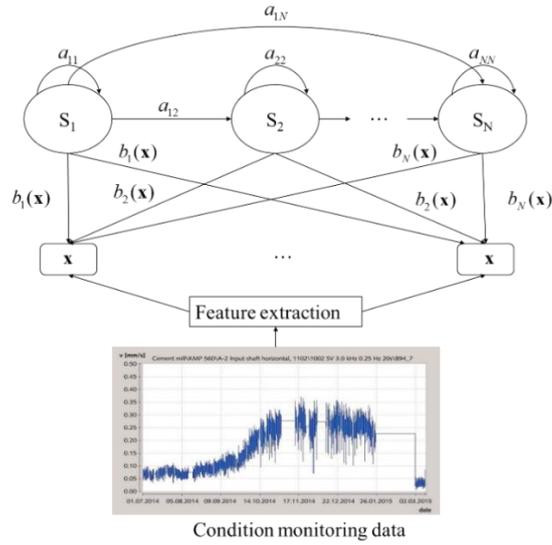


Figure 2-2. Description of the HM-GMM.

2.3.2 Degradation states estimation based on condition monitoring data

In this section, we show how to estimate the degradation states of the safety barriers based on the developed HM-GMM of the condition monitoring data. As shown in Figure 2-3, the estimation is made by an offline step and an online step. In the offline step, a HM-GMM is trained based on training data from a population of similar systems. The trained HM-GMM model, is, then, used in the online step for degradation state estimation based on the condition monitoring data.

The offline step starts from collecting training data, denoted by $\mathbf{c}_{T_r}^{(k)}(t), k = 1, 2, \dots, n_{T_r}, t = t_1, t_2, \dots, t_{T_r}$. The training data comprise of historical measurements of the degradation signals from a population of similar systems. To ensure the accuracy of HM-GMM training, it is required to collect as many as possible training samples, i.e., the sample size n_{T_r} should be as large as possible. The raw training data are preprocessed in a feature extraction step, as shown in Figure 2-3, to extract the health indicators $\mathbf{x}_{T_r}^{(k)}(t), k = 1, 2, \dots, n_{T_r}, t = t_1, t_2, \dots, t_{T_r}$. Depending on the nature of the degradation process condition, different feature extraction methods, e.g., time-domain, frequency domain, time-frequency analyses, etc., can be used [70]. Next, in the HM-GMM training step, the extracted degradation indicators are used to estimate the parameters $\lambda = \{\pi, A, \mu, \Sigma\}$ of the trained HM-GMM. In this chapter, the Expectation Maximization (EM) algorithm [71] is employed for training the HM-GMM (see section 2.3.2.1 for details). The parameters λ is the output of the offline step.

The online step starts from collecting the condition monitoring data for the safety barrier, denoted by $\mathbf{c}(t_k), k=1, 2, \dots, q$. The condition monitoring data should be of the same type and collected by the same sensors, as in the offline step. Then, the raw degradation signals are preprocessed and the health indicators $\mathbf{x}(t_k), k=1, 2, \dots, q$ of the target safety barrier are extracted, following the same procedures as in the offline step. Next, the degradation state of the safety barrier is estimated, based on the HM-GMM trained in the offline step. In this chapter, we use the forward algorithm for degradation state estimation [71], as presented in details in section 2.3.2.2. The estimated degradation state based on only condition monitoring data, denoted by $S_{CM}(t_k)$, is, then, integrated with inspection data for DRA in Section 2.4.

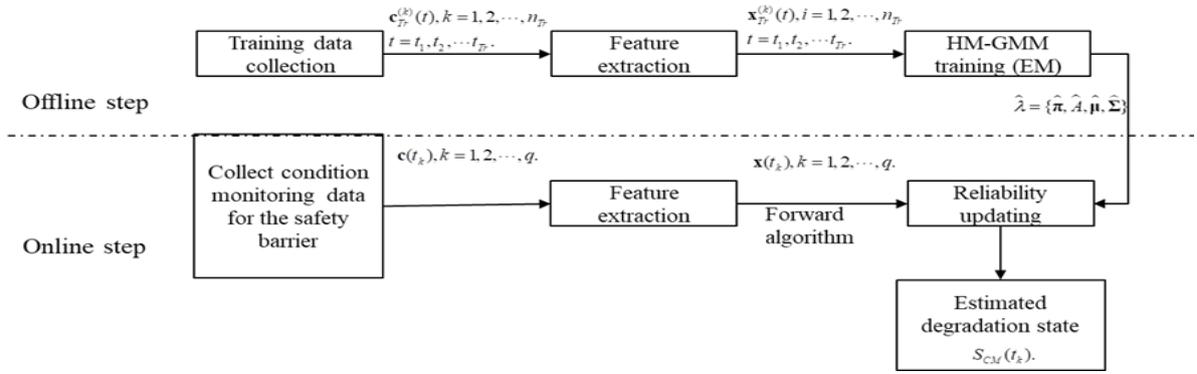


Figure 2-3. Degradation state estimation based on condition monitoring data.

2.3.2.1 HM-GM training

In this section, we present in detail how to do HM-GMM training in the offline step. The parameters

$\lambda = \{\pi, A, \mu, \Sigma\}$ are estimated by maximizing the likelihood of observing the $\mathbf{x}_{Tr}^{(k)}(t), k=1, 2, \dots, n_{Tr}, t = t_1, t_2, \dots, t_{Tr}$:

$$\begin{aligned} \lambda &= \arg \max_{\lambda} P(\mathbf{x}_{Tr}^{(1)}(t), \mathbf{x}_{Tr}^{(2)}(t), \dots, \mathbf{x}_{Tr}^{(n_{Tr})}(t) | \lambda) \\ &= \arg \max_{\lambda} \prod_{k=1}^{n_{Tr}} P(\mathbf{x}_{Tr}^{(k)}(t) | \lambda) \end{aligned} \quad (2.4)$$

Let $L \triangleq \prod_{k=1}^{n_{Tr}} P(\mathbf{x}_{Tr}^{(k)}(t) | \lambda)$ be the likelihood function of the observation data. Directly solving (2.4) is not possible

in practice, as the likelihood function in equation (2.4) contains unobservable variables (the true degradation states $S(t)$ in this case). Expectation Maximization (EM) algorithm [71] is applied to solve this problem, where the maximum likelihood estimator is found in an iterative way: the current values of the parameters are used to estimate the unobservable variables (Expectation phase); then, the estimated values of the unknown variables are substituted

into the likelihood function to update the maximum likelihood estimators of the parameters (Maximization phase). The iterative procedures are repeated until the maximum likelihood estimators converge.

To apply the EM algorithm to the HM-GMM model, two auxiliary variables need to be defined first, i.e., forward variable $\alpha_t(S_i)$ and backward variable $\beta_t(S_i)$. The forward variable is defined as the probability of observing the health indicators up to the current time t and that the true degradation state $S(t) = S_i$, given a known HM-GMM λ :

$$\alpha_t(S_i) = P(\mathbf{x}(t_1), \mathbf{x}(t_2), \dots, \mathbf{x}(t), S(t) = S_i | \lambda). \quad (2.5)$$

It is easy to verify that

$$\begin{aligned} \alpha_1(S_i) &= \pi_i b_i(\mathbf{x}(t_1)), \\ \alpha_{t+1}(S_j) &= b_j(\mathbf{x}_{t+1}) \left[\sum_{i=1}^Q \alpha_t(S_i) a_{ij} \right], 1 \leq i \leq Q, 1 \leq j \leq Q, 1 \leq t \leq t_{Tr} - 1, \end{aligned} \quad (2.6)$$

where t_{Tr} represents the observation time length and all the elements in π_i are zero, except the one that corresponds to the i -th element being one.

The backward probability $\beta_t(S_i)$ is defined as the probability of observing the health indicator $\mathbf{x}(t+1), \mathbf{x}(t+2), \dots, \mathbf{x}(t_{Tr})$ from $t+1$ to the end of the observations, given that $S(t) = S_i$ and the model parameters are λ :

$$\beta_t(S_i) = P(\mathbf{x}(t+1), \mathbf{x}(t+2), \dots, \mathbf{x}(t_{Tr}) | S(t) = S_i, \lambda). \quad (2.7)$$

It is easy to verify that $\beta_t(S_j) = \left[\sum_{i=1}^Q b_j(\mathbf{x}(t+1)) a_{ij} \right] \beta_{t+1}(S_i), 1 \leq i, 1 \leq j \leq Q, \beta_{t_{Tr}}(i) = 1, t = t_{Tr} - 1, t_{Tr} - 2, \dots, 1$.

The iterative estimators for the transition probabilities, denoted by a_{ij} , can, then, be derived as follows [72]:

$$a_{ij} = \frac{\sum_{k=1}^{n_{Tr}} \sum_{t=1}^{t_{Tr}} \xi_{Tr,t}^{(k)}(S_i, S_j)}{\sum_{k=1}^{n_{Tr}} \sum_{t=1}^{t_{Tr}} \gamma_{Tr,t}^{(k)}(S_i)}, \quad (2.8)$$

where $\xi_{Tr,t}^{(k)}(S_i, S_j)$ represents the probability of the k -th sample being in S_i at time t and state S_j at time $t+1$,

and is calculated by [72]:

$$\begin{aligned} \xi_{Tr,t}^{(k)}(S_i, S_j) &= P(S(t) = S_i, S(t+1) = S_j | \mathbf{x}_{Tr}^{(k)}(t+1), \lambda) \\ &= \frac{\gamma_{Tr,t}^{(k)}(S_i) a_{ij} b_{Tr,j}^{(k)}(\mathbf{x}_{Tr}^{(k)}(t+1)) \beta_{Tr,t+1}^{(k)}(S_j)}{\beta_{Tr,t}^{(k)}(S_i)}, \end{aligned} \quad (2.9)$$

where $\gamma_{Tr,t}^{(k)}(S_i)$ represents the probability of being in S_i at time t given the health indicator $\mathbf{x}_{Tr}^{(k)}(t)$ and λ for the k -th training sample:

$$\gamma_{Tr,t}^{(k)}(S_i) = \frac{\alpha_{Tr,t}^{(k)}(S_i)\beta_{Tr,t}^{(k)}(S_i)}{p(\mathbf{x}_{Tr}^{(k)}(t)|\lambda)} = \frac{\alpha_{Tr,t}^{(k)}(S_i)\beta_{Tr,t}^{(k)}(S_i)}{\sum_{i=1}^Q \alpha_{Tr,t}^{(k)}(S_i)\beta_{Tr,t}^{(k)}(S_i)}. \quad (2.10)$$

The estimator for the initial state probability $\pi_i, i=1,2,\dots,Q$ is calculated by [71]:

$$\pi_i = \frac{\sum_{k=1}^{n_{Tr}} \gamma_{Tr,t}^{(k)}(S_i)}{n_{Tr}}. \quad (2.11)$$

The estimators of the mean value vectors are derived as [72]:

$$\boldsymbol{\mu}_i = \frac{\sum_{k=1}^{n_{Tr}} \sum_{t=1}^{t_{Tr}} \gamma_{Tr,t}^{(k)}(S_i) \mathbf{x}_{Tr}^{(k)}(t)}{\sum_{k=1}^{n_{Tr}} \sum_{t=1}^{t_{Tr}} \gamma_{Tr,t}^{(k)}(S_i)}. \quad (2.12)$$

Similarly, the covariance matrices of the Gaussian output are calculated by [72]:

$$\boldsymbol{\Sigma}_i = \frac{\sum_{k=1}^{n_{Tr}} \sum_{t=1}^{t_{Tr}} \gamma_{Tr,t}^{(k)}(S_i) (\mathbf{x}_{Tr}^{(k)}(t) - \boldsymbol{\mu}_i)(\mathbf{x}_{Tr,t}^{(k)} - \boldsymbol{\mu}_i)'}{\sum_{k=1}^{n_{Tr}} \sum_{t=1}^{t_{Tr}} \gamma_{Tr,t}^{(k)}(S_i)}. \quad (2.13)$$

2.3.2.2 Degradation state estimation

In this chapter, the forward algorithm [71] is employed to estimate the degradation state of the safety barriers in the online step. Let S_{CM} denote the estimated degradation state from condition monitoring data and $P_{CM,t_k}(S_{CM}), k=1,2,\dots,q$ represent the posterior distribution of S_{CM} given the condition monitoring data up to t_k :

$$P_{CM,t_k}(S_{CM} = S_i) = P(S(t_k) = S_i | \mathbf{x}(t_1), \mathbf{x}(t_2), \dots, \mathbf{x}(t_k), \lambda) \quad (2.14)$$

The posterior probabilities defined in (2.14) can be easily calculated from the forward probabilities defined in (2.15):

$$\begin{aligned}
P_{CM,t_k}(S_{CM} = S_i) &= \frac{P(S(t_k) = S_i, \mathbf{x}(t_1), \mathbf{x}(t_2), \dots, \mathbf{x}(t_k) | \boldsymbol{\lambda})}{P(\mathbf{x}(t_1), \mathbf{x}(t_2), \mathbf{x}(t_3), \dots, \mathbf{x}(t_k) | \boldsymbol{\lambda})} \\
&= \frac{\alpha_{t_k}(S_i)}{\sum_{i=1}^Q \alpha_{t_k}(S_i)}.
\end{aligned} \tag{2.15}$$

In practice, the $\alpha_{t_k}(S_i)$ in (2.15) is calculated recursively, based on (2.5).

At each $t = t_k$, the most likely degradation state, denoted by $S_{CM,MAP}(t_k)$, is, then, determined by finding the state with maximal posterior probability:

$$S_{CM,MAP}(t_k) = \arg \max_{1 \leq i \leq Q} [P_{CM,t_k}(S_{CM} = S_i)], 1 \leq k \leq q. \tag{2.16}$$

2.4 Data integration for DRA

In this section, we first show how to integrate the condition monitoring data with inspection data for reliability updating and prediction of the safety barriers (section 2.4.1). Then, in section 2.4.2, we develop a DRA method based on the updated and predicted reliabilities.

2.4.1 A Bayesian network model for data integration

As in the previous sections, we illustrate the developed data integration method using the i -th safety barrier at $t = t_k$. For simplicity and to avoid confusion, we drop the i and t_k in the notations. To update and predict the reliability, one needs to estimate the degradation state first. Let S_{IN} denote the degradation state estimated from inspection data and S denote the true degradation state. In practice, S_{IN} is subject to uncertainty due to potential imprecision in the inspection and recording by the maintenance personnel. To model such uncertainty, in this chapter, we assume that the reliability of inspection is R_{IN} , and that the maintenance personnel correctly identify the true degradation state with a probability R_{IN} , whereas an inspection error can occur with probability $(1 - R_{IN})$. When an inspection error occurs, it is further assumed that the probabilities for each of the possible degradation states being erroneously identified as the true degradation state are equal to each other:

$$P(S_{IN} = S_i | S) = \begin{cases} R_{IN}, & S = S_i \\ \frac{1 - R_{IN}}{Q - 1}, & S \neq S_i \end{cases} \quad (2.17)$$

where Q is the number of degradation state. It should be noted that other inspection models might also be assumed, depending on the actual problem setting.

In this chapter, a BN is developed to describe the dependencies among S, S_{IN}, S_{CM} , as shown in Figure 2-4. The BN in Figure 2-4 is constructed based on the assumption that given the true degradation state S , the estimated degradation state from condition monitoring data and inspection data are conditional-independent.

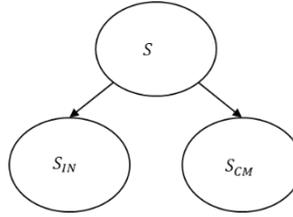


Figure 2-4. A BN model for data integration.

Based on the BN in Figure 2-4, we have

$$P(S, S_{IN}, S_{CM}) = P(S_{IN} | S) P(S_{CM} | S) P(S). \quad (2.18)$$

In (2.18), $P(S)$ measures the prior belief of the analysts on the current degradation states. We assume that is a uniform distribution over all the possible degradation states, indicating that there is no further information to distinguish the states. $P(S)$

The conditional probability distribution $P(S_{IN} | S)$ describes the uncertainty in the inspections and is derived based on (2.17). In (2.17), the reliability of the inspection can be estimated from historical data or assigned based on expert judgments. The conditional probability distribution $P(S_{CM} | S)$ measures the trust one has on the estimated degradation state based on condition monitoring data. Its values can be estimated from validation test data. However, in practice, as validation tests are not always available, $P(S_{CM} | S)$ might also be assigned by experts considering the measurement uncertainty of the sensors and the distance between the neighbouring degradation states.

Once the condition monitoring data and inspection data are available, the observed values of S_{IN} and S_{CM} are known. Suppose we have $S_{CM} = S_j$ and $S_{IN} = S_i$. It should be noted that we choose the state with maximal posterior probability from (2.16) as the observation value of S_{CM} . The two data sources can be naturally integrated by

calculating the posterior distribution of S given the two data sources, denoted by $P_{INT}(S)$. Based on the BN in Figure 2-4, we have:

$$\begin{aligned}
P_{INT}(S) &\triangleq P(S | S_{IN} = S_i, S_{CM} = S_j) \\
&= \frac{P(S, S_{IN} = S_i, S_{CM} = S_j)}{P(S_{IN} = S_i, S_{CM} = S_j)} \\
&= \frac{P(S_{IN} = S_i | S) P(S_{CM} = S_j | S) P(S)}{P(S_{IN} = S_i, S_{CM} = S_j)}
\end{aligned} \tag{2.19}$$

Given the estimated posterior distribution in (2.19), the reliability of the safety barrier can be updated. Suppose the current time is t_k , the updated reliability can be calculated by:

$$R_{SB}(t_k) = \sum_{S \in W} P_{INT, t_k}(S), \tag{2.20}$$

where W is the working set that contains all the working states; $P_{INT, t_k}(S)$ is the posterior probability of the true degradation state after integrating the two data sources at $t = t_k$ and is calculated from (2.19).

Furthermore, at $t = t_k$, we can also predict the reliability of the safety barriers at a future time t_{Fut} . For this, the distribution of the degradation states at $t = t_{Fut}$ is predicted first, using Chapman-Kolmogorov equation [73] and the trained model from the offline step:

$$P_{INT, t_{Fut}}(S) = P_{INT, t_k}(S) \times A^{(t_{Fut} - t_k)}. \tag{2.21}$$

The reliability at $t = t_k$, can be predicted as:

$$R_{SB}(t_{Fut}) = \sum_{S \in W} P_{INT, t_{Fut}}(S). \tag{2.22}$$

2.4.2 Dynamic risk assessment

The updated reliabilities from (2.20), can, then, be substituted into (2.2) for DRA:

$$r_{C_i}(t_k) = f_{ET}(R_{SB_1}(t_k), R_{SB_2}(t_k), \dots, R_{SB_K}(t_k), R_{SB_{K+1}}, \dots, R_{SB_M} | IE), i = 1, 2, \dots, N, \tag{2.23}$$

where in (2.23), $R_{SB_i}(t_k)$ is calculated by (2.20). Similarly, the risk index at a future time t_{Fut} can be predicted by:

$$r_{C_i}(t_{Fut}) = f_{ET}(R_{SB_1}(t_{Fut}), R_{SB_2}(t_{Fut}), \dots, R_{SB_K}(t_{Fut}), R_{SB_{K+1}}, \dots, R_{SB_M} | IE), i = 1, 2, \dots, N, \tag{2.24}$$

where $R_{SB_i}(t_{Fut})$ is calculated by (2.21) and (2.22).

2.5 Application

In this section, the developed method is applied for DRA of an ATWS accident of a NPP [52]. The description of the case study is briefly introduced in section 2.5.1. Then, in section 2.5.2, the developed HM-GMM and the data integration process are presented. The results of the DRA are presented and discussed in section 2.5.3.

2.5.1 System description

ATWS is an accident that can happen in a NPP. In this accident, the scram system, which is designed to shut down the reactor during an abnormal event (anticipated transient), fails to work [74]. An ET has been developed for PRA of the ATWS for a NPP in China [52], as shown in Figure 2-5. In Figure 2-5, T_1ACM represents the failure of the automatic scram system and is the initialling event (IE) considered. Eleven safety barriers ($SB_1 \sim SB_{11}$) are designed to contain the accident. Depending on the states of the safety barriers, 23 sequences can be generated ($SE_{01} - SE_{23}$) [52, 75]. The consequences of the sequences are grouped into two categories, based on their severity; the first group,

$$C_s = \{SE_{03}, SE_{06}, SE_{07}, SE_{08}, SE_{09}, SE_{12}, SE_{13}, SE_{14}, SE_{15}, SE_{18}, SE_{19}, SE_{20}, SE_{21}, SE_{22}, SE_{23}\}, \quad (2.25)$$

represents the event sequences with severe consequences, whereas the remaining event sequences have non-severe consequences [75]. The risk index *Risk* considered in this chapter is the conditional probability of having severe consequences, given the initialling event ($IE = T_1ACM$):

$$Risk \triangleq P(C_s | IE) = f_{ET}(R_{SB_1}, R_{SB_2}, \dots, R_{SB_M} | T_1ACM), \quad (2.26)$$

where the model function $f_{ET}(\bullet)$ is determined from the ET in Figure 2-5 and $R_{SB_1}, R_{SB_2}, \dots, R_{SB_M}$ are the reliabilities of the safety barriers, calculated based on the component failure probabilities. It should be noted that the failure probabilities for SB_7 and SB_8 change depending on the event sequence that occurs.

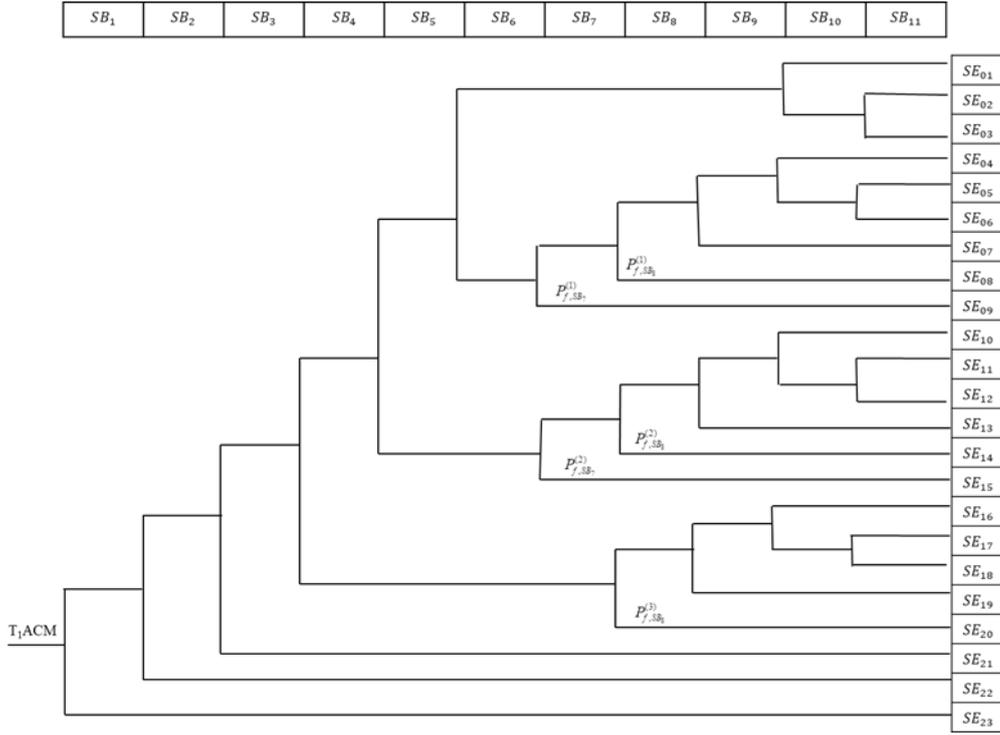


Figure 2-5. ET for the ATWS.

The condition monitoring data of the bearing come from the bearing degradation dataset from university of Cincinnati [76]. The dataset contains four samples and for each sample, raw condition monitoring data are collected in real time by measuring the vibration acceleration signals. On the other hand, the inspection can be performed at some given time instants to identify the different degradation states. In this case study, we consider four states (healthy, minor degradation, medium degradation, sever degradation).

2.5.2 Dynamic risk assessment

DRA of the ATWS is carried out following the procedures in Figure 2-3, where the real data set from [76] is used as historical training data. In the offline step, feature extraction needs to be conducted first. Three features are extracted from the vibration signals using the time domain method:

$$\begin{cases} x_1(t_i) = \frac{1}{(t_i - t_{i-1}) \cdot f} \sum_{j \in (t_{i-1}, t_i)} c_j^2 \\ x_2(t_i) = \sqrt{\frac{1}{(t_i - t_{i-1}) \cdot f} \sum_{j \in (t_{i-1}, t_i)} (c_j - \bar{c})^2} \\ x_3(t_i) = \frac{1}{(t_i - t_{i-1}) \cdot f} \sum_{j \in (t_{i-1}, t_i)} c_j \end{cases} \quad (2.27)$$

where x_1 is the average power of vibration, x_2 is the root mean square, x_3 is the mean value of vibration. In (2.27), f is the sampling frequency, $(t_i - t_{i-1}) \cdot f$ is the number of sampling points in time interval $[t_{i-1}, t_i]$, and c_j is the vibration signal. The results of data process are shown in Figure 2-6.

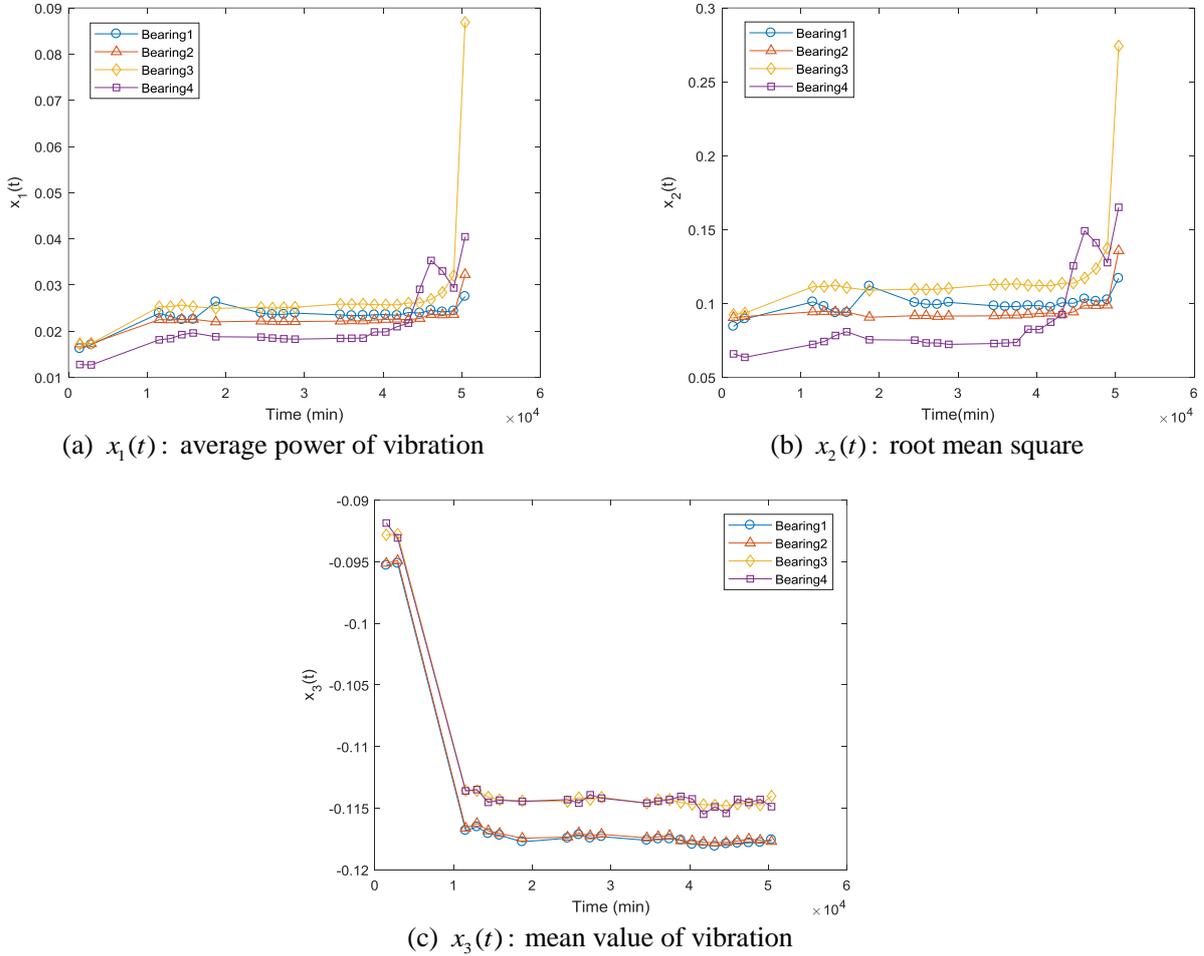


Figure 2-6. Extracted degradation indicators.

The estimated degradation state S_{IN} and S_{CM} are, then, integrated using (2.19). Note that in (2.17), the reliability of the inspection data is set to $R_{IN} = 0.8$. Then, the value of $P(S_{IN} | S)$ in (2.19) can be derived easily from (2.17). The values of $P(S_{CM} | S)$ are assigned by considering the distance between the neighbouring degradation states: the closer the states are, the more likely a misclassification might happen. For example, the normalized distance between S_2 and S_3 is:

$$\frac{d(\boldsymbol{\mu}_2, \boldsymbol{\mu}_3)}{\sum_{i=1}^4 d(\boldsymbol{\mu}_i, \boldsymbol{\mu}_3)} = 0.4807, \quad (2.28)$$

and the normalized distance between S_3 and S_4 is:

$$\frac{d(\boldsymbol{\mu}_4, \boldsymbol{\mu}_3)}{\sum_{i=1}^4 d(\boldsymbol{\mu}_i, \boldsymbol{\mu}_3)} = 0.1108, \quad (2.29)$$

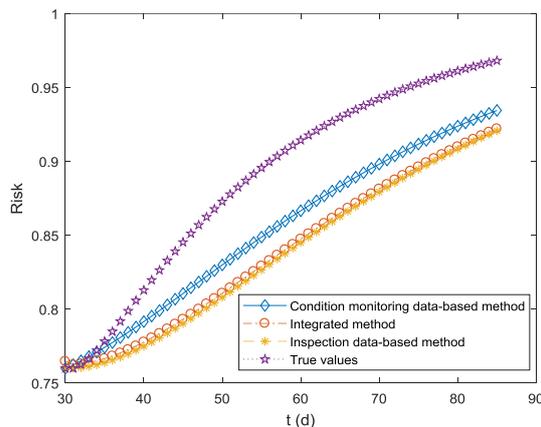
where $d(\cdot)$ is the Euclidean distance. Thus, we set $P(S_{CM} = S_2 | S = S_3) = 0.1$ and $P(S_{CM} = S_4 | S = S_3) = 0.2$. The value of the other elements in $P(S_{CM} | S)$ are determined in a similar way and reported in Table 2-1. Once the integrated estimation of the degradation state is obtained, risk updating and prediction can be performed by (2.23) and (2.24), respectively.

Table 2-1. Values of $P(S_{CM} | S)$.

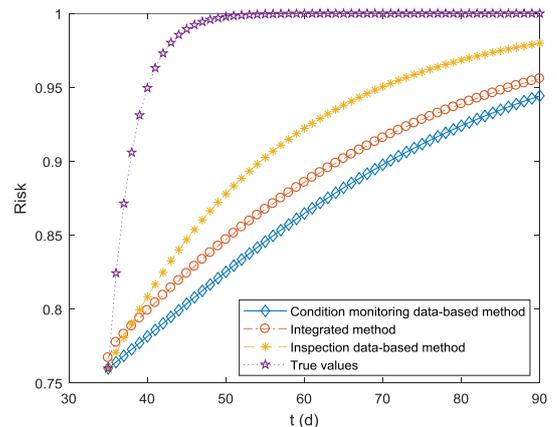
	$S = S_1$	$S = S_2$	$S = S_3$	$S = S_4$
$P(S_{CM} = S_1 S)$	0.9	0	0	0
$P(S_{CM} = S_2 S)$	0.05	0.9	0.1	0.1
$P(S_{CM} = S_3 S)$	0.05	0.1	0.9	0.1
$P(S_{CM} = S_4 S)$	0	0	0	0.8

2.5.3 Results

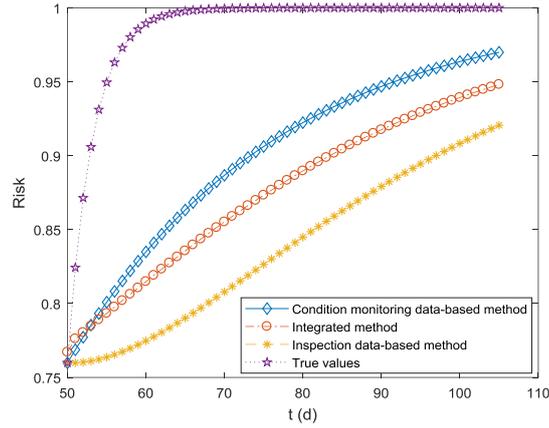
The results of risk updating and prediction at $t = 30, 35$ and $50(d)$ are given in Figure 2-7. In Figure 2-7, we also show the results from using only condition monitoring data and inspection data, for comparison.



(a) $t = 30 (d)$



(b) $t = 35 (d)$



(c) $t = 50$ (d)

Figure 2-7. The results of risk updating and prediction.

As shown in Figure 2-7(a), at $t = 30$ (d), the results from all the three methods are close to each other. However, when compared to the true risk values, the updated and predicted risks from all the three methods show relatively large discrepancies. This discrepancy is mainly due to the estimation errors in the offline step, as we have only four samples in the training data set. A possible way to increase the accuracy of risk updating is, then, to increase the sample size of the training data in the offline step.

At $t = 35$ (d), the inspection data give correct information on the current degradation state while condition monitoring data do not. From Figure 2-7(b), it can be seen that the developed data-integration method improves the DRA results from the condition monitoring data-based method, as it integrates the correct information from inspection data. On the other hand, when the inspection data fail to give the correct information ($t = 50$ (d)), it can be seen from Figure 2-7(c) that the developed data integration method can also correct the misleading results obtained from using only the inspection data. Hence, in general, applying the developed data integration method can achieve a more robust DRA result than using the two data sources individually.

2.6 Conclusion

In this chapter, a novel framework has been presented to integrate condition monitoring data and inspection data in DRA. A HM-GMM has been developed to estimate the degradation states of the safety barriers based on the condition monitoring data. The estimated degradation states are integrated with the inspection data for DRA by a BN model. A real-world application on a NPP accident risk assessment model (an ET) has been conducted. The results

show that, as expected, integrating the two data sources into the DRA gives more accurate and robust results than using any one of the two individual data sources.

Chapter 3 Dynamic business continuity assessment using condition monitoring data

Business organizations are faced with threats from various disruptive events, such as natural disaster, malicious attacks and equipment failures, etc. Business continuity management (BCM) has been demonstrated as a comprehensive and proactive method to prevent disruptive events from impacting the business operation and reduce the potential losses. However, most existing BCM models are developed for time-static problems, where the factors related to business continuity indexes are considered not varying over time. On the contrary, in practice, various time-dependent factors influence business continuity, such as the degradation of safety barriers, the dynamic behaviour of profits and losses, etc. The aim of this chapter is to develop a simulation-based scheme for dynamic business continuity assessment (DBCA) using condition monitoring data, accounting for the time-variant factors in the BCA process.

The reminder of this chapter is arranged as follows. Section 3.1 briefly reviews the business continuity assessment methods. Section 3.2 presents the proposed numerical metrics for DBCA. Section 3.3 presents an integrated framework of DBCA. Section 3.4 shows an application of the proposed framework on a NPP. Section 3.5 summarizes this chapter.

3.1 State of the art

Most existing researches on BCA only focus on qualitative analysis [17]. For instance, the necessity and benefit of implementing BCM in a supply chain has been discussed in [18]. In [77], a framework for the design, implementation and monitoring of BCM programs has been exploited. A framework that integrates business continuity and disaster recovery planning for efficiently resuming critical operation has been proposed in [10]. In [78], BCM has been compared with the conventional risk management methods, showing that BCM considers not only the protection of the system against the disruptive event, but also the recovery process during and after the accident. In [20], a framework for information system continuity management has been introduced. Standards

concerning BCM of the Brazilian gas supply chain have been discussed in [79]. In [21], the conceptual foundation of business continuity management has been presented in the context of societal safety.

From an engineering point of view, it is needed to define numerical indexes that support quantitative BCA. A few numerical indexes have been defined in [9]. e.g., maximum tolerable period of disruption (MTPD), minimum business continuity objective (MBCO) and recovery time objective (RTO). However, these numerical indexes are usually directly estimated based on expert judgements. Only a few attempts exist concerning developing quantitative models to evaluate these numerical indexes. For example, a statistical model integrating Cox's model and Bayesian networks has been proposed to model the business continuity process [38]. In [80], the BCM outsourcing and insuring strategies have been compared based on the organization characteristics and the relevant data through a two-step fuzzy cost-benefit analysis. Two probabilistic programming models have been developed to determine appropriate business continuity plans given epistemic uncertainty of input data in [39]. In [40], a new model for integrated business continuity and disaster recovery planning has been presented, considering the multiple disruptive incidents that might happen simultaneously. An integrated framework was developed for quantitative business continuity analysis, where four numerical metrics were proposed to quantify the business continuity level based on the potential loss caused by the disruptive event [14].

As shown in the reviews above, the existing quantitative BCM approaches only apply for time-static problems. On the contrary, in practice, various time-dependent factors influence the business continuity, such as the degradation of safety barriers, the dynamic behaviour of profits and losses, etc. On the other hand, as sensor technologies and computing resources advance, it is possible to capture these dynamic factors even in real-time, based on online-collected condition monitoring data [42, 81]. For example, a condition-based fault tree has been used for dynamic risk assessment (DRA) [43], where the condition monitoring data are used to update the failure rates of specific components and predict the reliability. In [44], a Bayesian reliability updating method has been developed for dependent components by using condition monitoring data. In [32], a holistic framework that integrates the condition monitoring data and statistical data has been proposed for DRA. A sequential Bayesian approach has been developed in [82] for dynamic reliability assessment and remaining useful life prediction for dependent competing failure processes.

3.2 Numerical metrics for dynamic continuity assessment

An integrated, quantitative framework for modeling BC has been developed in [14], based on the potential losses caused by the disruptive events. The business process is divided into four sequential stages: preventive stage, mitigation stage, emergency stage and recovery stage. Various safety measures are designed in different stages to guarantee the continuity of the business process. Business continuity value (BCV) was formally defined as [14]:

$$BCV([0, T]) = 1 - \frac{L([0, T])}{L_{\text{tol}}} \quad (3.1)$$

where L denotes the loss in $[0, T]$ from the disruptive event; T is the evaluation horizon for the assessment (e.g., the lifetime of the system); L_{tol} is the maximum loss that can be tolerated by an organization. Equation (3.1) measures the relative distance to a financially dangerous state by taking into account the possible losses generated by the business disruption. It should be noted that only one business process is considered in this chapter, while in practice, an organization might be involved in multiple business processes at the same time. For multiple-business system, the developed framework can be naturally extended based on the potential losses and profit generated by the different business processes together.

The business continuity metrics discussed above are time-static in nature. In practice, however, various factors influencing the business continuity are time-dependent. These dynamic influencing factors can be grouped into internal factors and external factors. Internal factors are related to the safety barriers within the system of interest, such as the dynamic failure behavior of the safety barriers (e.g., corrosion, fatigue crack and wear [60]). External factors refer to the influence from external environment. For example, variations in the price of products will affect the accumulated revenue of the organization, and, then, the tolerable loss in Equation (3.1). To consider these factors, the business continuity metrics are extended to the dynamic cases:

$$DBCV([t, t+T]) = 1 - \frac{L([t, t+T])}{L_{\text{tol}}(t)}, \quad (3.2)$$

where t is the time instant when the dynamic business continuity assessment is carried out; $DBCV([t, t+T])$ represents the business continuity value evaluated at time t , for a given evaluation horizon of T ; $L([t, t+T])$ represents the potential losses in $[t, t+T]$; $L_{\text{tol}}(t)$ denotes the maximal amount of losses that the company can tolerate at t , before having troubles in recovery. The physical meaning of DBCV is the relative distance to a financial dangerous state at time t , by considering the possible losses in $[t, t+T]$ due to business disruption; it measures the

dynamic behavior of business continuity in a time interval of interest $[t, t + T]$. By calculating the DBCV at different t , the dynamic behavior of business continuity can be investigated.

In [14], two kinds of losses need to be considered when calculating $L([t, t + T])$: direct loss and indirect loss. Direct loss, denoted by $L_d([t, t + T])$, represents the losses that are caused directly by the disruptive event. For example, in a NPP leakage event, $L_d([t, t + T])$ includes all equipment damage directly caused by the event. Indirect loss, denoted by $L_{in}([t, t + T])$, is the revenue loss suffered during the shutdown of the plant in the recovery process. Hence, the total loss is calculated by:

$$L([t, T + T]) = L_d([t, t + T]) + L_{in}([t, t + T]). \quad (3.3)$$

The DBCV defined in (3.2) is a random variable. Three numerical metrics are, then, proposed for its quantification:

$$EDBCV = E[DBCV] \quad (3.4)$$

$$P_{BI}([t, t + T]) = \Pr(BCV < 1, t) \quad (3.5)$$

$$P_{BF}([t, t + T]) = \Pr(BCV < 0, t) \quad (3.6)$$

where $EDBCV$ denotes the expected value of the dynamic business continuity value. A higher value $EDBCV$ indicates higher business continuity. $P_{BI}([t, t + T])$ represents the probability that at least one disruptive event causes business interruption in time interval $[t, t + T]$; $P_{BF}([t, t + T])$ is the probability of business failure $[t, t + T]$, meaning that the losses caused by the disruptive event are beyond the system tolerable losses. It measures the risk that a business cannot recover from disruptive events, if a plant with an age of t continues operation for other T units of time.

3.3 An integrated framework for dynamic business continuity assessment

In this section, we first present an integrated modeling framework for the dynamic business continuity metrics defined in Section 3.2. Then, particle filtering (PF) is used to estimate the potential loss L_{tol} in real time using condition monitoring data (section 3.3.1). The quantification of tolerable losses L_{tol} is, then, discussed Section 3.3.2.

3.3.1 The integrated modeling framework

To model the dynamic business continuity, we make the following assumptions:

- (1) The evolution of the disruptive event is modeled by an ET. The possible consequences of the disruptive event are classified as $C_i, i = 1, 2, \dots, n$ based on the severity of the consequence.
- (2) Some safety barriers in the ET are subject to degradation failure processes. Condition monitoring data are available for these safety barriers at predefined time instants $t_k, k = 1, 2, \dots, q$.
- (3) The other safety barriers have constant failure probabilities.
- (4) Recovery means repairing the failed component and restarting the business. The time from the recovery for consequence C_i is a random variable $t_{recv,i}$, with a probability density function (PDF) $f_{recv,i}$.

An integrated framework for DBCA is presented in Figure 3-1. The DBCA starts from collecting condition monitoring data, denoted as c_k , which is collected from sensors and can be used to characterize the degradation states of the component. The degradation of the safety barriers is estimated based on the condition monitoring data and used to update the estimated losses. Then, the potential profits are predicted and used to calculate the tolerable losses. Finally, the dynamic business continuity metrics can be calculated.

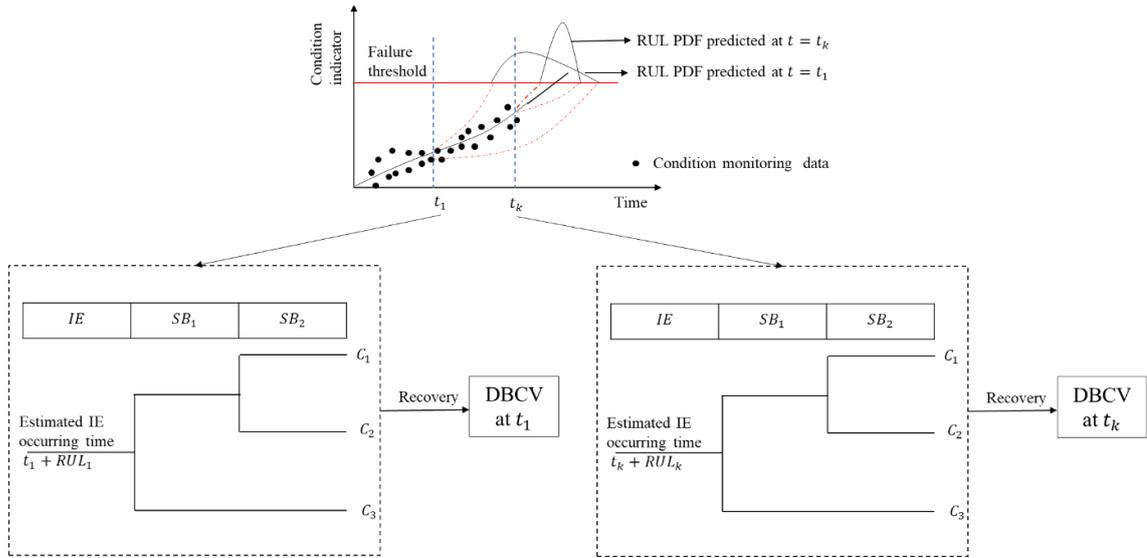


Figure 3-1. An integrated model for DBCA.

3.3.2 Loss modeling

To capture the dynamic failure behavior of the safety barrier, PF is employed to estimate its degradation and predict its remaining useful life (RUL) [83-86]. Suppose the degradation process of a safety barrier can be described by Equation (3.7), in which the current state x_k at the k -th time step depends on the previous state x_{k-1} . Here, f is a non-linear function and v_k represents process noise that follows a known distribution. In practice, Equation (3.7) is often determined based on physics-of-failure models [83, 87]:

$$\mathbf{x}_k = f(\mathbf{x}_{k-1}, \mathbf{v}_k) \quad (3.7)$$

A sequence of condition monitoring data \mathbf{z}_k is assumed to be collected at predefined time points t_k . The sequence of measurement values is assumed to follow an observation function:

$$\mathbf{z}_k = h(\mathbf{x}_k, \boldsymbol{\sigma}_k) \quad (3.8)$$

where h is the observation function (possibly nonlinear), $\boldsymbol{\sigma}_k$ is the observation noise vector sequence of known distribution. The measurement data \mathbf{z}_k are assumed to be conditionally independent given the state process \mathbf{x}_k . Equation (3.8) quantifies the observation noise from the sensors.

The PF follows two steps [88]:

- 1) Filtering step, where the available condition monitoring data \mathbf{Z}_k are used to estimate the current degradation state of the system.
- 2) Prediction step, in which the RUL is predicted based on the estimated degradation state and the condition monitoring data.

In the filtering step, the posterior PDF of variable \mathbf{X}_k is approximated by the sum of weighted particles

$\{\mathbf{x}_k^{(i)}, \omega_k^{(i)}\}$:

$$p(\mathbf{x}_k | z_1, z_2, \dots, z_k) \approx \sum_{i=1}^{N_s} \omega_k^{(i)} \delta(\mathbf{x}_k - \mathbf{x}_k^{(i)}) \quad (3.9)$$

where $p(\mathbf{x}_k | z_1, z_2, \dots, z_k)$ is the estimated posterior PDF of \mathbf{x}_k , δ is the Dirac Delta function, $\omega_k^{(i)}$ is the weight assigned to particle $\mathbf{x}_k^{(i)}$ and is generated by sequential importance sampling [87]. When the new measurement z_k is available, the required posterior distribution of the current state \mathbf{x}_k can be obtained by updating the prior distribution:

$$p(\mathbf{x}_k | \mathbf{z}_k) = \frac{p(z_k | \mathbf{x}_k) p(\mathbf{x}_k | \mathbf{z}_{k-1})}{\int p(z_k | \mathbf{x}_k) p(\mathbf{x}_k | \mathbf{z}_{k-1}) d\mathbf{x}_k} \quad (3.10)$$

where $p(z_k | \mathbf{x}_k)$ is the likelihood function that can be derived from the observation function (3.8). Generally, if the samples $\mathbf{x}_k^{(i)}$ are drawn from the sampling distribution $p(\mathbf{x}_k | \mathbf{z}_k)$, then, the particle weight can be updated with a new observation z_k , as follows [32]:

$$\omega_k^{(i)} = \omega_{k-1}^{(i)} \frac{p(z_k | \mathbf{x}_k^{(i)}) p(\mathbf{x}_k^{(i)} | \mathbf{x}_{k-1}^{(i)})}{p(\mathbf{x}_k^i | \mathbf{x}_{0:k-1}^i, \mathbf{z}_k)}. \quad (3.11)$$

Note that the weights are normalized as $\sum_{i=1}^{N_s} \omega_k^{(i)} = 1$.

Then, in the prediction step, the RUL associated to the i -th particle at $t = t_k$ can be estimated through state function (3.7) by simulating the evolution trajectory of the particles until they reach the failure threshold z_{th} :

$$RUL_k^{(i)} = \left\{ (T_{th}^{(i)} - 1 - k) \mid x_{T_m^{(i)}-1} < z_{th}, x_{T_m^{(i)}} \geq z_{th} \right\}, \quad (3.12)$$

where $T_{th}^{(i)}$ is the first time the particle reaches the threshold z_{th} . Thus, the PDF of the RUL can be generated by:

$$p(RUL|\mathbf{z}_k, z_{th}) \approx \sum_{i=1}^{N_k} \omega_k^{(i)} \delta(RUL - RUL_k^{(i)}). \quad (3.13)$$

The predicted $RUL_k^{(i)}, i=1,2,\dots,N_s$ can, then, be used in a simulation process to generate samples of the total loss L , according to Equation (3.3). The procedures are summarized in Algorithm 2, where P_{ID} is the indirect loss per unit of time.

3.3.3 Tolerable losses modeling

Budget limitations are the primary driver of resilience-enhancing investments [89], which influence protection, prevention, and recovery capabilities of system. Tolerable losses L_{tol} depend on the cash flow of the company and also the risk appetite of the decision maker [9]. Therefore, we assume that the tolerable loss at t_k is proportional to the cash flow $Q(t_k)$ of the company at t_k ,

$$L_{tol}(t_k) = Q(t_k) \cdot \alpha \quad (3.14)$$

For example, $\alpha = 0.1$ means 10% of the current cash flow can be used to withstand potential losses caused by a disruptive event.

We make the following assumptions to model the dynamic behavior of cash flows:

(1) At $t = 0$, there is an initial capital of Q_0 .

(2) Installment is used for the company to purchase the asset, where an equal repayment of C_p is paid each month for N_p months.

It is noteworthy that the cash flow $Q(t)$ depends on the profit earned by the normal operation of the asset:

$$Q(t_k) = Q_0 + I(t_k) - C_o(t_k) - \sum_{i=1}^k (\Psi \cdot C_p(t_i)), \quad (3.15)$$

where Q_0 is the initial capital, $I(t_k)$ is the accumulated revenues of the organizations up to t_k by selling the product of the asset. For example, in a NPP, $I(t_k)$ is determined by the electricity price [90], in the oil exploitation, $I(t_k)$ depends on the petroleum price [91]. $C_o(t_k)$ is the operational cost in $[0, t_k]$, $C_p(t_i)$ is the amount of repayment of the installment in $[t_{i-1}, t_i]$, which can be modeled by [92]:

$$C_p = \frac{(IN_{\text{tot}} - D_p)}{N_p} (1 + \rho)^{N_p}, \quad (3.16)$$

where IN_{tot} denotes the total investment and equals the whole value of the system, D_p represents the down payment, ρ is the interest rate, Ψ is an indicator function:

$$\Psi = \begin{cases} 1, & \text{if } t \leq N_p \\ 0, & \text{otherwise} \end{cases}, \quad (3.17)$$

where N_p is the repayment period.

3.4 Case study

In this section, we consider a NPP for the DBCA, as a case study [62]. The developed methods are utilized to evaluate the business continuity of the NPP at different ages $t = 1, 2, \dots, 40$ (year) and different evaluation horizons $T = 1, 2, \dots, 60$ (year). The evaluation is made with reference to a specific risk scenario, SGTR event.

The targeted system is briefly introduced in section 3.4.1. Subsequently, Section 3.4.2 presents the RUL prediction for a SGTR and the modeling of the potential losses. The time-dependent L_{tot} is calculated in section 3.4.3. Section 3.4.4 presents the results and discussions.

3.4.1 System description

For illustrative purposes, it is assumed that the NPP has one reactor with a capacity of 550 MW. It is also assumed that the NPP is subject to the threat of only one disruptive event, the SGTR. The whole value of the NPP is 10^9 € and the operator purchases the NPP using an installment, where the down payment is $5 \cdot 10^8$ € and the repayment period is 10 years with an interest rate of 2%.

SGTR is a potential accident that is induced by the degradation of the tubes in the steam generator, which can lead to tube cracking and rupture [93]. Steam generator tubes transfer the heat from the reactor core to the cooling water that is transformed into steam to drive turbines and produce electricity [62]. The steam generator tube is often manufactured with alloy material to attain the high structural integrity and prevent leakage of radioactive materials. An ET has been developed for probabilistic risk assessment (PRA) of the SGTR for a NPP in South Korea, as shown

in Figure 3-2. In Figure 3-2, eight safety barriers ($SB_1 \sim SB_8$) are designed to control the accident and mitigate its impact. Depending on the states of the safety barriers, 28 sequences are generated ($S_1 \sim S_{28}$). Based on the degree of their severities, the consequence of the sequences can be categorized into two groups. The first group,

$$C_{S1} = \{SE_1, SE_2, SE_4, SE_6, SE_7, SE_9, SE_{11}, SE_{12}, SE_{14}, SE_{16}, SE_{20}, SE_{24}\} \quad (3.18)$$

represents the event sequences in which a SGTR occurs but the consequence is contained by the safety barriers without causing severe damages. The remaining event sequences form the second group C_{S2} represent severe consequences of core damage. Regarding C_{S1} , albeit no severe losses have been caused, normal production of the NPP is disturbed because the ruptured tube has to be repaired. For C_{S2} , it is assumed that the NPP has to be shut down permanently and the losses incurred are denoted by C_{CD} .

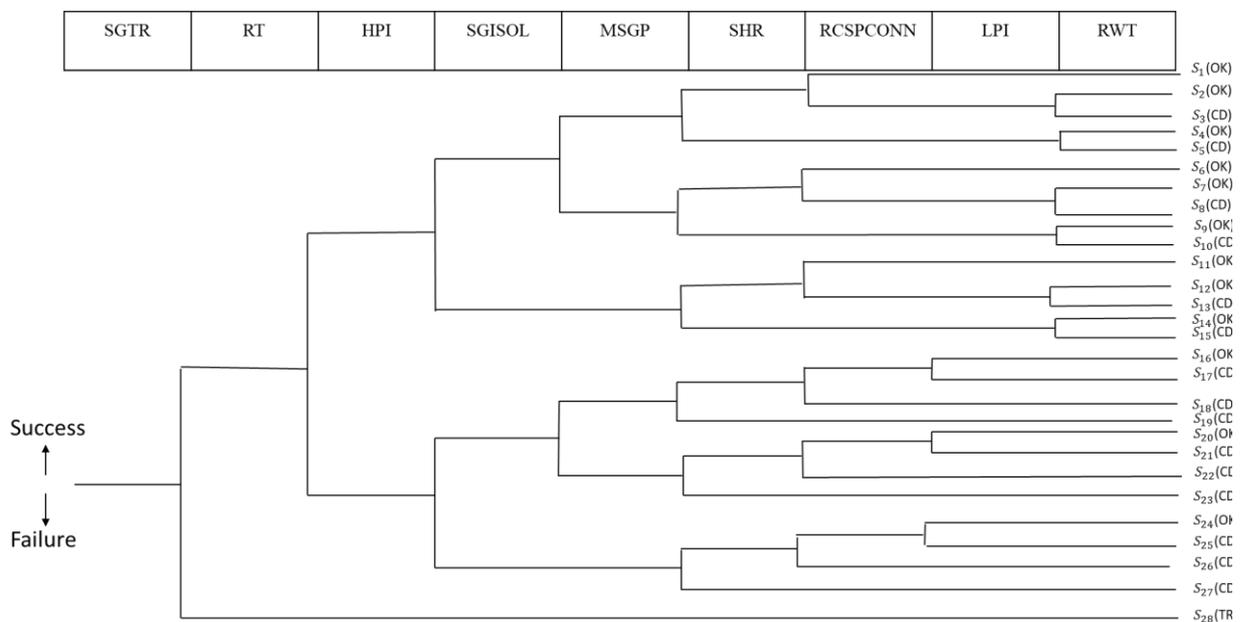


Figure 3-2. ET for SGTR accident [62].

Table 3-1. Safety barriers in the target system [94, 95].

Safety barrier	Failure probability	Description
Reactor trip (RT)	$P_{RT} = 1.8 \times 10^{-4}$	When there is off-normal condition, the protection system automatically inserts control rods into the reactor core to shut down nuclear reaction.
High pressure safety injection (HPI)	$P_{HPI} = 4.6 \times 10^{-4}$	Inject cool water (at a pressure of about 13.79 MPa) into the reactor coolant system (RCS) to cool the reactor core and provide RCS inventory make-up.
Main steam isolation valve (SGISOL)	$P_{SGI} = 1.0 \times 10^{-4}$	A valve used to isolate the affected steam generator (SG).
Maintain the affected SG pressure (MSGP)	$P_M = 1.5 \times 10^{-4}$	Maintain the affected SG pressure through the pressurizer.
Secondary heat removal (SHR)	$P_{SHR} = 3.4 \times 10^{-5}$	Heat removal by unaffected SG.

Reactor coolant system pressure control (RCSPCON)	$P_{\text{RCSM}} = 1.0 \times 10^{-2}$	Open the turbine bypass valve to control the secondary side pressure.
Low pressure safety injection (LPI)	$P_{\text{LPI}} = 4.6 \times 10^{-4}$	Inject cool water (at a pressure of about 1.03MPa) to cool down the RCS and provide RCS inventory make-up.
Refill RWT (RWT)	$P_{\text{RWT}} = 2.4 \times 10^{-8}$	Refill water storage tank.

The crack growth process that leads to SGTR can be monitored through non-destructive inspection (e.g., ultrasonic testing [96], eddy current testing [97]). In practice, this is done during planned shutdowns of the NPP, often during the refueling stage. The condition monitoring data collected from these inspections are, then, used for the dynamic business continuity assessment.

3.4.2 Particle filtering and loss modeling

The first step is to update the occurrence probability of the initiating event, based on the condition monitoring data. For illustrative purposes, the evolution of the tube crack growth process is assumed to follow the Paris-Erdogan model, which has been applied to model SGTR in [24, 95],

$$\frac{da}{dt} = C(\Delta K)^m, \Delta K = \Delta\sigma\sqrt{\pi a}, \quad (3.19)$$

where a is the crack length, C and m are constant parameters related to the component material properties, ΔK is the stress intensity factor, $\Delta\sigma$ is the stress range. The model can be rewritten in the form of a state transition function [98]:

$$a_k = C_k (\Delta\sigma\sqrt{\pi a_k})^{m_k} dt + a_{k-1} \quad (3.20)$$

The crack size a_k at $t = t_k$ is obtained from non-destructive inspection, such as ultrasonic testing; the corresponding observation z_k is:

$$z_k = a_k + \delta_k, \quad (3.21)$$

where δ_k is the observation noise.

PF is used to estimate the degradation state and predict the RUL. The results are shown in Figure 3-3 and Figure 3-4, respectively. The number of particles simulated is $N_s = 5000$. It should be noted that for the tube degradation process, the state vector \mathbf{x} includes the crack size a and the model parameter variables C , m . The initial values for these variables are drawn uniformly from the intervals of values listed in Table 3-2.

$$\begin{cases} C_k = C_{k-1} + N(0, \sigma_c^2) \\ m_k = m_{k-1} + N(0, \sigma_m^2) \end{cases} \quad (3.22)$$

Table 3-2. Initial intervals for the parameters.

Parameters	Initial interval
C	[0.1,0.2]
m	[1.1,1.3]
σ_c	$[0.9 \times 10^{-3}, 0.2 \times 10^{-2}]$
σ_m	$[0.9 \times 10^{-3}, 0.2 \times 10^{-2}]$
σ_o	[0.65,0.85]

The results of PF are shown in Figure 3-4, where we find that the RUL prediction results become more accurate when more condition monitoring data are available.

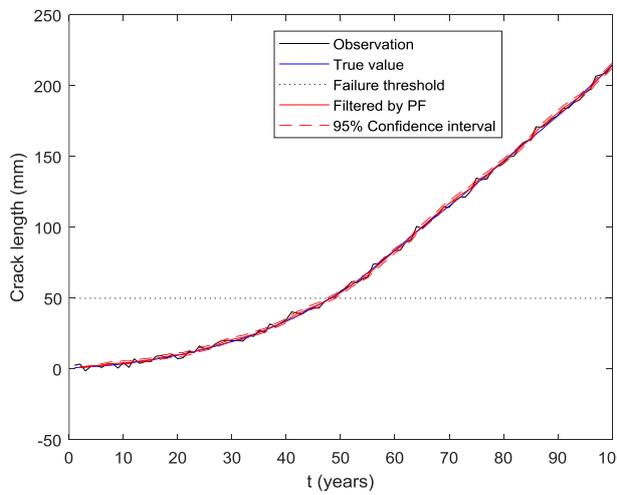


Figure 3-3 Crack growth process.

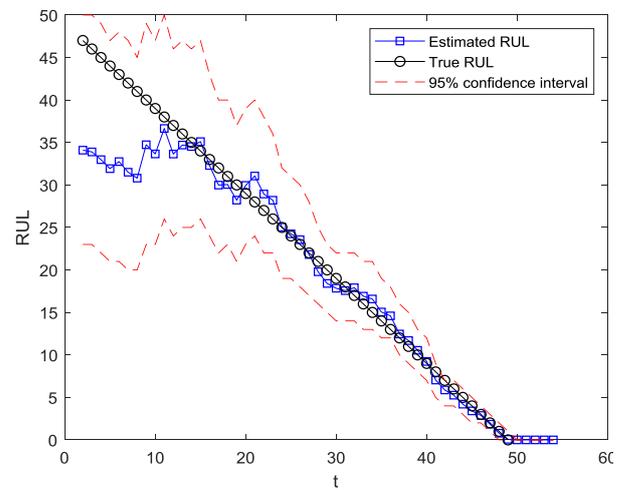


Figure 3-4 RUL Prediction results.

Afterwards, the loss $L([t, t+T])$ in Equation (3.1) can be calculated. The losses caused by a SGTR event, include the direct losses and indirect losses. In this case study, the direct losses, denoted by L_d , equal to the value of the damaged equipment. For the consequence C_{S1} , L_d is identical to the value of the ruptured tube. For the consequence C_{S2} , L equals the whole value of the NPP production since the NPP needs to be shutdown. In this chapter, we assume that if C_{S2} occurs, we have $L = 5 \cdot 10^9 \text{ €}$ [90].

The indirect losses L_{in} are calculated considering the revenue losses during the recovery process, which depends on the recovery time and electricity price. Due to the common use of lognormal distribution for modeling the repair process [99-101], we also assume that the recovery time follows a lognormal distribution with the parameters summarized in Table 3-3, where ε and β are parameters of the lognormal distribution, whose PDF is

$$f(t_{recv}) = \begin{cases} \frac{1}{\sqrt{2\pi}\sigma t_{recv}} e^{-\frac{(\ln(t_{recv})-\varepsilon)^2}{2\beta^2}}, & t_{recv} > 0 \\ 0, & t_{recv} \leq 0. \end{cases} \quad (3.23)$$

Then, the value of L_{in} is calculated by Monte Carlo simulation [102].

Table 3-3. Values of the recovery model parameters.

Parameter	Description	Value
ε	The mean value of the lognormal distribution.	1 year
β	The variance value of the lognormal distribution.	0.1 year ²

3.4.3 Tolerable loss modeling

As discussed in Section 3.3, the tolerable loss is proportional to the cash flow and should be modeled through Equation (3.15). For the NPP, $I(t_k)$ depends on the electricity price, which often exhibits large variabilities. In this chapter, we use the following model to simulate the stochastic behavior of the electricity price [103]:

$$dx_t = \theta\tau(t)(\mu_p - x_t)dt + \sigma\sqrt{\tau(t)}dW_t + dZ_t \quad (3.24)$$

where x_t is the electricity price at t , $\theta > 0$ and μ_p is the mean value of the price, W_t is a standard Brownian motion and Z_t is a compound Poisson process with levy measure $\nu(dx) = \lambda g(x)dx$, λ is the jump intensity and g is the density of the jump size distribution, $\tau(t)$ is a positive stochastic process which satisfies:

$$\tau(t) = s(t) + \nu(t) \quad (3.25)$$

where $s(t)$ is a deterministic, time-dependent and positive seasonal component, which is often modeled by a trigonometric function:

$$S_1(t) = a_1 \sin\left(\frac{a_2 + 2\pi t}{5}\right) + a_3 \left(\frac{a_4 + 2\pi t}{251}\right) + a_5. \quad (3.26)$$

The values of the seasonal component parameters are shown in Table 3-4.

Table 3-4. Values of the seasonal component parameters of the spot prices.

Parameter	Value
a_1	0.41

a_2	1.90
a_3	0.40
a_4	43.11
a_5	0.29

$\nu(t)$ is a stochastic process, representing the stochastic part of the time change. The Cox-Ingersoll-Ross process [104] is used to model $\nu(t)$,

$$d\nu(t) = \kappa(\eta - \nu(t))dt + \sqrt{\nu(t)\sigma_2}dW_2(t). \quad (3.27)$$

By using Itô's lemma [103], Equation (3.24) can be solved and we can derive the following form:

$$x(t) = x(0) + \int_0^t \theta(\mu - x(t))dt + \int_0^t \sigma\sqrt{\tau(t)}dB(t) + \int_0^t dZ(t). \quad (3.28)$$

The parameters of the stochastic electricity model are tabulated in Table 3-5, which is estimated from the German EEX¹ (a market platform for energy and commodity products), from 12.03.2009 until 31.12.2013. The interested readers may refer to details and derivations in [103].

Table 3-5. Parameters in the stochastic electricity model [103].

Parameter	Value
x_0	40
θ	0.22
μ	50
σ	5.98
dt	1
λ	0.12
μ_1	1.02
σ_1	1.35

Eventually, the generated stochastic electricity price trajectory can be used to model the profit and potential losses. The operation cost $C_o(t_k)$ in Equation (3.15) is set as constant 20€/MWh, which includes the cost of uranium fuel and the cost of disposing used fuel and wastes [105]. Finally, the cash flow at different time points is shown in Figure 3-5. We can see that the accumulated profit is small at the beginning. This is because this period is still under

¹ <https://www.eex.com>, accessed 2019-09-12

the repayment period and a large amount of the revenue is used for repaying the installment. After $t = 10$ years, the repayment is paid off and, thus, the profit increases significantly.

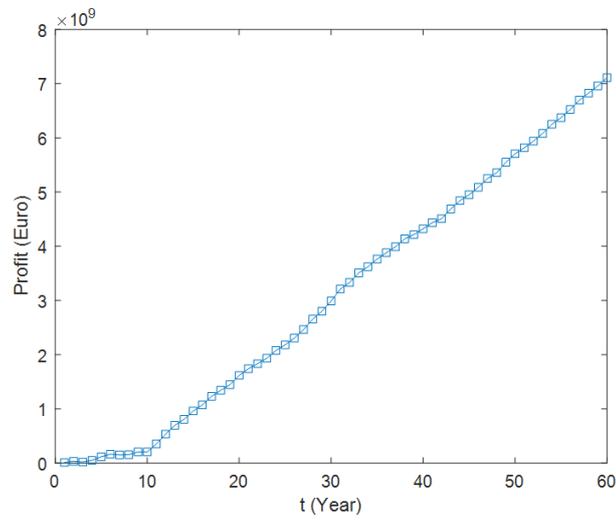
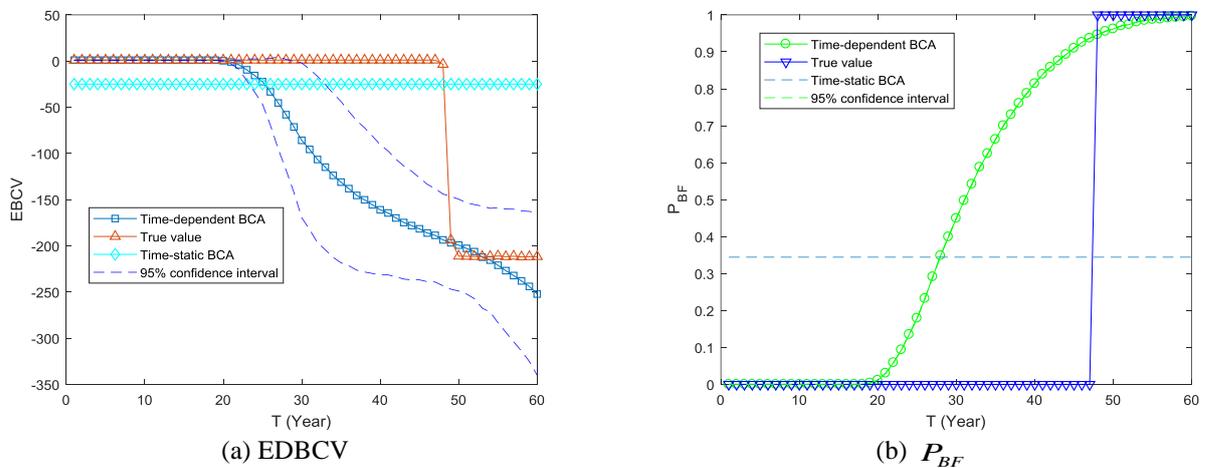
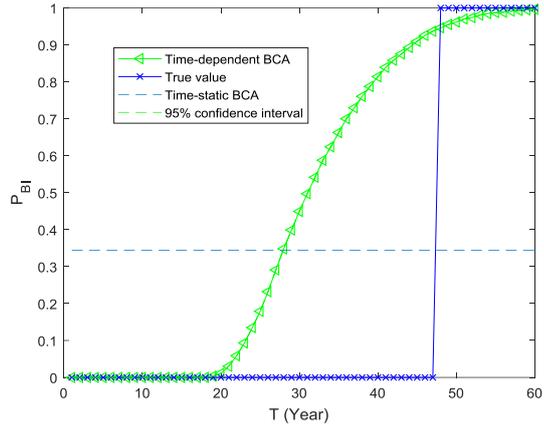


Figure 3-5. Profit trajectory at different estimation points.

3.4.4 Results

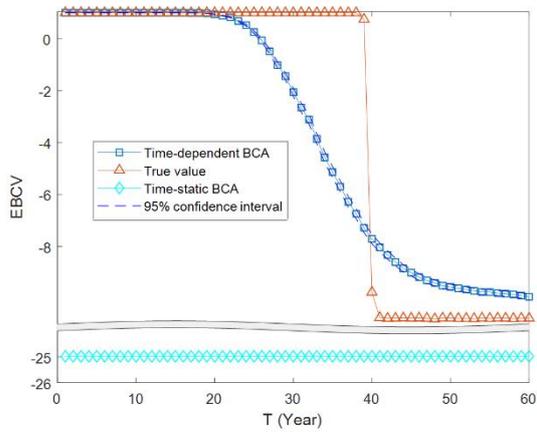
The results from the time-static and time-dependent business continuity analyses are compared in Figure 3-6, Figure 3-7 and Figure 3-8, where the true value is generated based on a theoretical model with known parameters. Abscissa axis shows the estimation horizon T , and the vertical axis stands for the different BCV indexes. Therefore, these results show the business continuity of NPPs at different age (t), if it is operated for different lengths of time (T). It can be seen from the Figures that:



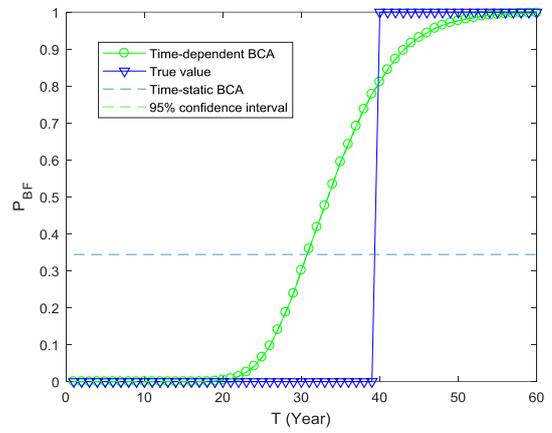


(c) P_{BI}

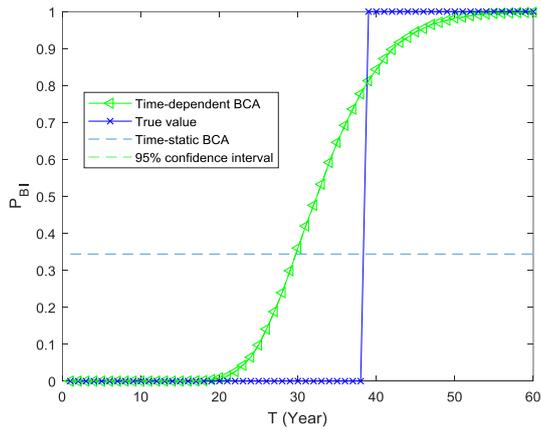
Figure 3-6. Business continuity metrics at $t=1$ year.



(a) EBCV



(b) P_{BF}



(c) P_{BI}

Figure 3-7. Business continuity metrics at $t=10$ years.

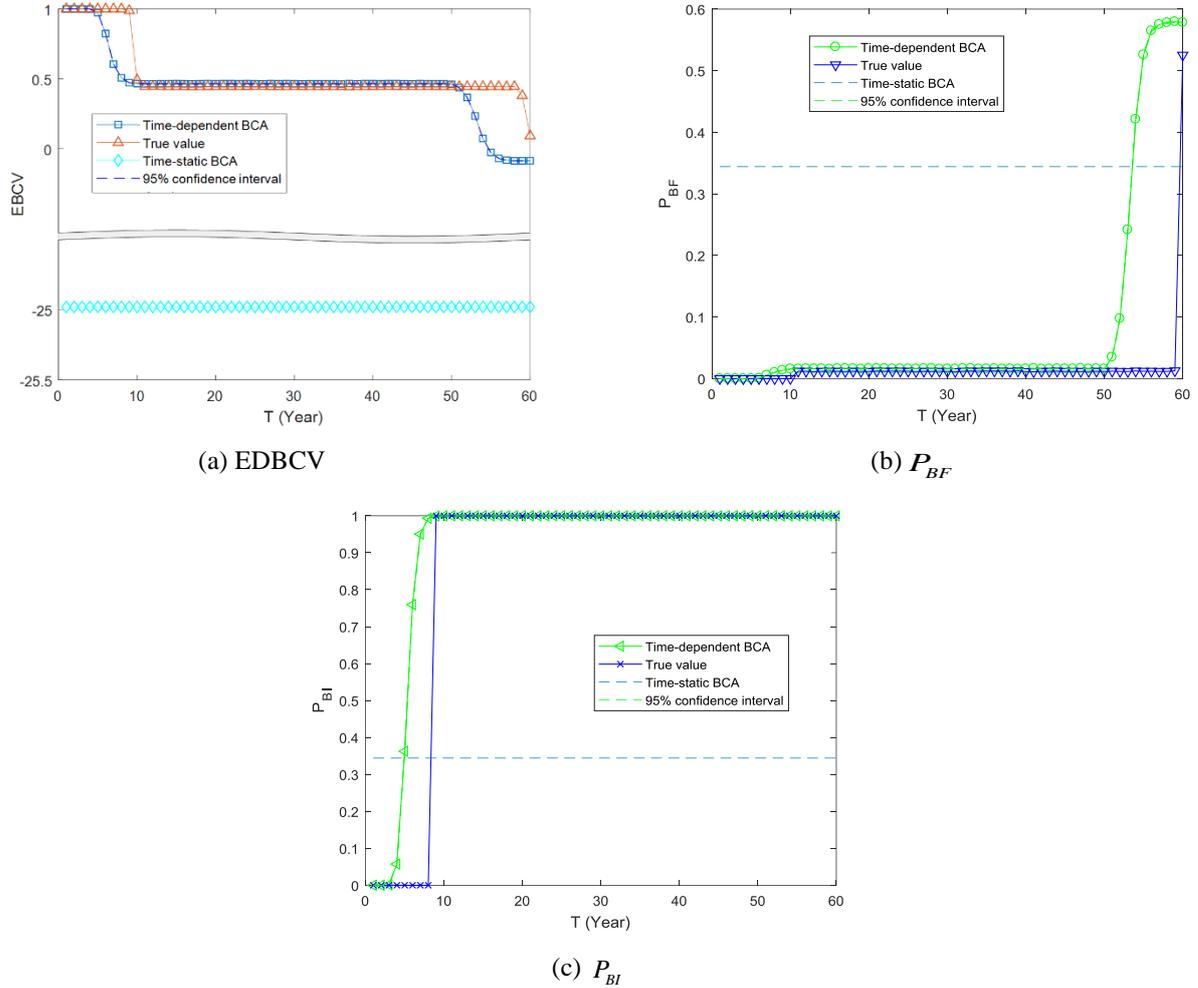


Figure 3-8. Business continuity metrics at $t=40$ years.

(1) At each t with the increase of the estimation horizon T the DBCV decreases. This means that regardless of the age t of the NPP, the longer the NPP is operated, the worse its business continuity. This is logical as it is primarily caused by the tube's degradation process. No rupture is supposed to occur at the beginning of system operation. Subsequently, as the crack grows, rupture will occur eventually and lead to system failure. In addition, the dynamic business continuity (DBC) indexes curves drop significantly after a certain value. In practice, intervention measures like overhauls need to be taken before this T , in order to prevent serious losses from occurring failures and ensure the business continuity.

(2) For the same estimation horizon T , with the increase of NPP age t the EDBC moves toward left, which means the financial safety margin is narrowing overtime t . This is because the steam generator tube is getting closer to a dangerous state as the NPP ages.

(3) The comparison between DBC and static business continuity shows that the results from the DBCA

using condition-monitoring data are closer to the true BCV than those of the static business continuity. This is because the DBC using condition monitoring data can capture the time-dependent behaviour of SGTR degradation. Moreover, with more condition monitoring data the DBCV estimation results are more accurate.

(4) Confidence interval quantifies the level of confidence that the BCV metrics are captured by the interval. From Figures Figure 3-6~Figure 3-8, we can see that with more data available, the width of confidence interval is narrowing. That is because that with more condition monitoring, more precise of the component state estimation and less uncertainty of the BCA results.

3.5 Conclusion

In this chapter, a dynamic business continuity assessment method that integrates condition monitoring data is proposed. Two factors that influence the dynamic behaviour of business continuity are considered explicitly. The first one is the dynamics of the degradation-to-failure process affecting the safety barriers. Condition monitoring data are used to update and predict the time-dependent failure behaviour by PF. The second factor is the time-dependent profit and tolerable losses. This is quantified by applying a stochastic price model and an installment model. A simulation-based framework is developed to calculate the time-dependent business continuity metrics originally introduced. A case study regarding the analysis of an accident initiated by SGTR in a NPP shows that the proposed framework allows capturing the dynamic character of business continuity. The outcomes of such dynamic analysis can provide insights to stakeholders and decision-makers, that can help them to identify when best to take actions for preventing serious losses and ensuring business continuity.

Chapter 4 Joint optimization for enhancing business continuity

In this chapter, a joint optimization model for business continuity that considers prevention, mitigation, emergency and recovery processes is proposed. Generally, the resources to guarantee a system's continuity are often limited. How to allocate and arrange the limited resources to keep the continuous service of the target system is a paramount issue. Conventionally, organizations treat the different phases in BCA separately. In order to capture the coordination of the four phases and extract useful information on resource allocation, a joint optimization model is developed. A case study of the SGTR in a NPP is conducted to illustrate the utility of the joint resource allocation model.

The remainder of this chapter is organized as follows. Section 4.1 presents a brief literature review on the optimal allocation of loss-reduction resources. Section 4.2 elaborates the joint optimization framework. Section 4.3 shows the MIGA method that is used for solving the resource allocation problem. Section 4.4 illustrates the utility of the proposed framework through a NPP case study. Finally, Section 4.5 concludes this chapter.

4.1 State of the art

Most of research for improving system reliability, safety and resilience concentrates on individual or partial stages, especially for the situation under limited resources. Some attempts focus on resource allocation for preventive stage. For example, in [106], optimization of preventive upgrading interventions on the bridges of a highway network has been conducted to improve the bridge reliability under earthquake disruption. In addition, a combination of the knapsack problem and a risk matrix has been presented to carry out a cost-benefits analysis to efficiently make prevention investment decision within a predefined budget in [107]. An optimal portfolio of prevention measures for time-dependent accident scenarios has been proposed in [108], using Dynamic Bayesian network (DBN) represent the temporal evolution of component failure. These model concentrates on risk prevention, in other words, reducing frequency of disruptive event.

In [109, 110], a multi-objective multi-decisionmaker resource allocation framework has been represented to model resource allocation process within large-scale hierarchical systems, aiming at mitigating the risks from the viewpoint of subsystem and overall system. Additionally, some studies have developed to allocate the system resources to improve system resilience for a specific disaster. For instance, multi-systems' joint restoration processes resilience modeling has been addressed and the effectiveness of five different restoration strategies has been compared in [45]. In [46], a two-stage mixed-integer programming resource allocation model for lifeline system has been proposed to improve the efficiency of restoration. A multi-objective optimization model of emergency organization allocation for sustainable disaster supply chain has been developed to design optimized strategies of emergency organization allocation [47]. with the objective of minimizing the expected outage duration of loads, multiple microgrids have been used to real-time optimize resources and restore critical loads [48]. In [49], a restoration resource allocation model has been proposed to enhance resilience of interdependent infrastructure systems. A resiliency-based optimization methodology has been performed over the set of feasible restoration policies, information investments, and human resource availability to determine optimal customer and system-wide monetary utility [50]. A stochastic optimization technique has been developed to allocate scarce national resources to coping with multiple simultaneous disasters happening across the nation [51]. All of the above-mentioned researches concern post-disruption decision making, assuming the disruption has happened.

The objective of resource allocation in business continuity is searching an integrated optimization method to improve the system continuity level, with consideration of the necessary measures in the whole stage, including pre-disruption and post-disruption. As a matter of fact, this problem has not been sufficiently addressed in the available literature and we are motivated to fill the above gaps by mathematically formulating the business continuity enhancing based resource allocation problem and developing a joint optimization approach to identify the system resource allocation on four phases. With respect to system business continuity, we adopt the quantitative metrics proposed in [14]. Regarding the solution of optimization, MIGA is applied to solve the joint optimization model due to its parallel searching and efficient interactions characteristics [111, 112].

4.2 Joint optimization

BCM starts from hazard identification, and follows with event evolution analysis, as well as safety barriers identification. The concrete definitions, functions and the characteristics of the four stages addressed in BCM are shown as follows:

- (1) Prevention stage often takes inherent safety measures to lower the probability of disruptive event [16].
- (2) Mitigation phase is usually equipped with passive strategies, aiming at minimizing consequence of disruptive events. The designer's choice of business continuity maximization on mitigation phase is captured by redundancy design, especially for the safety barriers arrangement which are used to mitigate the system consequence induced by a disruptive event [113]. Regarding a corresponding redundant system, if one component collapses, corresponding redundant system will substitute it to work for a period of time. Redundancy allocation problem is an important topic in system reliability design, and also plays a key role in engineering resilience[114-116].
- (3) Emergency phase starts after the mitigation phase and prior to the recovery phase. Corresponding emergency safety measures are activated when mitigation measures fail to contain damage. sometimes human intervention are required in this phase [14].
- (4) Recovery phase mainly focuses on restoring a system timely to normal operation following disruptive events [117]. Recovery ability refers to the ability of a system repairing itself [118]. The cost of this phase mainly focuses on system investment on repair crews, vehicles, equipment and replacement components [119].

The overall cost (based on the cost of deployment of safety barriers on prevention phase, mitigation, emergency and recovery phase) must not exceed a budget constrain. Considering the system business continuity, one paramount objective of BCM is maximizing BC given limited budget or resources.

As reviewed in Section 3.2, the metric of EBCV directly reflect business continuity level which can be used as an objective for the optimizing system resource allocation.

A joint optimization framework considering all safety barriers and corresponding cost in four phases is presented in Figure 1-1. Prevention cost, redundancy arrangement and recovery investment are used to minimize the loss level in BCM [19, 120].

$$\max \text{EBCV} = f(C_{SB_P}, C_{SB_M}, C_{SB_E}, C_{SB_R}) \quad (4.1)$$

$$\text{s.t. } C_{SB_P} + C_{SB_M} + C_{SB_E} + C_{SB_R} \leq C_{total} \quad (4.2)$$

$$C_{SB_P} \geq 0, C_{SB_M} \geq 0, C_{SB_E} \geq 0, C_{SB_R} \geq 0. \quad (4.3)$$

The objective function maximizes system business continuity under SGTR over T time periods. where C_{SB_P} , C_{SB_M} , C_{SB_E} , C_{SB_R} , are the cost allocated in preventive, mitigation, emergency and recovery stage, respectively. C_{total} denotes total resource budget. The first constraint in the joint optimization model is the constrain on the maximal allowable resource budget. With limited budget, it is essential to allocate resource budget in an effective way, in order to maximize system business continuity.

4.3 Solution method

In this section, the solution of this joint optimization issue is presented. The joint optimization model shown by Equation (4.1) can be solved with methods such as Lagrange multiplier. However, due to the computation complexity of parameters and the nonlinear characteristic of the function, hereby, MIGA is used to solving the model. MIGA is a powerful stochastic search algorithm that has been successfully used in literature for solving optimization problems in critical infrastructure resilience [45]. The procedures to search for an optimal solution to the joint optimization problem can be described by following steps [121, 122].

- (1) Encoding. Express each solution of cost allocation by a genotype $e = (e_1, e_2, \dots, e_7)^T$. The different combinations of structured cost make up different chromosomes. The initial solution is randomly generated according to constraint conditions in Equations (4.2) and (4.3).
- (2) Fitness assessment. Calculate the fitness value of each genotype. The fitness value of each genotype represents the business continuity value of the system of interests [123].
- (3) Selection, crossover and mutation. Offsprings are produced by these three types of operator, and then return to the second step until the maximum generation is reached. The section operator chooses a genotype with a probability depending on its fitness value. Two selected genotypes produce two descendants by using crossover operator that exchange substrings of the codes of the two chosen genotypes. Then, each descendant generates an offspring by using the mutation operator with a mutation

probability. The genotype in the final generation with the maximum EBCV corresponds to the optimum resource allocation.

With abovementioned steps, the genotype is selected according to their fitness value in each generation. The rule for stop is the convergence of the optimal fitness value between two generations. When the algorithm stops, the genotype corresponding to the minimal fitness value is the optimal solution for the budget allocation model.

4.4 Case study

A Zion PWR NPP is considered to illustrate the developed model. It is assumed that the NPP has one SG that is equipped with a bundle of 3592 inverted U tubes. Each U tube has a mean outside diameter of 22.23 mm and a mean wall thickness of 1.27 mm and is subject to SGTR caused by diverse degradation mechanisms like stress corrosion cracking (SCC), fatigue, pitting corrosion and fretting wear. A detail list of the NPP parameter values is presented in

Table 4-1 Parameters of the NPP.

Parameter	Value
Capacity of NPP (C)	1100 Mwh
Number of tubes (n_{tube})	3592
Outer diameter (d)	$N(22.23, 0.1667)$ mm
Pressure different (ΔP)	$N(8.3, 0.33)$ Mpa
Thickness (b)	$N(1.27, 0.0592)$ mm

4.4.1 Event modelling

For illustrative purposes, it is assumed that the NPP is only subject to the threat of one disruptive event, SGTR. SGTR is a potential accident that is induced by the degradation of tubes in steam generator, which can further lead to tube cracking and rupture event. In principal, steam generator tube is designed for transferring heat produced by steam generator to drive turbine for producing electricity [124]. To analyze business continuity of the NPP considering SGTR, a schematic event tree (ET) on SGTR is investigated, as shown in Figure 4-1.

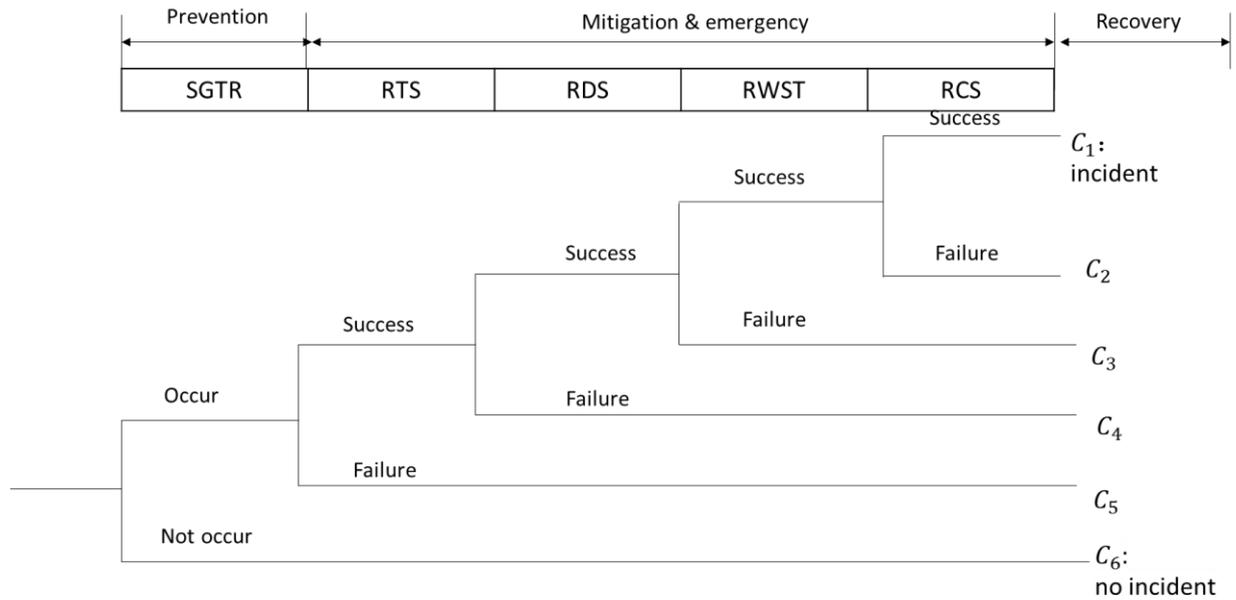


Figure 4-1. Schematic ET model of SGTR accident ($C_2 \sim C_5$ core damage) [120].

Depending on the performance of different safety barriers, six consequences, i.e., C_1, C_2, \dots, C_6 can result from the SGTR event. These consequences can be grouped into three categories based on their severity: no incident C_1 , incident C_{II} and core damage C_{III} as tabulated in Table 4-2. Different consequences are formulated based on the variant performances of safety barrier.

Table 4-2. Classification of consequences.

Consequences	Group	Meaning
C_6	C_I	No SGTR occurs, the NPP is operating normally.
C_1	C_{II}	SGTR occurs, but the consequence is successfully controlled by the mitigation and emergency barriers. The power generation business is temporarily terminated.
$C_2 \sim C_5$	C_{III}	Core damage is caused by SGTR, the power generation business is terminated for a long time.

4.4.2 Business continuity modelling

The aim of this section is determining the different measures on the four phases and their functions on NPP business continuity, as well as the corresponding costs.

4.4.2.1 Prevention phase

The main prevention measures for crack growth are regular inspection on the crack size, and, also the timely preventive maintenance, such as plugging the cracked tube. Additionally, the rupture probability depends on the inspection interval and plugging threshold, as formulated in:

$$P_{rupt,tube} = f(x, y_{th}). \quad (4.4)$$

The costs in this preventive stage include inspection cost and plugging cost. In practice, the inspection is conducted regularly during the NPP refueling process. The tube crack size usually is measured by eddy current test and once the crack reaches a given threshold y_{th} , the associated tube is plugged to prevent occurrence of tube rupture [24].

A two-stage crack model is used to simulate the crack progression. For the first stage, it is assumed that from the initial crack to the critical length of 0.1 mm, which indicates after this critical length the crack propagates faster. The duration of first stage is described through a lognormal distribution [24]. The propagation stage can be formulated through a scott model [95], which is an empirical model that illustrate the crack growth rate as a function of stress:

$$\frac{da}{dt} = \alpha(K - K_{th})^m, \quad (4.5)$$

$$K = F\sigma\sqrt{\frac{\pi a}{2}}, \quad (4.6)$$

$$\sigma = \frac{\Delta P \cdot d}{2b} \quad (4.7)$$

where $\frac{da}{dt}$ is the crack growth rate, a is the crack length, α , K_{th} and m are constant parameters related to the component material properties, σ is the stress at the crack tip, ΔP denotes the pressure difference. In this case study, Alloy 600 material is considered for the steam generator tube. Based on the material properties, the values for parameters in Equations (4.5)-(4.7) can be determined.

Therefore, the cost in prevention phase can be formulated as:

$$\begin{aligned} C_{SB_p} &= C_{insp} \cdot n_{insp} + n_{tube} \cdot P_{plug} \cdot C_{plug} \\ &= \frac{T}{x} \cdot C_{insp} + n_{tube} \cdot P_{plug} \cdot C_{plug}, \end{aligned} \quad (4.8)$$

where T denotes the time horizon of business continuity assessment; C_{insp} is the cost of a single inspection of the health of the tube; C_{plug} is the unit price for plugging one tube; and p_{plug} is the plugging rate of the tube, which is also dependent on the values of x and y_{th} , can be further calculated through simulation.

4.4.2.2 Mitigation & emergency phase

Failures of safety barriers in mitigation phase can increase accident severity. In this regard, redundancy system or component are often considered to improve reliability and availability of these component. Evidently, as the number of redundancy augments, the cost of the whole system proportionally grows. In this context, achieving an optimal number of redundancy components by which total costs of the system is minimized could be interesting.

From the event tree model in Figure 4-1, the losses generated from different consequences $C_1 \sim C_6$ can be quantified as a function of the event probabilities along the sequences:

$$p_{C_i} = f_{ET}(p_{SGTR}, p_1, p_2, p_3, p_4) \quad (4.9)$$

where p_{SGTR} is the probability of a single tube rupture, p_1, p_2, p_3, p_4 represent the failure probability of RTS, RDS, RWST and RCS, respectively.

The performance of mitigation & emergency measures, i.e., RTS, RDS, RWST and RCS in Figure 4-1, can be represented by their failure probabilities. Redundancy design can be an appropriate way used for reducing failure probabilities. In this chapter, we assume that parallel redundancy using the same type of equipment is considered for the four mitigation and emergency safety barriers. It is easy to show that the failure probability of the i -th measure becomes:

$$p_i = (p_{i,b})^{n_i+1} \quad (4.10)$$

where $p_{i,b}$ is the failure probability of the i -th safety barrier system and n_i is the number of redundant system added to the original system.

The cost for improving mitigation and emergency performance can, then, calculated by:

$$C_{SB_M} = \sum_{i=1}^4 C_{R,i} \cdot n_i \quad (4.11)$$

where $C_{R,i}$ is the price for adding one i -th redundancy measure.

4.4.2.3 Recovery phase

Afterwards, the analysis goes to the recovery phase. That means when consequence C_B and C_C occur, the NPP becomes temporarily unavailable for producing electricity, until the recovery measures are applied to the system to restore the system to normal operation. It is assumed that the basic recovery time T_{bs,C_i} , where $i = I, II, III$ follow lognormal distributions [99, 101] whose probability density function is:

$$f(T_{recv,C_i}) = \begin{cases} \frac{1}{\sqrt{2\pi}\beta_i T_{recv,C_i}} e^{-\frac{(\ln(T_{recv,C_i}) - \varepsilon_i)^2}{2\beta_i^2}}, & T_{recv,C_i} > 0 \\ 0, & T_{recv,C_i} \leq 0 \end{cases} \quad (4.12)$$

where $f(\cdot)$ is the probability density function (PDF) of T_{recv,C_i} ; ε_i and β_i are the mean value and standard deviation value of the lognormal distribution, respectively. The values of ε_i and β_i are depends on the recovery ability of the target organization. It is noting that the more serious consequence i , the smaller of $c_{e,i}$.

In practice, given the budget on BCM, with more resources allocated on recovery process, the recovery rate will be improved, and, then, lead to more efficient recovery process to reduce the potential indirect losses in system operation. The allocated resources on the recovery process is often assumed following logarithmic function [125, 126], which is defined as Equation (4.13),

$$T_{recv,i} = \frac{T_{recv,C_i}}{1 + \ln(1 + c_e \cdot \mu \cdot C_{SB_R})}. \quad (4.13)$$

where $T_{recv,i}$ is a random variable that represents the time needed to recover from the i -th consequence; T_{recv,C_i} denotes the basic recovery time for consequence C_i , which is dependent on the basic requirement on recovery time, C_{SB_R} denotes the resources invested on the recovery process and c_e is the effective parameters of resources on II -th consequence; μ denotes the relationship between different cost-effective parameters for different consequences; Its value should be set by decision makers based on the capability of the organization.

4.4.3 Joint optimization

To formulate the objective EBCV, the indirect losses caused in the recovery process can, be modelled by:

$$L_{In,C_i} = P_e \cdot C \cdot T_{recv,C_i}, \quad (4.14)$$

where L_{in,C_i} represents the indirect losses in the recovery process for consequence C_i ; P_e is the unit electricity price; C is the generation capacity of the NPP; T_{recv,C_i} denotes the recovery time for the i – th consequence.

Then, the EBCV in Equation (4.1) can be formulated by:

$$\begin{aligned}
E(L([0,T])) &= \sum_{i=1}^3 E(L_{C_i}([0,T])) \cdot P_{C_i} \\
&= \sum_{i=1}^3 E(L_{d,C_i} + L_{in,C_i}) \cdot P_{C_i} \\
&= P_{C_{II}} \cdot (L_{d,C_{II}} + P_e \cdot C \cdot E(T_{recv,C_{II}})) + P_{C_{III}} \cdot (L_{d,C_{III}} + P_e \cdot C \cdot E(T_{recv,C_{III}})).
\end{aligned} \tag{4.15}$$

According to the cost analysis in whole process, the explicit form of the resource allocation model is given as:

$$\max \text{ EBCV} = f(x, y_{th}, n_1, n_2, n_3, n_4, C_{SB_R}) \tag{4.16}$$

$$\text{s.t. } C_{SB_P} + C_{SB_M} + C_{SB_R} \leq C_{Th}, \tag{4.17}$$

$$C_{SB_P}, C_{SB_M}, C_{SB_R} \geq 0, \tag{4.18}$$

$$x \in [6, 7, 8, \dots, 17, 18] \tag{4.19}$$

$$P_{plug} \leq P_{th}, \tag{4.20}$$

$$\begin{aligned}
n_{L,1} &\leq n_1 \leq n_{U,1}, n_1 \in N \\
n_{L,2} &\leq n_2 \leq n_{U,2}, n_2 \in N \\
n_{L,3} &\leq n_3 \leq n_{U,3}, n_3 \in N \\
n_{L,4} &\leq n_4 \leq n_{U,4}, n_4 \in N
\end{aligned} \tag{4.21}$$

The first constrain in Equation (4.17) regards the total budget on all safety barriers which cannot exceeds a limited value C_{th} . In Equation (4.17), the cost $C_{SB_P}, C_{SB_M}, C_{SB_R}$ are further calculated by Equations (4.8), (4.11) and (4.13). The constrain in Equation (4.19) defines the possible value of inspection interval (in months). In this work, it is assumed that the inspection can reveal the exact state of the tube. The constraint in Equation (4.20) means the total number of plugged tubes can exceed a maximum value. The value of p_{th} is determined based on the power generation efficiency requirement of the NPP. According to the nuclear regulations, a steam generator of the type employed in Zion PWR NPP can tolerate up to 30% plugged tubes before a significant reduction in efficiency occurs [127]. Therefore, here, we see $p_{th} = 0.3$. The last constraint in Equation (4.21) describes the minimal and maximal number of redundant system for the mitigation measures and is employed to describe the redundancy number of mitigation measures.

MIGA is applied to solve the previously defined joint optimization problem. And the flowchart of MIGA is shown in Figure 4-2. The parameters of the algorithm are tabulated in Table 4-3.

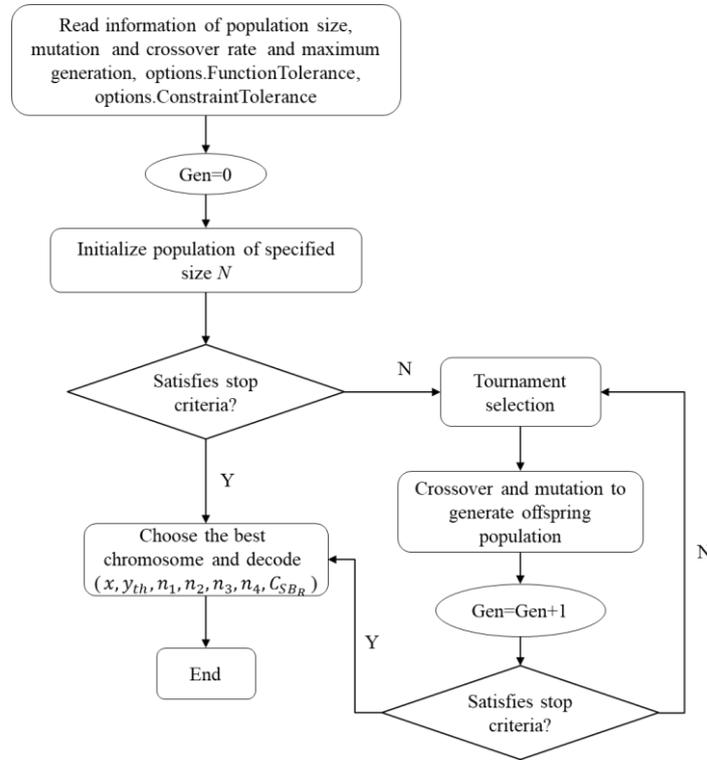


Figure 4-2 Schematic of MIGA.

Table 4-3. Parameters of the MIGA algorithm.

Parameters	Values
Population size	40
Crossover rate	0.9
Mutation rate	0.5
Maximum generation	500

4.4.4 Results and sensitivity analysis

In this section, we firstly present comparative results on single phase resource allocation and joint optimization within predefined budget. Subsequently, sensitivity analysis on budget, cost-effectiveness parameter as well as failure probability of mitigation barriers are carried out to discuss the influence of these variables on EBCV.

4.4.4.1 Results

The optimization problem is solved numerically through the MIGA described in Section 4.3 to show insights on resource allocation under budget constraint. Parameter values used in this case study are tabulated in Table 4-4.

Through simulation on crack onset and propagation, we can obtain the probability that the crack onset time outside of T is 0.6498.

Table 4-4 Parameter values used in the case study.

Parameter	Meaning	Value
C	Capacity of NPP	1100 (Mw)
L_{tol}	Tolerable losses	5×10^6 (k€)
C_{insp}	Cost for one inspection	500 (k€)
C_1	Cost for adding one redundant RTS	20 (k€)
C_2	Cost for adding one redundant RDS	4 (k€)
C_4	Cost for adding one redundant RWST	11 (k€)
C_4	Cost for adding one redundant RCS	5 (k€)
C_{plug}	Cost for plugging one tube	5 (k€)
c_e	Cost effectiveness parameter for consequence C_{II}	0.001
μ	The relationship of cost-effectiveness parameter between C_{II} and C_{III}	0.5
$T_{recv,1}$	Basic recovery time for consequence C_{II}	Lognormal (3.9828, 0.4724) (days)
$T_{recv,2}$	Basic recovery time for consequence C_{III}	Lognormal (6.5922, 0.4724) (days)
P_e	Unit price of electricity	50€/MWh
$n_{L,1}, n_{L,2}, n_{L,3}, n_{L,4}$	Lower bound of mitigation measures' number	0
$n_{U,1}, n_{U,2}, n_{U,3}, n_{U,4}$	Upper bound of mitigation measures' number	4
C_{total}	Total budget	8000 (k€)

To investigate the effectiveness of the joint optimization method based on business continuity, a comparison among individual optimization of the three phase is conducted. We program the joint optimization model 10 times and consider all the results of the objective during the 10 times, and the results are shown in Figure 4-3, where green circle means the EBCV calculated from spend all the resources on prevention phase and boxplot is the 10 times simulation results for the proposed model. We can see that the joint optimization works better targeting maximizing EBCV than the strategy only investing all the budget on preventive stage.

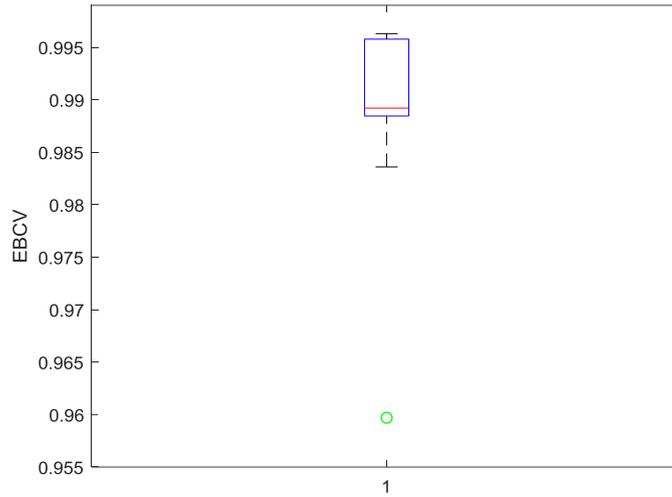


Figure 4-3 Results comparison for proposed joint optimization and prevention only ($c_{total} = 8000(\text{k€})$).

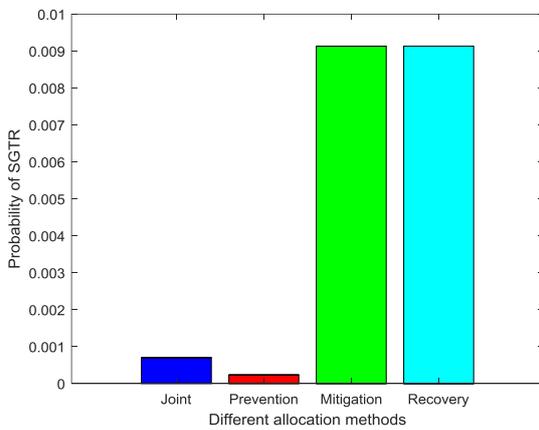
The details on one of the simulation results is shown in Table 4-5. It is worth noting that a tolerance level $\varepsilon = 1.0 \times 10^{-4}$ is enforced for MIGA. It can be seen that the joint optimal design solution requires to do a periodical inspection of the steam generation tube every 15 months and the tube will be plugged when crack length exceeds 7.8780(mm). Additionally, redundancy design for mitigation & emergency measures is $n_1 = 1, n_2 = 4, n_3 = 4, n_4 = 4$. For the recovery stage, additional resources $C_{SB_R} = 108.2384$ (k€) are allocated to improve the recovery efficiency and reduce the indirect losses L_{in} .

Table 4-5 Comparison results for business continuity under different strategies.

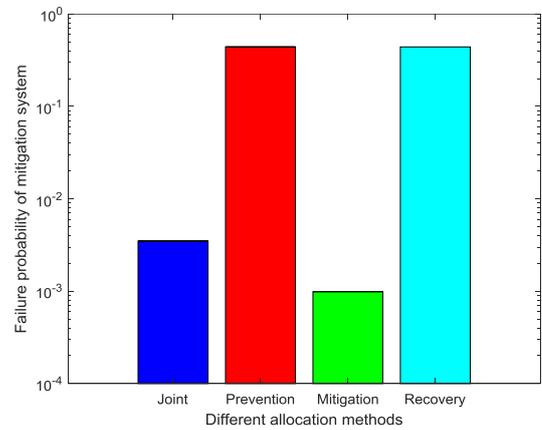
Variable	Joint optimization	Prevention measures only	Mitigation & emergency	Recovery (only)
x	1.25 (year)	1 (year)	/	/
y_{th}	7.8780 (mm)	14.1347(mm)	/	/
n_1	1	/	4	0
n_2	4	/	4	0
n_3	4	/	4	0
n_4	4	/	4	0
C_{SB_p}	7792.2 (k€)	7990.05(k€)	0	0
C_{SB_M}	90 (k€)	0	120 (k€)	0

C_{SB_R}	108.2384 (k€)	0	0	8000 (k€)
$C_{total, cost}$	7990.4 (k€)	7999.05 (k€)	120 (k€)	8000 (k€)
EBCV	0.9963	0.9597	-20.06	-44.63

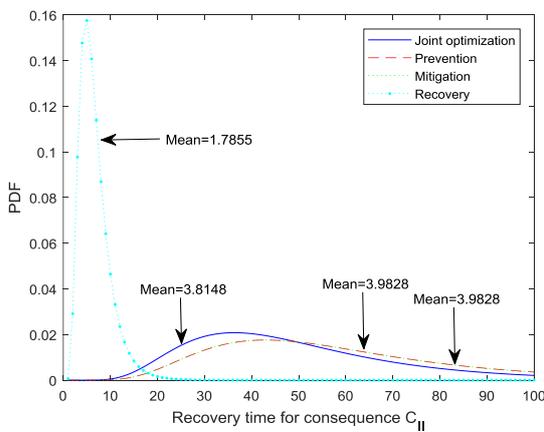
We pick the best solution among the 10 trails to show the details of the simulation result. The behavioral indexes proposed Section 4.2 are shown in Figure 4-4. We can see that the probability of SGTR dramatically decrease after jointly optimizing resource allocation comparing with only spend the budget on mitigation and recovery phase. For the mitigation & emergency phase, the best option is spent all the budget only for mitigation phase. Regarding the recovery phase, as expected, if all the budget is spent on recovery phase, the recovery time will be significantly reduced (both regarding consequences C_{II} and C_{III}).



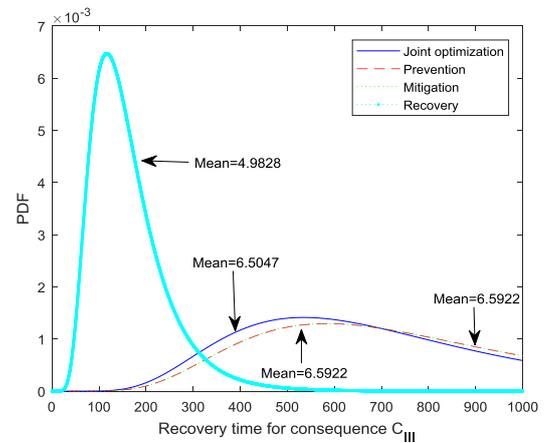
(a) Probability of SGTR.



(b) Failure probability of mitigation system.



(c) PDF of recovery time for consequence C_{II} .



(d) PDF of recovery time for consequence C_{III} .

Figure 4-4 Behavioural indexes in prevention, mitigation & emergency, and recovery phases.

It can be seen from Figure 4-4 that if all the budget C_{total} is spent on the preventive phase, the SGTR occurrence probability can be reduced from 6.9998×10^{-4} to 2.9998×10^{-4} comparing with the joint optimization model. However, the respective EBCV is lower than the joint optimization model. This is because in the preventive phase only considered model the total cost is invested to prevent crack growth. The mitigation & emergency and recovery processes are ignored, becoming bottlenecks to the business continuity of the NPP. Similar results can be found in Figure 4-4 (b), (c) and (d): although the solution obtained from the joint optimization model not be optimal with respect to each safety barrier, it can achieve an overall optimal performance with respect to business continuity. That is because that the proposed joint optimization method takes all the factors into account, thus, there is not any special shortcoming/bottleneck in the resource allocation process.

4.4.4.2 Sensitivity analysis

In this context, one major engineering interests lie in the quantification of the sensitivity of the business continuity of the system with respect to the different parameters of basic variables. Because sensitivity analysis on parameters provides insights on how these parameters affect the optimal allocation of budgets. In this work, a sensitivity analysis is conducted in terms of system total budget, failure probability of mitigation measures and the cost-effectiveness parameter in recovery stage. For each parameter sensitivity investigated, the other parameter values are kept the same. The problem size is limited due to the significant computations required by the optimization models to obtain joint optimal solutions [128].

For the changing budget's effect on system business continuity, the result is presented in Figure 4-5. As expected, with the growth of total budget, corresponding EBCV increases. This is due to more available resources allocated to keep system business continuity, the higher preventive ability, mitigation & emergency ability and recovery ability. Subsequently, the less loss in the evaluated time horizon $[0, T]$, and eventually, the higher business continuity. Additionally, when the total budget is bigger than 8000k€, the changing on EBCV is small, indicating marginal degradation of the budget. This is mainly because the limitations on the plugging rate, mitigation measures redundancy. This result can provide insights on how many budgets should be arranged to keep system business continuity.

Through the comparison results showed in Figure 4-5, we find that when total budget is relative less, the value of cost-effectiveness has more significant influence on system business continuity. This is mainly probably caused

by the fact when the budget is small, the cost-effectiveness parameter c_e plays a more important role in budget allocation. Higher cost-effectiveness makes budget allocated on recovery stage more effective, and then, effectively reduce the system indirect losses.

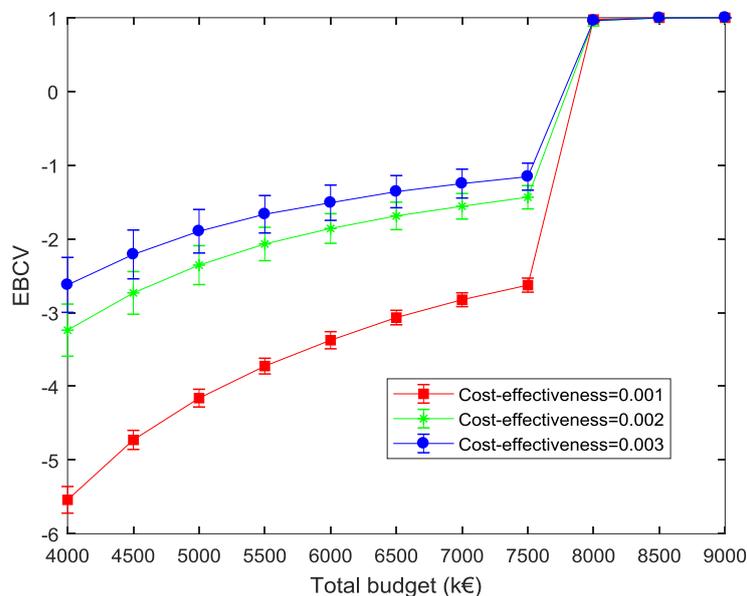


Figure 4-5 Comparison of EBCV with different cost effectiveness parameters.

To investigate the influence of cost-effective parameter on the joint optimization model, a comparison among total budget with 7500k€ and 8000k€ are studied respectively. Figure 4-6 shows the results of the EBVC comparison as a function of the changing cost-effective parameter. As can be seen from Figure 4-6, the smaller the budget, the more sensitive of cost-effectiveness parameters on business continuity, which is also verified in Figure 4-5. Moreover, when the cost-effectiveness parameter increases from 0.001 to 0.006, the corresponding EBCV increase. Additionally, when the cost-effectiveness parameter increases from 0.006 to 0.01, the change in system EBCV is relatively small due to the marginal decreasing rate of EBCV with increase of cost-effectiveness.

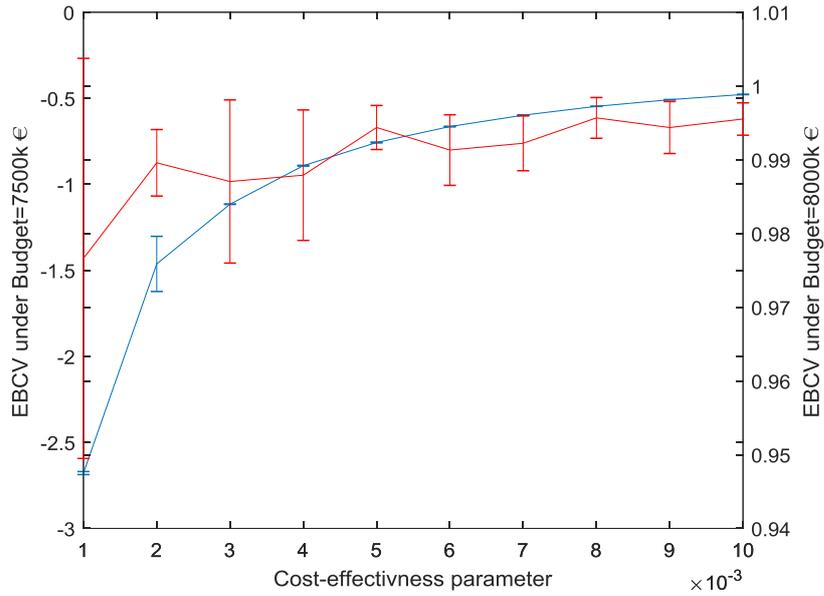


Figure 4-6 Comparison EBCV with different cost-budget.

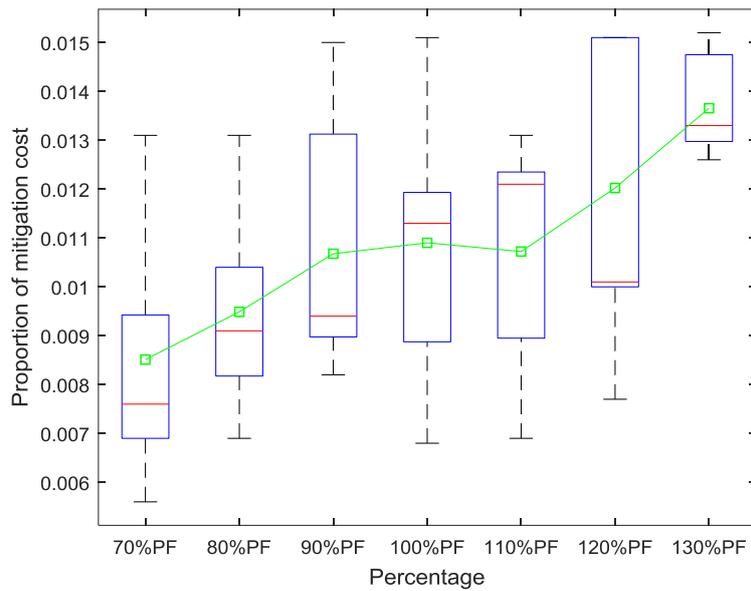


Figure 4-7 Schematic of changing failure probability of mitigation measures (70%~130%) under budget $C_{total} = 8000k€$.

To test the influences of mitigation measures' failure probability on the budget allocation, we do a sensitivity analysis by changing mitigation failure probability from 70% PF to 130% PF (where $PF = [p_{RTS}, p_{RDS}, p_{RWST}, p_{RCS}]$). As shown in Figure 4-7, with increase of failure probability of mitigation measure,

the proportion of mitigation cost increases accordingly. This is a strait forward conclusion: more redundant safety systems are needed if the safety system has higher failure probability.

4.5 Conclusion

A mathematical model is formulated in this chapter to jointly optimize system limited resources for enhancing system business continuity. The joint optimization model is based on the proposed BCA metrics, which aims at calculating system business continuity level given an estimation horizon. The proposed joint model considers reducing system potential loss under a given disruptive event from a comprehensive viewpoint. The case study on a NPP against SGTR event demonstrates the utility of the model in decision making. Through this case study, we can see that the proposed joint optimization model works better than the other models that deal with four phases individually. Through the results of sensitivity analysis, we can infer that: (1) larger resource budget can result in higher business continuity; the change of EBCV is marginally decreasing with the increase of the budget; there is an optimal budget for the given NPP; (2) higher failure probability of the safety measures in mitigation phase, less redundancy is needed; (3) the smaller the budget, the more sensitive of cost-effectiveness parameters on business continuity. The optimization method can jointly provide a better scheme than separative optimization strategies for decision makers under limited budget or resources.

Chapter 5 Conclusion and future work

5.1 Conclusion

This dissertation aims at developing an integrated framework and computational tools for the assessment and optimization of system business continuity. More particularly, the works in this dissertation can be summarized as follows.

Firstly, a framework for DRA was developed to integrate condition monitoring data and inspection data. A HM-GMM was developed to estimate the degradation states of the safety barriers based on the condition monitoring data. The estimated degradation states were integrated with inspection data for DRA by a BN model. An application showed that integrating two data sources into the DRA gives more robust results than using the two data sources individually.

Secondly, a simulation-based DBCA method was developed to analyse system business continuity that allows considering the time-dependent feature of safety barriers' states and target system revenues. A PF model was used to predict the RUL of the safety barriers from condition monitoring data. The time-dependent revenue was modelled by an instalment model. The proposed DBCA framework was applied to a NPP, taking into account a SGTR event. The results of the case study showed that the proposed framework allows capturing the dynamic behaviour of business continuity and can aid decision-makers.

Thirdly, a mathematical model was formulated to jointly optimize the system limited resources for enhancing business continuity. The model aims at reducing the system potential loss from a comprehensive viewpoint in which prevention, mitigation, emergency, and recovery phases are considered jointly. MIGA was employed to obtain the optimal solution of the comprehensive model. A comparative study was carried out to verify the effectiveness of the proposed BCA based decision-making. A sensitivity analysis was done on the cost-effectiveness parameter, budget, failure probability of mitigation measures.

In summary, the findings of this work demonstrate the feasibility and the importance of the developed methods for risk-informed analysis and BCM of energy systems, taking into account different available knowledge, information, and data. More specifically, the original contributions of this thesis include: (1) the developed data-

integrated method in DRA can achieve a more robust DRA results than using the two data sources individually, which is all-important for safety critical system; (2) the proposed quantitative DBA method taking into account the time-dependent factors in BCM provides a robust indicator on when to do maintenance, overhaul etc; (3) the four stages in accident evolution process are integrally considered in the resource allocation process which can provide a better performance in BCM to energy system.

5.2 Perspectives

Some limitations still exist on the methods developed in this thesis, which deserve potential future work.

Firstly, an ET is applied in our study (Chapter 2) for modelling the disruptive event evolution process, and describing the protection, mitigation, emergency and recovery phases. ET is mainly a static method, which cannot capture the time-dependent of the behaviour of the safety barriers. Advanced modelling method, such as BN, dynamic fault tree, can be applied in the business continuity modelling framework.

Secondly, the proposed BCM framework (Chapter 3) only considers one kind of disruptive event. In practice, the increasing number of hazards is forcing organizations to build BCM against numerous types of disruptions and their consequences [40, 119]. Therefore, multi-event based BCM can be investigated in future work, where business continuity under multiple hazards and multi-objective optimization considering multiple disruptions can be extensively investigated.

Lastly, in the modelling process of DBA, the current method considers a discrete-time discrete state Markov model as the degradation model. A potential future work might be to extend the developed framework to other degradation models, e.g. the Brownian motion model [129], Gamma process model [130], etc. Moreover, in the current framework, the parameters of HM-GMM are estimated offline; in the future, online updating of the parameters can be considered, aiming to improve the accuracy of the DRA.

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Paper I

J. Xing, Z. Zeng, E. Zio. A framework for dynamic risk assessment with condition monitoring data and inspection data. *Reliability Engineering and System Safety* (2019), 191, 106552.

A framework for dynamic risk assessment with condition monitoring data and inspection data

Jinduo Xing ¹, Zhiguo Zeng ¹, Enrico Zio ^{1,2,3}

¹Chair System Science and the Energy Challenge, Fondation Electricité de France (EDF), CentraleSupélec,
Université Paris Saclay, Gif-sur-Yvette, France

²Energy Department, Politecnico di Milano, Milan, Italy

³Department of Nuclear Engineering, College of Engineering, Kyung Hee University, Republic of Korea
jinduo.xing@centralesupelec.fr, zhiguo.zeng@centralesupelec.fr, enrico.zio@ecp.fr

Abstract

In this paper, a framework is proposed for integrating condition monitoring and inspection data in Dynamic Risk Assessment (DRA). Condition monitoring data are online-collected by sensors and indirectly relate to component degradation; inspection data are recorded in physical inspections that directly measure the component degradation. A Hidden Markov Gaussian Mixture Model (HM-GMM) is developed for modelling the condition monitoring data and a Bayesian network (BN) is developed to integrate the two data sources for DRA. Risk updating and prediction are exemplified on an Event Tree (ET) risk assessment model. A numerical case study and a real-world application on a Nuclear Power Plant (NPP) are performed to demonstrate the application of the proposed framework.

Keywords

Dynamic risk assessment (DRA), Condition monitoring data, Inspection data, Hidden Markov Gaussian Mixture model (HM-GMM), Bayesian network (BN), Probabilistic Risk Assessment (PRA), Prognostic and Health Management (PHM), Event Tree (ET), Nuclear Power Plant (NPP).

Acronyms

ATWS	Anticipated Transient Without Scram
BN	Bayesian Network
DRA	Dynamic Risk Assessment
EM	Expectation Maximization
ETA	Event Tree Analysis
FTA	Fault Tree Analysis
HM-GMM	Hidden Markov Gaussian Mixture Model
IE	Initiating Event
NPP	Nuclear Power Plant
PF	Particle Filtering

Notation

A	Transition probability matrix
π	Initial state distribution of the Markov degradation process
$b_i(\mathbf{x})$	Probability distribution of the degradation indicator \mathbf{x} when the degradation state is S_i
C_i	The i -th consequence in the ET
$c_i(t_k)$	Condition monitoring data from the i -th safety barrier at $t = t_k$
$\mathbf{c}_{Tr}^{(k)}(t)$	Condition monitoring data from the k -th training sample at t
$d(\cdot)$	Euclidean distance
$f_{ET}(\cdot)$	ET model
K	Number of safety barriers with time-dependent failure probabilities
M	Number of safety barriers in a system
N	Number of consequences in the ET
$n_{feature}$	Number of features extracted from condition monitoring data
n_{Tr}	Number of samples in the training data set
P_{C_i}	Probability that consequence i occurs, given that the initiating event has occurred

$P_{CM,t_k}(S_{CM})$	Posterior distribution of the estimated degradation state from condition monitoring data, evaluated at t_k
$P_{INT,t_k}(S)$	Posterior distribution of the estimated degradation state by integrating condition monitoring data and inspection data, evaluated at t_k
Q	Number of health states
R_{IN}	Reliability of the inspection
R_{SB_M}	Reliability of the M -th safety barrier
S_{CM}	Estimated degradation state from condition monitoring data
$S_{CM,MAP}$	Most likely degradation state given the condition monitoring data
S_{IN}	Estimated degradation state from inspection data
S	True degradation state
t_{Tr}	Length of the observation period for the training samples
W	Working set that contains all the working states
$\mathbf{x}_{Tr}^{(k)}(t)$	Health indicator of k -th training data at t
$\mathbf{x}(t)$	Health indicator of safety barrier at t
$\boldsymbol{\mu}$	Vector of the mean values of the multivariate Gaussian distribution
$\boldsymbol{\Sigma}$	Covariance matrices of the multivariate Gaussian distribution
$\alpha_t(S_i)$	Forward variable
$\beta_t(S_i)$	Backward variable

1. Introduction

Probabilistic Risk Assessment (PRA) is widely applied to critical systems like space shuttles, nuclear power plants, etc [1]. Traditional PRA methods, like Event Tree Analysis (ETA) and Fault Tree Analysis (FTA), assume that the failure probabilities of the safety barriers are independent on time and their values are estimated based on statistical data [2]. However, in practice, the safety barriers undergo degradation processes like wear [3], fatigue [4], crack growth [5], etc., which increase their failure probabilities with time. Furthermore, the operational and environmental conditions of the system change with time and can also lead to time-dependent failure probabilities of the safety barriers [6, 7].

Safety barriers are the physical and/or non-physical means installed in the system of interest, aiming to prevent, control, or mitigate undesired events or accidents [8]. Examples are, a sprinkler system in a chemical plant [9], a reactor trip system in a nuclear power plant (NPP) [10]. To account for the time-dependent failure behavior of safety barriers, Dynamic Risk Assessment (DRA) frameworks have been developed, which use data and information collected during the system life to update the estimated risk indexes [11]. Bayesian theory has been used to update the probabilities of the events in an ET [12, 13]. Near miss and precursor data have been exploited in a hierarchical Bayesian model of DRA for the offshore industry [14, 15]. A real-time DRA has been performed in [16, 17], based on a dynamic loss function that considers multiple key state variables in the process industry. In [18], BN and Bow-tie model have been employed for the dynamic safety assessment of a natural gas station. A condition-based PRA has been performed in [6] for a spontaneous steam generator tube rupture accident. A data-driven DRA model has been developed for offshore drilling operations, where real time operational data have been used to update the probability of the kick event [19]. In [20], statistical failure data and condition monitoring data have been integrated in a hierarchical Bayesian model for DRA. DRA of an ET has been developed in [10] by using condition monitoring data to update the events probabilities.

In the existing methods, the data used for DRA can be broadly divided into two categories: statistical failure data and condition monitoring data. Statistical failure data refer to counts of accidents, incidents or near misses collected from similar systems [21]. For instance, in [22] and [23], DRA has been performed using near misses and incident data from similar processes. In [24], Bayesian theorem has been applied to update the failure probabilities of the safety barriers in a Bow-tie model for DRA. Statistical failure data are collected from a population of similar systems, which are seldom available in large number and this limits the application of the statistical failure data-

based DRA methods in practice. Also, statistical data refer to a population of similar systems and do not necessarily capture the plant-specific features of the target system. To address these issues, condition monitoring data are often used in DRA. Condition monitoring data refer to the online monitoring data collected by sensors that are installed in the target system for monitoring the degradation process of the safety barrier. For example, a condition-based fault tree has been used for DRA, where the condition monitoring data have been used to update the failure rates of the specific components and predict the reliability [25, 26]. Particle filtering (PF) has been used for DRA based on condition monitoring data from a nonlinear non-Gaussian process [27]. In [28], a Bayesian reliability updating method has been developed by using condition monitoring data considering the dependencies between two components. In [5], condition monitoring data from a passive safety system have been used for DRA, without considering the uncertainty in the condition monitoring data.

Inspection data are collected by physical inspections performed by maintenance personnel [29]. They have been widely used for online reliability assessment. For example, a Bayesian method has been developed to merge experts' judgment with continuous and discontinuous inspection data for the reliability assessment of multi-state systems [30]. A two-stage recursive Bayesian approach has been developed in [31], in order to update system reliability based on imperfect inspection data. Condition monitoring data and inspection data on wind turbine blades have been used separately for remaining useful life estimation in [32]. As inspections directly measure the component degradation, they provide valuable information complementary to condition monitoring data for DRA and can help reducing the impact of the uncertainty in the condition monitoring data on the result of DRA. However, to the best of our knowledge, no previous work has considered integrating condition monitoring data and inspection data for DRA.

In this paper, we develop a new framework to integrate condition monitoring data and inspection data in DRA. Compared to the existing works, the original contributions lie in:

- (1) a Hidden Markov-Gaussian Mixture Model is developed for modeling condition monitoring data;
- (2) a Bayesian network model is developed to integrate condition monitoring data and inspection data for DRA;
- (3) a real-world application is performed.

The rest of the paper is organized as follows. Sect. 2 introduces the engineering motivation and formally defines the problem. In Sect. 3, a HM-GMM is developed for reliability updating and prediction of the failure of safety barriers based on condition monitoring data. A Bayesian network model is developed in Sect. 4 to integrate the inspection data and condition monitoring data for DRA. The framework is tested in Sect. 5 through a numerical

example. In Sect. 6, it is applied for the DRA of a real-world NPP. Finally, conclusions and potential future works are discussed in Sect. 7.

2. Problem definitions

The framework developed in this paper is motivated by real-world PRA practices. We consider an event tree model developed for the PRA of an Anticipated Transient Without Scram (ATWS) accident of a NPP [2]. The occurrence probabilities of the basic events, associated to the reliability of the safety barriers in the ET, are estimated from statistical data and assumed to remain constant throughout the life of the NPP [2]. However, the safety barriers in practice degrade. For example, a safety barrier in the aforementioned ET is the recirculation pump [2]; according to [33], most failures of the recirculation pump are caused by the degradation of the bearings, which makes the reliability of the pump time-dependent. DRA is best suited to capture such time-dependencies.

Two types of data can be used for the DRA of the ATWS accident. The first is inspection data. Take the bearing mentioned above as an example: through inspections, the degradation state of the bearing can be identified, e.g., healthy, minor degradation (e.g., outer race defect), medium degradation (e.g., roller element defect), severe degradation (e.g., inner race defect), etc. (see Figure 1). The second type of data is condition monitoring data: some observable signals, e.g., temperature, vibration, etc., that contain information on the degradation process are measured and used to infer the degradation state. For example, the vibration signals of bearings are often used as condition monitoring data to estimate the degradation state and update the reliability of bearings [34]. Inspection data usually give discrete degradation states, with uncertainty due to state classification by the maintenance operator. Condition monitoring data are subject to uncertainty due to observation noises and degradation state estimation errors. In this paper, a new framework is proposed to integrate condition monitoring data and inspection data for improving the accuracy and reducing the uncertainty of the risk assessment.

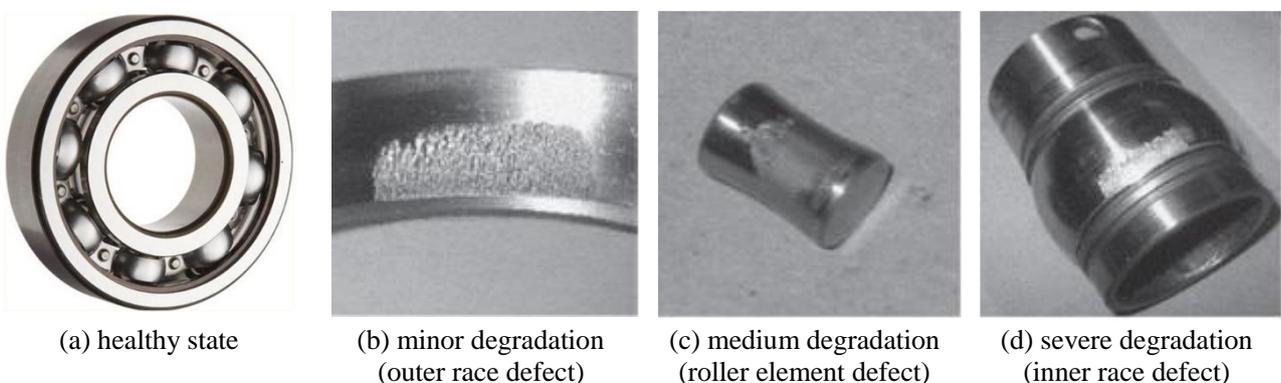


Figure 1 Degradation states of bearing [35].

Without loss of generality, we consider a generic Even Tree (ET) model for DRA, but the framework is applicable to other risk assessment models as well. Let IE represent the initialing event of the ET and assume that there are M safety barriers (SB) in the ET, denoted by $SB_i, i = 1, 2, \dots, M$, whose states can be working or failure. The sequences that emerge from the IE depend on the states of the SB s and lead to N possible consequences, denoted by C_1, C_2, \dots, C_N . The generic risk index considered in this paper is the conditional probability that a specific consequence C_i occurs, given that the IE has occurred:

$$P_{C_i} = P\{C_i \text{ occurs} | IE \text{ has occurred}\}, i = 1, 2, \dots, N. \quad (1)$$

Conditioning on the occurrence of the IE , these probabilities are functions of the reliabilities $R_{SB_i}, i = 1, 2, \dots, M$ of the safety barriers along the specific sequences:

$$P_{C_i} = f_{ET}(R_{SB_1}, R_{SB_2}, \dots, R_{SB_M}), i = 1, 2, \dots, N. \quad (2)$$

where $f_{ET}(\cdot)$ is the ET model function. For example, in the ET in Figure 2, the risk index P_{C_2} of the consequence C_2 of the second accident sequence, in which the IE occurs with certainty, the first SB_1 functions successfully and the second SB_2 fails to provide its function, can be calculated as:

$$\begin{aligned} P_{C_2} &= f_{ET}(R_{SB_1}, R_{SB_2}) \\ &= R_{SB_1}(1 - R_{SB_2}). \end{aligned} \quad (3)$$

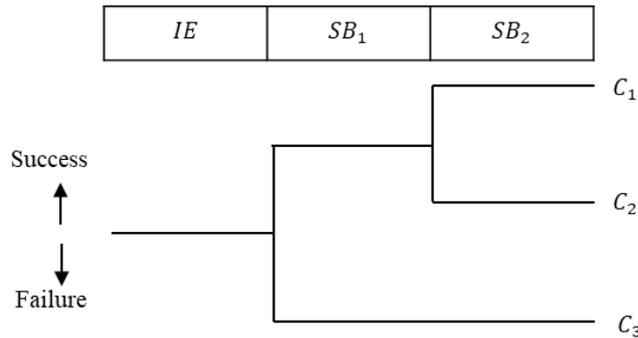


Figure 2 Illustrative Event Tree model.

Without loss of generality, we assume that in the ET:

- (5) Safety barriers SB_1, SB_2, \dots, SB_K are subject to degradation processes and, therefore, their reliability functions are time-dependent, whereas $SB_{K+1}, SB_{K+2}, \dots, SB_M$ do not degrade and have constant reliability

values;

- (6) Condition monitoring data are collected for SB_1, SB_2, \dots, SB_K at predefined time instants $t = t_k, k = 1, 2, \dots, q$;
- (7) The collected condition monitoring data on the i -th safety barrier at $t = t_k$ are denoted by $c_i(t_k)$, where $i = 1, 2, \dots, K, k = 1, 2, \dots, q$ and $\mathbf{c}_i(t) = [c_i(t_1), c_i(t_2), \dots, c_i(t_q)]$ is a vector containing all the signals that are monitored, where q is the length of the time series;
- (8) At $t = t_m$, inspections are performed on the safety barriers $SB_i, i = 1, 2, \dots, K$. The inspection data are denoted by $S_{IN,i}, i = 1, 2, \dots, K$.

The DRA tasks are formally defined as:

- (1) risk updating: at time $t = t_k, k = 1, 2, \dots, q$, update the estimated risk indexes at the current time t_k , based on the integration of condition monitoring and inspection data available up to t_k ;
- (2) risk prediction: at time $t = t_k$, predict the values of the risk indexes at future times, based on the integration of condition monitoring and inspection data available up to t_k .

3. A Hidden Markov Gaussian Mixture Model for modeling condition monitoring data

In this section, we develop a HM-GMM to model condition monitoring data. In Sect. 3.1, we formally define the HM-GMM. Then, in Sect 3.2, we show how to use the developed HM-GMM to estimate the degradation state of a safety barrier using condition monitoring data. The estimated degradation states are, then, used in Sect. 4 for data integration in DRA.

3.1 Model formulations

Without loss of generality, we illustrate the HM-GMM using the i -th safety barrier in the ET. For simplicity of presentation, we drop the subscript i in the notations. An illustration of the model is given in Figure 3. It is assumed that the safety barrier degrades during its lifetime and the degradation process follows a discrete state discrete time Markov model $S(t)$ with a finite state space $S(t) \in \{S_1, S_2, \dots, S_Q\}$, where $S(t)$ represents the health state of the safety barrier, Q is the number of health states, and S_1, S_2, \dots, S_Q are in descending order of health (S_1 is the perfect functioning state, S_Q is the failure state). The evolution of the degradation process is characterized by the transition probability matrix of the Markov process, denoted by A , where $A = \{a_{ij}\}$ and

$a_{ij} = P(S(t_{k+1}) = S_j | S(t_k) = S_i), k = 1, 2, \dots, q, 1 \leq i, j \leq Q$. The initial state distribution of the Markov process is denoted by $\boldsymbol{\pi} = [\pi_1 \ \pi_2 \ \dots \ \pi_Q]$, where $\pi_i = P(S(t_0) = S_i), 1 \leq i \leq Q$. It should be noted that repairs are not considered in this paper, just to simplify the calculation. Then, $S(t)$ can only transit to a worse state and cannot move backwards to a better state. Besides, the failure state S_Q is an absorbing state, such that $p(S(t_{k+1}) = i | S(t_k) = S_Q) = 1$ if and only if $i = S_Q$ and $p(S(t_{k+1}) = i | S(t_k) = S_Q) = 0$ for all other values of i . However, the model accommodates the case of repairable components, where the transition matrix has non-zero entries also for backward state transitions, which represent the repairs of the safety barriers. The developed algorithms, can, then, be extended naturally.

The discrete time discrete state Markov process model is chosen because it is widely applied for quantitatively describing discrete state degradation processes in many practical applications [36]. For example, a discrete state Markov model has been used to model the bearing degradation process in [35]. The degradation process of a safety instrumented system is modeled by a Markov model for availability analysis [37, 38]. Although only Markov process-based degradation models are discussed in this paper, the developed methods for data integration into DRA can be easily extended to other degradation models.

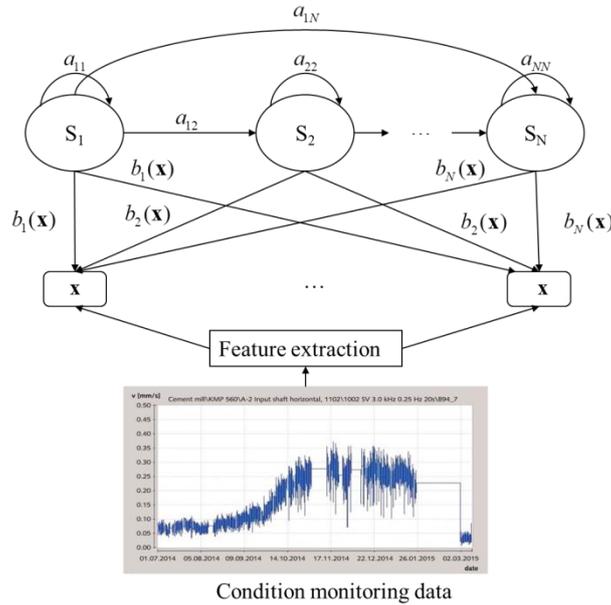


Figure 3 Description of the HM-GMM.

As described in Sect. 2.1, condition monitoring data $\mathbf{c}(t)$ are available at $t = t_k, k = 1, 2, \dots, q$. In practice, $\mathbf{c}(t)$ contains only raw signals, which cannot be directly used for degradation modeling and analysis. Feature extraction,

as shown in Figure 3, is needed to extract degradation features from $\mathbf{c}(t)$. For example, vibration signals are usually used as condition monitoring data for bearings [24]. The raw vibration signals, however, need to be preprocessed to extract features for degradation characterization. The commonly used degradation features include entropy, root mean square (RMS), kurtosis, etc [39]. In this paper, we refer to these extracted features as degradation indicators and denote them by $\mathbf{x}(t)$, where $\mathbf{x}(t) = [x_1(t), x_2(t), \dots, x_{n_{feature}}(t)]$ and $n_{feature}$ is the number of the degradation features.

As the safety barrier degrades, the degradation indicator $\mathbf{x}(t)$ exhibits distinct patterns. To capture such patterns and the uncertainty associated with them, it is assumed that at each degradation state $S_i, 1 \leq i \leq Q$, the values of the degradation indicators \mathbf{x} follow a multivariate Gaussian distribution $b_i(\mathbf{x}) = p(\mathbf{x}|S(t) = S_i) = N(\mathbf{x}|\boldsymbol{\mu}^{(i)}, \boldsymbol{\Sigma}^{(i)}), i = 1, 2, \dots, Q$, as shown in Figure 3. The mean values vector $\boldsymbol{\mu}^{(i)}$ captures the degradation pattern at each degradation state, while the covariance matrix $\boldsymbol{\Sigma}^{(i)}$ captures the uncertainty in the condition monitoring data. An overall picture of the HM-GMM is given in Figure 3. Conceptually, we denote the HM-GMM compactly as $\boldsymbol{\lambda} = \{\boldsymbol{\pi}, A, \boldsymbol{\mu}, \boldsymbol{\Sigma}\}$, where $\boldsymbol{\pi}$ is the initial state distribution, A is the transition probability matrix, $\boldsymbol{\mu} = [\boldsymbol{\mu}_1, \boldsymbol{\mu}_2 \dots, \boldsymbol{\mu}_Q]$ is a vector of the mean values and $\boldsymbol{\Sigma} = [\boldsymbol{\Sigma}^{(1)}, \boldsymbol{\Sigma}^{(2)}, \dots, \boldsymbol{\Sigma}^{(Q)}]$ is a collection of the covariance matrices of the multivariate Gaussian distribution, respectively.

3.2 Degradation states estimation based on condition monitoring data

In this section, we show how to estimate the degradation states of the safety barriers based on the developed HM-GMM of the condition monitoring data. As shown in Figure 4, the estimation is made by an offline step and an online step. In the offline step, a HM-GMM is trained based on training data from a population of similar systems. The trained HM-GMM model, is, then, used in the online step for degradation state estimation based on the condition monitoring data.

The offline step starts from collecting training data, denoted by $\mathbf{c}_{T_r}^{(k)}(t), k = 1, 2, \dots, n_{T_r}, t = t_1, t_2, \dots, t_{T_r}$. The training data comprise of historical measurements of the degradation signals from a population of similar systems. To ensure the accuracy of HM-GMM training, it is required to collect as many as possible training samples, i.e., the sample size n_{T_r} should be as large as possible. The raw training data are preprocessed in a feature extraction step, as shown in Figure 4, to extract the health indicators $\mathbf{x}_{T_r}^{(k)}(t), k = 1, 2, \dots, n_{T_r}, t = t_1, t_2, \dots, t_{T_r}$. Depending on the nature of the degradation process condition, different feature extraction methods, e.g., time-domain, frequency domain, time-

frequency analyses, etc., can be used [39]. Next, in the HM-GMM training step, the extracted degradation indicators are used to estimate the parameters $\lambda = \{\pi, A, \mu, \Sigma\}$ of the trained HM-GMM. In this paper, the Expectation Maximization (EM) algorithm [40] is employed for training the HM-GMM (see Sect. 3.2.1 for details). The parameters λ is the output of the offline step.

The online step starts from collecting the condition monitoring data for the safety barrier, denoted by $\mathbf{c}(t_k), k = 1, 2, \dots, q$. The condition monitoring data should be of the same type and collected by the same sensors, as in the offline step. Then, the raw degradation signals are preprocessed and the health indicators $\mathbf{x}(t_k), k = 1, 2, \dots, q$ of the target safety barrier are extracted, following the same procedures as in the offline step. Next, the degradation state of the safety barrier is estimated, based on the HM-GMM trained in the offline step. In this paper, we use the forward algorithm for degradation state estimation [40], as presented in details in Sect. 3.2.2. The estimated degradation state based on only condition monitoring data, denoted by $S_{CM}(t_k)$, is, then, integrated with inspection data for DRA in Sect. 4.

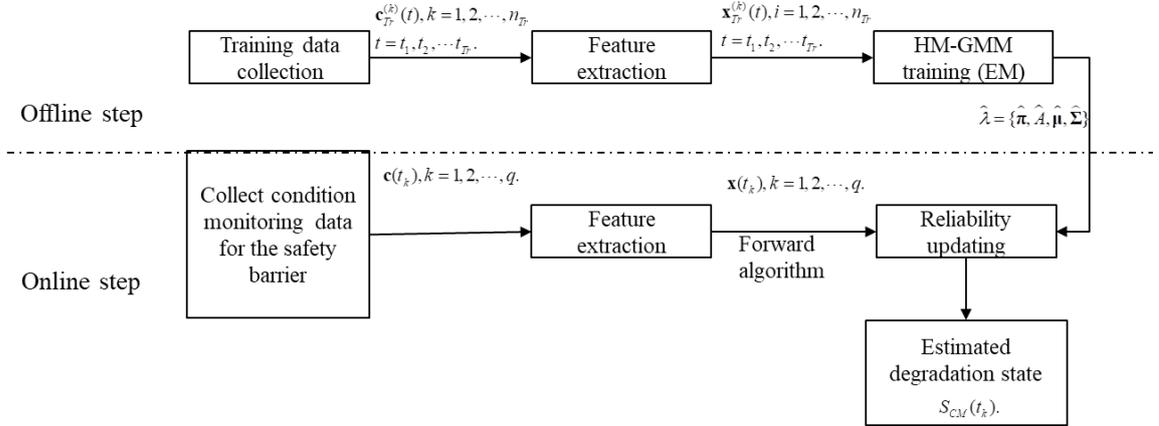


Figure 4 Degradation state estimation based on condition monitoring data.

3.2.1 HM-GMM training

In this section, we present in detail how to do HM-GMM training in the offline step. The parameters $\lambda = \{\pi, A, \mu, \Sigma\}$ are estimated by maximizing the likelihood of observing the $\mathbf{x}_{Tr}^{(k)}(t), k = 1, 2, \dots, n_{Tr}, t = t_1, t_2, \dots, t_{Tr}$:

$$\begin{aligned} \lambda &= \arg \max_{\lambda} P\left(\mathbf{x}_{Tr}^{(1)}(t), \mathbf{x}_{Tr}^{(2)}(t), \dots, \mathbf{x}_{Tr}^{(n_{Tr})}(t) | \lambda\right) \\ &= \arg \max_{\lambda} \prod_{k=1}^{n_{Tr}} P\left(\mathbf{x}_{Tr}^{(k)}(t) | \lambda\right) \end{aligned} \quad (4)$$

Let $L \triangleq \prod_{k=1}^{n_{Tr}} P(\mathbf{x}_{Tr}^{(k)}(t) | \boldsymbol{\lambda})$ be the likelihood function of the observation data. Directly solving (4) is not possible in

practice, as the likelihood function in (4) contains unobservable variables (the true degradation states $S(t)$ in this case). Expectation Maximization (EM) algorithm [40] is applied to solve this problem, where the maximum likelihood estimator is found in an iterative way: the current values of the parameters are used to estimate the unobservable variables (Expectation phase); then, the estimated values of the unknown variables are substituted into the likelihood function to update the maximum likelihood estimators of the parameters (Maximization phase). The iterative procedures are repeated until the maximum likelihood estimators converge.

To apply the EM algorithm to the HM-GMM model, two auxiliary variables need to be defined first, i.e., forward variable $\alpha_i(S_i)$ and backward variable $\beta_i(S_i)$. The forward variable is defined as the probability of observing the health indicators up to the current time t and that the true degradation state $S(t) = S_i$, given a known HM-GMM $\boldsymbol{\lambda}$:

$$\alpha_i(S_i) = P(\mathbf{x}(t_1), \mathbf{x}(t_2), \dots, \mathbf{x}(t), S(t) = S_i | \boldsymbol{\lambda}). \quad (5)$$

It is easy to verify that

$$\begin{aligned} \alpha_1(S_i) &= \pi_i b_i(\mathbf{x}(t_1)), \\ \alpha_{t+1}(S_j) &= b_j(\mathbf{x}_{t+1}) \left[\sum_{i=1}^Q \alpha_t(S_i) a_{ij} \right], 1 \leq i \leq Q, 1 \leq j \leq Q, 1 \leq t \leq t_{Tr} - 1, \end{aligned} \quad (6)$$

where t_{Tr} represents the observation time length and all the elements in π_i are zero, except the one that corresponds to the i -th element being one.

The backward probability $\beta_i(S_i)$ is defined as the probability of observing the health indicator $\mathbf{x}(t+1), \mathbf{x}(t+2), \dots, \mathbf{x}(t_{Tr})$ from $t+1$ to the end of the observations, given that $S(t) = S_i$ and the model parameters are $\boldsymbol{\lambda}$:

$$\beta_i(S_i) = P(\mathbf{x}(t+1), \mathbf{x}(t+2), \dots, \mathbf{x}(t_{Tr}) | S(t) = S_i, \boldsymbol{\lambda}). \quad (7)$$

It is easy to verify that $\beta_t(S_j) = \left[\sum_{i=1}^Q b_j(\mathbf{x}(t+1)) a_{ij} \right] \beta_{t+1}(S_j), 1 \leq i, 1 \leq j \leq Q, \beta_{t_{Tr}}(i) = 1, t = t_{Tr} - 1, t_{Tr} - 2, \dots, 1$.

The iterative estimators for the transition probabilities, denoted by a_{ij} , can, then, be derived as follows [41]:

$$a_{ij} = \frac{\sum_{k=1}^{n_{Tr}} \sum_{t=1}^{t_{Tr}} \xi_{Tr,t}^{(k)}(S_i, S_j)}{\sum_{k=1}^{n_{Tr}} \sum_{t=1}^{t_{Tr}} \gamma_{Tr,t}^{(k)}(S_i)}, \quad (8)$$

where $\xi_{Tr,t}^{(k)}(S_i, S_j)$ represents the probability of the k -th sample being in S_i at time t and state S_j at time $t+1$, and is calculated by [41]:

$$\begin{aligned} \zeta_{Tr,t}^{(k)}(S_i, S_j) &= P(S(t) = S_i, S(t+1) = S_j | \mathbf{x}_{Tr}^{(k)}(t+1), \boldsymbol{\lambda}) \\ &= \frac{\gamma_{Tr,t}^{(k)}(S_i) a_{ij} b_{Tr,j}^{(k)}(\mathbf{x}_{Tr}^{(k)}(t+1)) \beta_{Tr,t+1}^{(k)}(S_j)}{\beta_{Tr,t}^{(k)}(S_i)}, \end{aligned} \quad (9)$$

where $\gamma_{Tr,t}^{(k)}(S_i)$ represents the probability of being in S_i at time t given the health indicator $\mathbf{x}_{Tr}^{(k)}(t)$ and $\boldsymbol{\lambda}$ for the k -th training sample:

$$\gamma_{Tr,t}^{(k)}(S_i) = \frac{\alpha_{Tr,t}^{(k)}(S_i) \beta_{Tr,t}^{(k)}(S_i)}{p(\mathbf{x}_{Tr}^{(k)}(t) | \boldsymbol{\lambda})} = \frac{\alpha_{Tr,t}^{(k)}(S_i) \beta_{Tr,t}^{(k)}(S_i)}{\sum_{i=1}^Q \alpha_{Tr,t}^{(k)}(S_i) \beta_{Tr,t}^{(k)}(S_i)}. \quad (10)$$

The estimator for the initial state probability $\pi_i, i = 1, 2, \dots, Q$ is calculated by [40]:

$$\pi_i = \frac{\sum_{k=1}^{n_{Tr}} \gamma_{Tr,t}^{(k)}(S_i)}{n_{Tr}}. \quad (11)$$

The estimators of the mean value vectors are derived as [41]:

$$\boldsymbol{\mu}_i = \frac{\sum_{k=1}^{n_{Tr}} \sum_{t=1}^{t_{Tr}} \gamma_{Tr,t}^{(k)}(S_i) \mathbf{x}_{Tr}^{(k)}(t)}{\sum_{k=1}^{n_{Tr}} \sum_{t=1}^{t_{Tr}} \gamma_{Tr,t}^{(k)}(S_i)}. \quad (12)$$

Similarly, the covariance matrices of the Gaussian output are calculated by [41]:

$$\boldsymbol{\Sigma}_i = \frac{\sum_{k=1}^{n_{Tr}} \sum_{t=1}^{t_{Tr}} \gamma_{Tr,t}^{(k)}(S_i) (\mathbf{x}_{Tr}^{(k)}(t) - \boldsymbol{\mu}_i) (\mathbf{x}_{Tr}^{(k)}(t) - \boldsymbol{\mu}_i)'}{\sum_{k=1}^{n_{Tr}} \sum_{t=1}^{t_{Tr}} \gamma_{Tr,t}^{(k)}(S_i)}. \quad (13)$$

Algorithm 1 below summarizes the procedures for training the HM-GMM based on the EM algorithm. In Algorithm 1, $\|\cdot\|$ measures the distance between the current and the previous estimators. In this paper, we use the absolute value for its calculation, and tol is the tolerance of the error. In this paper, we set $tol = 1 \times 10^{-4}$.

Algorithm 1: HM-GMM training based on EM algorithm.

Inputs: $\lambda_0 = \{\pi_0, A_0, \mu_0, \Sigma_0\}, \mathbf{x}_{Tr}^{(1)}(t), \mathbf{x}_{Tr}^{(2)}(t), \dots, \mathbf{x}_{Tr}^{(n_{Tr})}(t)$;

Outputs: $\lambda = \{\pi, A, \mu, \Sigma\}$;

Step 1: $\lambda = \lambda_0$;

Step 2: Expectation phase: calculate the forward and backward variables, based on (5) and (7), respectively, using the current value of λ ;

Step 3: Maximization phase: update λ based on (8), (11)-(13), respectively;

Step 4: If $\|\lambda - \lambda_{prev}\| < tol$, End;

Else, $\lambda_{prev} = \lambda$, go to Step 2.

3.2.2 Degradation state estimation

In this paper, the forward algorithm [40] is employed to estimate the degradation state of the safety barriers in the online step. Let S_{CM} denote the estimated degradation state from condition monitoring data and $P_{CM,t_k}(S_{CM}), k = 1, 2, \dots, q$ represent the posterior distribution of S_{CM} given the condition monitoring data up to t_k :

$$P_{CM,t_k}(S_{CM} = S_i) = P(S(t_k) = S_i | \mathbf{x}(t_1), \mathbf{x}(t_2), \dots, \mathbf{x}(t_k), \lambda) \quad (14)$$

The posterior probabilities defined in (14) can be easily calculated from the forward probabilities defined in (15):

$$\begin{aligned} P_{CM,t_k}(S_{CM} = S_i) &= \frac{P(S(t_k) = S_i, \mathbf{x}(t_1), \mathbf{x}(t_2), \dots, \mathbf{x}(t_k) | \lambda)}{P(\mathbf{x}(t_1), \mathbf{x}(t_2), \mathbf{x}(t_2), \dots, \mathbf{x}(t_k) | \lambda)} \\ &= \frac{\alpha_{i_k}(S_i)}{\sum_{i=1}^q \alpha_{i_k}(S_i)}. \end{aligned} \quad (15)$$

In practice, the $\alpha_{i_k}(S_i)$ in (15) is calculated recursively, based on (5).

At each $t = t_k$, the most likely degradation state, denoted by $S_{CM,MAP}(t_k)$, is, then, determined by finding the state with maximal posterior probability:

$$S_{CM,MAP}(t_k) = \arg \max_{1 \leq i \leq Q} [P_{CM,t_k}(S_{CM} = S_i)], 1 \leq k \leq q. \quad (16)$$

Algorithm 2 below summarizes the major steps used for estimating the degradation state.

Algorithm 2 Forward algorithm for degradation state estimation at $t = t_k$.

Input: $\lambda = \{\pi, A, \mu, \Sigma\}, \alpha_{t_{k-1}}(S_i), i = 1, 2, \dots, Q, \mathbf{x}(t_k)$;

Output: $P_{CM,t_k}(S_{CM}), S_{CM,MAP}(t_k)$;

Step 1: Calculate $\alpha_{t_k}(S_i), i = 1, 2, \dots, Q$, by (6);

Step 2: Calculate the posterior probability $P_{CM,t_k}(S_{CM})$ by (15);

Step 3: Estimate the degradation state $S_{CM,MAP}(t_k)$ by (16).

4. Integrating condition monitoring data with inspection data for DRA

In this section, we first show how to integrate the condition monitoring data with inspection data for reliability updating and prediction of the safety barriers (Sect. 4.1). Then, in Sect. 4.2, we develop a DRA method based on the updated and predicted reliabilities.

4.1. A Bayesian network model for data integration

As in the previous sections, we illustrate the developed data integration method using the i -th safety barrier at $t = t_k$. For simplicity and to avoid confusion, we drop the i and t_k in the notations. To update and predict the reliability, one needs to estimate the degradation state first. Let S_{IN} denote the degradation state estimated from inspection data and S denote the true degradation state. In practice, S_{IN} is subject to uncertainty due to potential imprecision in the inspection and recording by the maintenance personnel. To model such uncertainty, in this paper, we assume that the reliability of inspection is R_{IN} , and that the maintenance personnel correctly identify the true degradation state with a probability R_{IN} , whereas an inspection error can occur with probability $(1 - R_{IN})$. When an inspection error occurs, it is further assumed that the probabilities for each of the possible degradation states being erroneously identified as the true degradation state are equal to each other:

$$P(S_{IN} = S_i | S) = \begin{cases} R_{IN}, & S = S_i \\ \frac{1 - R_{IN}}{Q - 1}, & S \neq S_i, \end{cases} \quad (17)$$

where Q is the number of degradation states. It should be noted that other inspection models might also be assumed, depending on the actual problem setting.

In this paper, a BN is developed to describe the dependencies among S, S_{IN}, S_{CM} , as shown in Figure 5. The BN in Figure 5 is constructed based on the assumption that given the true degradation state S , the estimated degradation state from condition monitoring data and inspection data are conditional-independent.

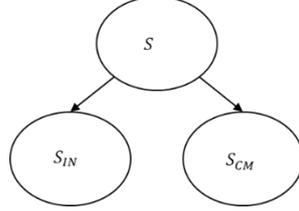


Figure 5 A BN model for data integration.

Based on the BN in Figure 5, we have

$$P(S, S_{IN}, S_{CM}) = P(S_{IN} | S) P(S_{CM} | S) P(S). \quad (18)$$

In (18), $P(S)$ measures the prior belief of the analysts on the current degradation states. We assume that $P(S)$ is a uniform distribution over all the possible degradation states, indicating that there is no further information to distinguish the states.

The conditional probability distribution $P(S_{IN} | S)$ describes the uncertainty in the inspections and is derived based on (17). In (17), the reliability of the inspection can be estimated from historical data or assigned based on expert judgments. The conditional probability distribution $P(S_{CM} | S)$ measures the trust one has on the estimated degradation state based on condition monitoring data. Its values can be estimated from validation test data. However, in practice, as validation tests are not always available, $P(S_{CM} | S)$ might also be assigned by experts considering the measurement uncertainty of the sensors and the distance between the neighboring degradation states. We give an example of how to determine $P(S_{CM} | S)$ in the case study of Sect. 6.

Once the condition monitoring data and inspection data are available, the observed values of S_{IN} and S_{CM} are known. Suppose we have $S_{CM} = S_j$ and $S_{IN} = S_i$. It should be noted that we choose the state with maximal posterior probability from (16) as the observation value of S_{CM} . The two data sources can be naturally integrated by calculating the posterior distribution of S given the two data sources, denoted by $P_{INT}(S)$. Based on the BN in Figure 5, we have:

$$\begin{aligned}
P_{INT}(S) &\triangleq P(S | S_{IN} = S_i, S_{CM} = S_j) \\
&= \frac{P(S, S_{IN} = S_i, S_{CM} = S_j)}{P(S_{IN} = S_i, S_{CM} = S_j)} \\
&= \frac{P(S_{IN} = S_i | S) P(S_{CM} = S_j | S) P(S)}{P(S_{IN} = S_i, S_{CM} = S_j)}
\end{aligned} \tag{19}$$

Given the estimated posterior distribution in (19), the reliability of the safety barrier can be updated. Suppose the current time is t_k , the updated reliability can be calculated by:

$$R_{SB}(t_k) = \sum_{S \in W} P_{INT, t_k}(S), \tag{20}$$

where W is the working set that contains all the working states; $P_{INT, t_k}(S)$ is the posterior probability of the true degradation state after integrating the two data sources at $t = t_k$ and is calculated from (19).

Furthermore, at $t = t_k$, we can also predict the reliability of the safety barriers at a future time t_{Fut} . For this, the distribution of the degradation states at $t = t_{Fut}$ is predicted first, using Chapman-Kolmogorov equation [42] and the trained model from the offline step:

$$P_{INT, t_{Fut}}(S) = P_{INT, t_k}(S) \times A^{(t_{Fut} - t_k)}. \tag{21}$$

The reliability at $t = t_k$, can be predicted as:

$$R_{SB}(t_{Fut}) = \sum_{S \in W} P_{INT, t_{Fut}}(S). \tag{22}$$

4.2. Dynamic risk assessment

The updated reliabilities from (20), can, then, be substituted into (2) for DRA:

$$r_{C_i}(t_k) = f_{ET}(R_{SB_1}(t_k), R_{SB_2}(t_k), \dots, R_{SB_K}(t_k), R_{SB_{K+1}}, \dots, R_{SB_M} | IE), i = 1, 2, \dots, N, \tag{23}$$

where in (23), $R_{SB_i}(t_k)$ is calculated by (20). Similarly, the risk index at a future time t_{Fut} can be predicted by:

$$r_{C_i}(t_{Fut}) = f_{ET}(R_{SB_1}(t_{Fut}), R_{SB_2}(t_{Fut}), \dots, R_{SB_K}(t_{Fut}), R_{SB_{K+1}}, \dots, R_{SB_M} | IE), i = 1, 2, \dots, N, \tag{24}$$

where $R_{SB_i}(t_{Fut})$ is calculated by (21) and (22).

Figure 6 summarizes the major steps for the developed DRA method by integrating condition monitoring data with inspection data. It should be noted that in Figure 6, the risk updating is made at $t = t_k$, while risk prediction is made for a given future time t_{Fut} .

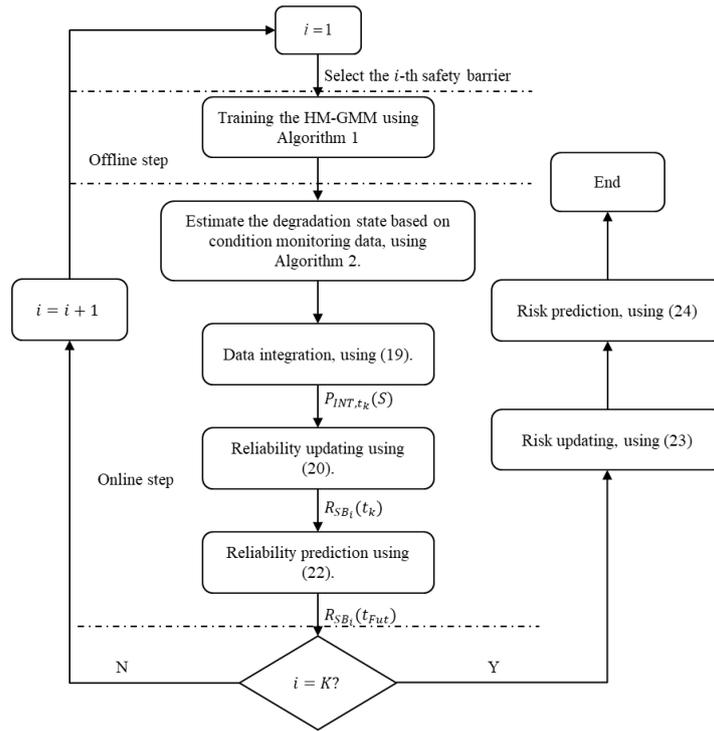


Figure 6 Procedures for DRA based on condition monitoring and inspection data.

5. Numerical case study

In this section, we apply the DRA framework for data integration (see Sect. 4.1) on a numerical case study. The purpose is to test the updating and prediction of safety barrier reliability. Hence, only reliability updating and prediction are considered. The application of the overall DRA framework is done in Sect. 6 on a real-world case.

Consider a component whose degradation process follows a discrete state discrete time Markov chain $S(t)$ with four discrete degradation states S_1, S_2, S_3, S_4 , where $S_1 \sim S_4$ have increasing degrees of degradation from S_1 perfect state, to S_4 failure state. The condition monitoring data are generated from a HM-GMM with known parameters values:

$$\begin{aligned}
\mathbf{A} &= \begin{pmatrix} 0.6 & 0.2 & 0.1 & 0.1 \\ 0 & 0.5 & 0.25 & 0.25 \\ 0 & 0 & 0.5 & 0.5 \\ 0 & 0 & 0 & 1 \end{pmatrix}, \\
\boldsymbol{\pi} &= [1 \ 0 \ 0 \ 0], \\
\boldsymbol{\mu} &= \begin{pmatrix} 0.0588 & 0.1424 & 0.1842 & 1 \\ 0.1268 & 0.1597 & 0.2432 & 1 \\ 0.0946 & 0.9744 & 0.8648 & 0.8449 \end{pmatrix}, \\
\boldsymbol{\Sigma}^{(i)} &= \begin{pmatrix} 0.001 & 0 & 0 \\ 0 & 0.001 & 0 \\ 0 & 0 & 0.001 \end{pmatrix}, \text{ with } i = 1, 2, 3, 4.
\end{aligned} \tag{25}$$

The degradation indicator comprises of three features, denoted by x_1, x_2 and x_3 , respectively. The size of the generated training data is 10^4 and $t = t_1, t_2, \dots, t_{23}$ are the time instants of data collection. Then, the training data can be represented as $\mathbf{x}_{Tr}^{(k)}(t), k = 1, 2, \dots, 10^4, t = t_1, t_2, \dots, t_{23}$, where $\mathbf{x}_{Tr}^{(k)}(t) = [x_{Tr,1}^{(k)}(t), x_{Tr,2}^{(k)}(t), x_{Tr,3}^{(k)}(t)]$. The training data are used in the offline step for estimating the model parameters. Then, another sample, denoted by $\mathbf{x}_{CM}(t), t = 1, 2, \dots, t_{CM}$, is generated from the HM-GMM in (25) and used as condition monitoring data collected on the safety barrier monitored in the online step, as shown in Figure 7.

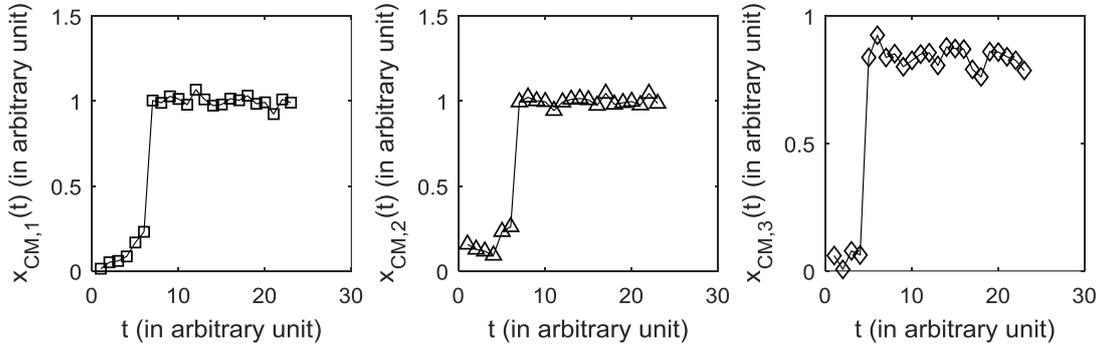


Figure 7 The generated condition monitoring data for the monitored safety barrier.

Based on the generated condition monitoring data, the reliability updating and prediction can be done using Algorithm 1 and equations (20) and (22). Due to the noise in the condition monitoring data, the updated reliability is subject to uncertainty. The method in Figure 6 is applied to solve this problem by integrating condition monitoring data with inspection data. In this section, we test the performance of the developed data integration method under three possible scenarios:

- (1) Both condition monitoring data and inspection data correctly estimate the degradation state: this scenario

is represented by choosing the time point $t=t_3$, where the estimated degradation state from condition monitoring data and the true degradation state are both S_2 . The inspection data at t_k is generated to be exactly $S_{IN}(t_3) = S_2$.

- (2) Condition monitoring data correctly estimate the degradation state, but inspection data do not: this scenario is represented by choosing the time point $t=t_7$, where the estimated degradation state from condition monitoring data and the true state are both S_3 , whereas the inspection data at t_7 is randomly sampled from $S_k, k = 1, \dots, Q, k \neq 3$. The state from the inspection data is $S_{IN}(t_3) = S_2$.
- (3) Inspection data correctly estimate the degradation state, but condition monitoring data do not: this scenario is generated by choosing the time point $t=t_5$, where the estimated degradation state from condition monitoring data is $S_{CM}(t_5) = S_2$, whereas the true degradation state is $S(t_5) = S_3$. The inspection data at t_5 are generated to be $S_{IN}(t_5) = S(t_5) = S_3$.

In subsections 5.1-5.3, we apply the developed data integration method on the three scenarios above.

5.1 Scenario I : Both data sources are reliable

The reliability updating and prediction processes are conducted following the procedures in Figure 6, at $t = t_3$. The updated and predicted reliability are compared to those calculated based on only condition monitoring data and only inspection data, respectively. The comparison is shown in Figure 8. We also show the relative errors of the three methods with respect to the true values in Table 1.

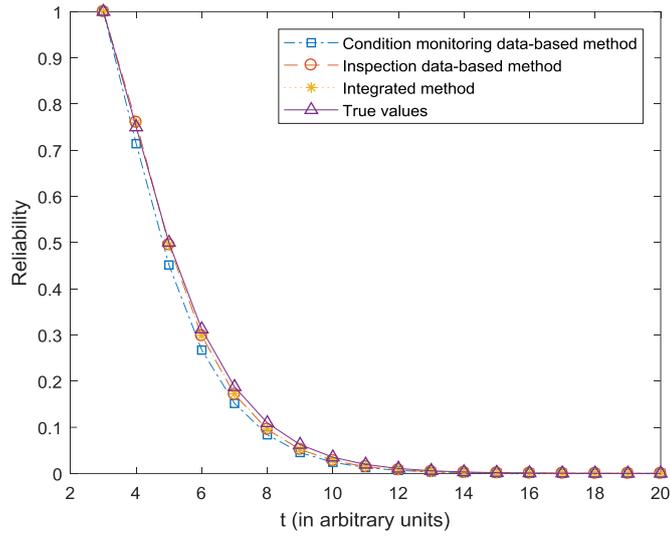


Figure 8 Updated and predicted reliability at $t = t_3$ (scenario I).

Table 1 Relative errors of the scenario I

	$t = t_3$	$t = t_4$	$t = t_5$	$t = t_6$	$t = t_7$	$t = t_8$	$t = t_9$	$t = t_{10}$
Condition monitoring data-based method	0	4.8%	9.7%	14.5%	19%	23%	27%	31%
Inspection data-based method	0	1.34%	0.9%	4.6%	8.7%	12.9%	17%	21%
Integrated method	0	1.2%	0.9%	4.3%	7%	11.7%	15%	18.6%

As shown in Figure 8 and Table 1, the proposed method provides a more accurate estimation and prediction of the reliability than the other two methods. This is because condition monitoring data are affected by noise from the data collection process, which results in uncertainty in the estimated degradation state. In this case, the state distribution estimated by the condition monitoring data is

$$P_{CM,t_3}(S_{CM}) = [0 \quad 0.8263 \quad 0.1737 \quad 0], \quad (26)$$

whereas the one estimated by integrating the two data sources is

$$P_{INT,t_3}(S) = [0.01 \quad 0.98 \quad 0.01 \quad 0]. \quad (27)$$

It can be seen that integrating the two data sources reduces the uncertainty in the degradation state estimation (note that at $t = t_3$, the true degradation state is S_2). Therefore, the updated and predicted reliabilities are more accurate than only using condition monitoring data.

On the other hand, the transition probability matrix A estimated from the offline step is

$$A = \begin{bmatrix} 0.6010 & 0.2125 & 0.0865 & 0.1 \\ 0 & 0.4483 & 0.3121 & 0.2395 \\ 0 & 0 & 0.4938 & 0.5062 \\ 0 & 0 & 0 & 1 \end{bmatrix}. \quad (28)$$

Comparing (28) to the true values in (25), it can be seen that when the current state is S_2 , the estimated A tends to underestimate the reliability as it overestimates the transition probabilities to the failure states. As the inspection data estimate that the system is in S_2 , using only inspection data tends to underestimate the reliability. Integrating the two data sources, as shown in (27), predicts that the safety barrier is also likely to be in S_1 , which compensates the errors in the estimated λ and results in more accurate reliability estimates.

5.2 Scenario II: Condition monitoring data are reliable but inspection data are not

The reliability updating and prediction processes are conducted following the procedures in Figure 6, at $t = t_7$. The updated and predicted reliability are compared to those calculated based on only condition monitoring data and only inspection data, respectively. The comparison is shown in Figure 9. We also present the relative error of the three methods by comparing them to the true values in Table 2.

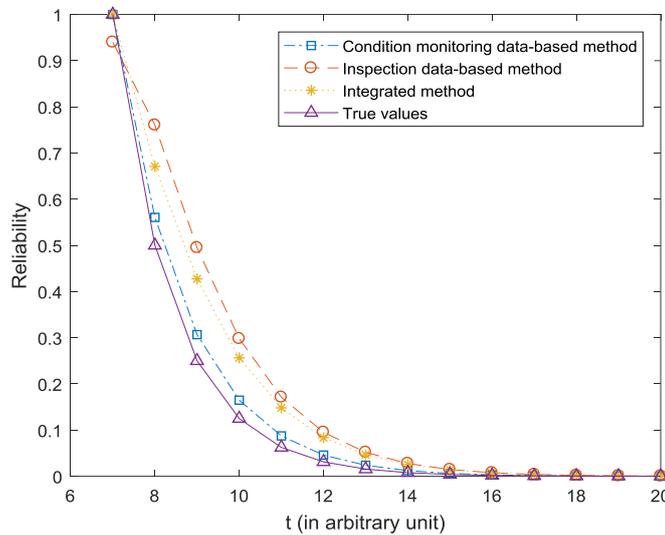


Figure 9 Updated and predicted reliability at $t = t_7$ (scenario II).

Table 2 Relative errors of the scenario II.

$t = t_7$	$t = t_8$	$t = t_9$	$t = t_{10}$	$t = t_{11}$	$t = t_{12}$	$t = t_{13}$	$t = t_{14}$
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Condition monitoring data-based method	0	12%	22%	33%	39%	46%	52%	57%
Inspection data-based method	0	52%	98%	138%	173%	204%	232%	255%
Integrated method	6%	34%	71%	96%	105%	137%	158%	197%

As shown in Figure 9 and Table 2, the results obtained by the inspection-data based method have the largest estimation error. The proposed data integration method provides more accuracy than the inspection data-based method. This is expected, as in this case the inspection data fail to correctly estimate the degradation state. By integrating condition monitoring data, the incorrect information from inspection data can be somewhat corrected. On the contrary, the estimation error of the data integration method is larger than that of the condition monitoring data-based method. This is because the data integration method is affected by the incorrect information from the inspection data. Trustworthiness of the inspection becomes essential, then.

5.3 Scenario III: Inspection data are reliable but condition monitoring data are not

The reliability updating and prediction are conducted following the procedures in Figure 6, at $t = t_5$. The updated and predicted reliability are compared to those calculated based on only condition monitoring data and only inspection data, respectively. The comparison is shown in Figure 10. We also present the relative errors of the three methods by comparing them to the true values in Table 3.

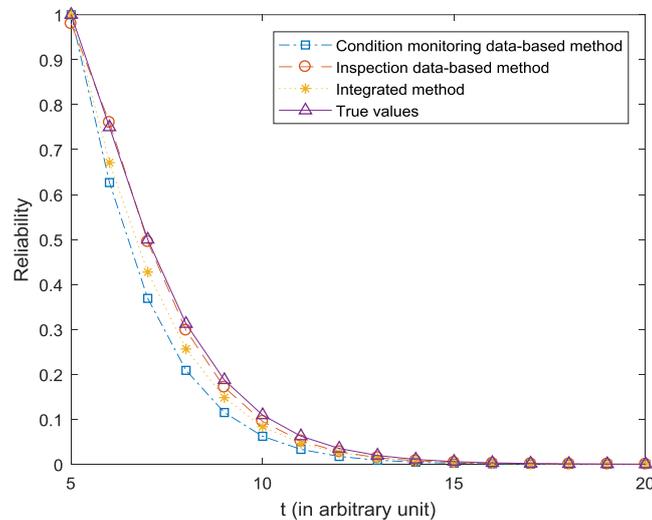


Figure 10 Updated and predicted reliability at $t = t_5$ (scenario III).

Table 3 Relative errors of the scenario III.

	$t = t_5$	$t = t_6$	$t = t_7$	$t = t_8$	$t = t_9$	$t = t_{10}$	$t = t_{11}$	$t = t_{12}$
Condition monitoring data-based method	0	16%	26%	14.5%	33%	38.5%	43%	46%
Inspection data-based method	0	1.39%	2.9%	4.6%	8.6%	12.9%	16.9%	21%

Integrated method	2%	10%	14%	17%	20%	23%	25%	27%
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As shown in Figure 10 and Table 3, the results obtained by the condition monitoring data-based method have the largest estimation errors. This is expected as in this case, the condition monitoring data fail to correctly estimate the degradation state. The proposed data integration method provides a more accurate result than the condition monitoring data-based method. This is because, by integrating inspection data, the incorrect estimation from the condition monitoring data can be compensated. However, the estimation error is larger than that of the inspection data-based method. This is because the data integration method also considers the incorrect information from the condition monitoring data.

In practical operation, the developed method can help the stakeholder/decision-makers to determine when to perform preventive maintenance on critical safety barriers. This is done by setting a minimum acceptable value for reliability and calculating the first time the reliability drops below this value. However, the reliability estimation can sometimes be imprecise. The developed method, can, then, provide a more realistic assessment to support decision making regarding when a preventive replacement is needed.

6. Application

In this section, the developed method is applied for DRA of an Anticipated Transient Without Scram (ATWS) accident of a NPP [2]. The description of the case study is briefly introduced in Sect. 6.1. Then, in Sect. 6.2, the developed HM-GMM and the data integration process are presented. The results of the DRA are presented and discussed in Sect. 6.3.

6.1 System description

ATWS is an accident that can happen in a NPP. In this accident, the scram system, which is designed to shut down the reactor during an abnormal event (anticipated transient), fails to work [43]. An ET has been developed for PRA of the ATWS for a NPP in China [2], as shown in Figure 11. In Figure 11, T₁ACM represents the failure of the automatic scram system and is the initialing event (IE) considered. Eleven safety barriers ($SB_1 \sim SB_{11}$) are designed to contain the accident (Table 4). Depending on the states of the safety barriers, 23 sequences can be generated ($SE_{01} - SE_{23}$) [2, 44]. The consequences of the sequences are grouped into two categories, based on their severity; the first group,

$$C_s = \{SE_{03}, SE_{06}, SE_{07}, SE_{08}, SE_{09}, SE_{12}, SE_{13}, SE_{14}, SE_{15}, SE_{18}, SE_{19}, SE_{20}, SE_{21}, SE_{22}, SE_{23}\}, \quad (29)$$

represents the event sequences with severe consequences, whereas the remaining event sequences have non-severe consequences [44]. The risk index *Risk* considered in this paper is the conditional probability of having severe consequences, given the initialing event ($IE = T_1ACM$):

$$Risk \triangleq P(C_S | IE) = f_{ET}(R_{SB_1}, R_{SB_2}, \dots, R_{SB_M} | T_1ACM), \quad (30)$$

where the model function $f_{ET}(\bullet)$ is determined from the ET in Figure 11 and $R_{SB_1}, R_{SB_2}, \dots, R_{SB_M}$ are the reliabilities of the safety barriers, calculated based on the component failure probabilities in Table 4. It should be noted that the failure probabilities for SB_7 and SB_8 change depending on the event sequence that occurs (see, e.g., $P_{f,SB_7}^{(1)}$ and $P_{f,SB_7}^{(2)}$ in Figure 11 and Table 4).

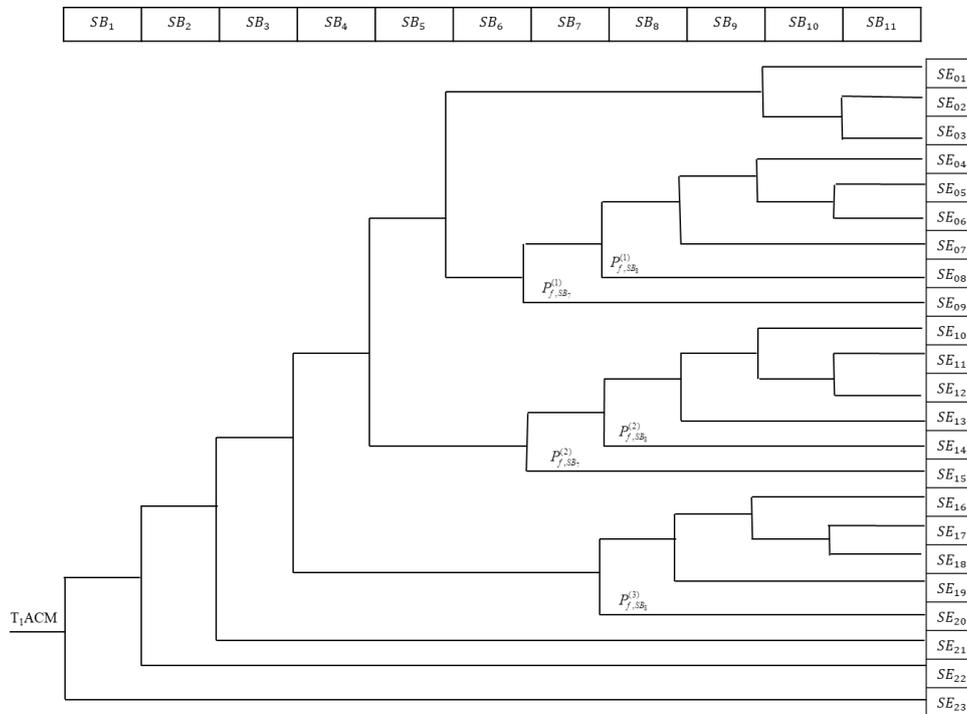


Figure 11 ET for the ATWS [44]; at each branching, the upper branch corresponds to the non-failure of the safety barrier and the low branch corresponds to the failure of the safety barrier.

In this original ETA of the ATWS, the failure probabilities in Table 4 are assumed to be constant values. In practice, however, these probabilities might change due to various degradation mechanisms. Take the recirculation pump as an example. According to [33], most field failures of the recirculation pump are caused by the degradation

of the bearing inside the pump, which makes the failure probability of the recirculation pump time-dependent. In this paper, we make a DRA on the ET in Figure 11, considering the degradation of the bearing in the recirculation pump.

The condition monitoring data of the bearing come from the bearing degradation dataset from university of Cincinnati [45]. The dataset contains four samples and for each sample, raw condition monitoring data are collected in real time by measuring the vibration acceleration signals. An illustration of the raw data is given in Figure 12. On the other hand, the inspection can be performed at some given time instants to identify the different degradation states. As shown in Figure 1, we distinguish from four degradation states in this case study.

Table 4 Safety barriers in the target system [2].

Safety barrier	Failure probability (P_f)	Description
Recirculation pump (SB_1)	1.96×10^{-3}	Once the plant fails to scram, the recirculation pump is activated and used to limit power generation of the NPP.
Safety valve (SB_2)	1.01×10^{-5}	Safety valves are opened to prevent over-pressurization of the reactor.
Boron injection (SB_3)	1×10^{-5}	Liquid boron should be injected manually by the operator within the allowable time to shut down the reactor safely.
Automatic Depressurization System (ADS) inhibit (SB_4)	1.37×10^{-2}	ADS is designed to decrease the pressure of the reactor in order to start the low-pressure system.
Early high-pressure makeup (SB_5)	8.45×10^{-2}	The system is supposed to work automatically when automatic actuation alarm appears, indicating that the water level is lowering to level 2.
Long-term high- pressure makeup (SB_6)	2.13×10^{-3}	The long-term high-pressure system is used to maintain the water level in the vessel 24 hours after the start.
Manual reactor depressurization (SB_7)	$P_{f,SB_7}^{(1)} = 0.45,$ $P_{f,SB_7}^{(2)} = 0.9$	The operator depressurizes the vessel manually to avoid core melt-down. In $SE_{04} - SE_{09}$, the failure probability is $P_{f,SB_7}^{(1)}$, whereas, in $SE_{10} - SE_{15}$, the failure probability is $P_{f,SB_7}^{(2)}$.
Reactor inventory makeup at low pressure (SB_8)	$P_{f,SB_8}^{(1)} = 1.12 \times 10^{-6},$ $P_{f,SB_8}^{(2)} = 3.4 \times 10^{-6},$ $P_{f,SB_8}^{(3)} = 9.49 \times 10^{-5}$	If the low pressure system fails as well as the high-pressure system, then the reactor inventory makeup at lower pressure needs to be activated. In $SE_{04} - SE_{07}$, the failure probability is $P_{f,SB_8}^{(1)}$, while, in $SE_{10} - SE_{14}$, the failure probability is $P_{f,SB_8}^{(2)}$. In $SE_{16} - SE_{20}$, the failure probability is $P_{f,SB_8}^{(3)}$.
Vessel overfill prevention (SB_9)	0.875	The operator needs to monitor the water level and make sure the level is not too high to cause core melt-down.
Long-term heat removal (SB_{10})	2.03×10^{-5}	The long-term heat removal system is initialized to cool down the suppression pool and containment in order to maintain the other supporting systems in working states.

Vessel inventory makeup after containment (SB_{11})	0.4	This measure supplies the proper amount of water to protect the fuel from melting when containment failure happens.
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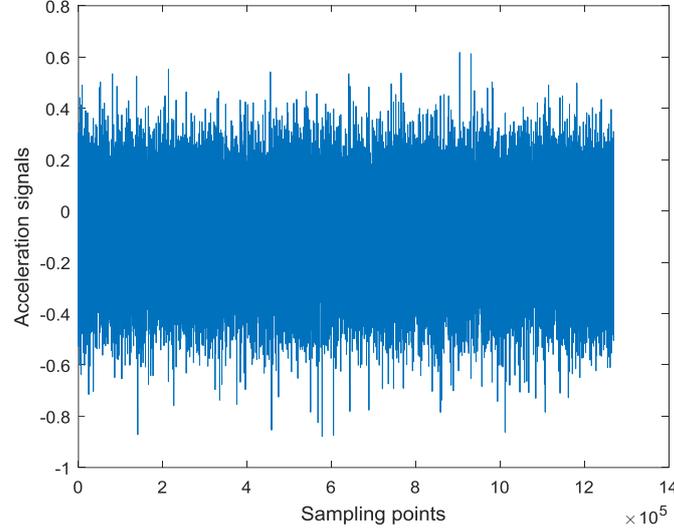


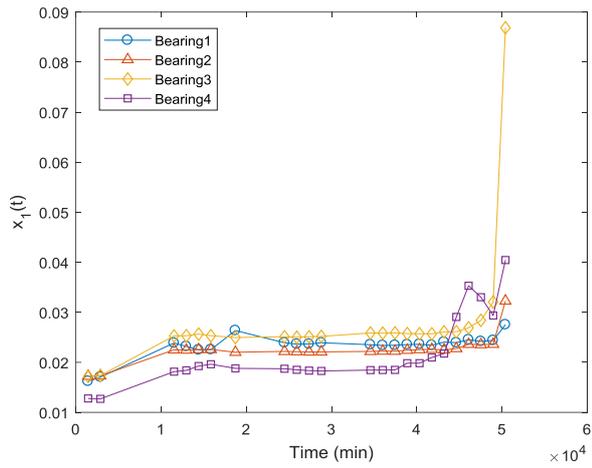
Figure 12 Raw data for the bearing 1 in the test #1 at 10 minutes.

6.2 Dynamic risk assessment

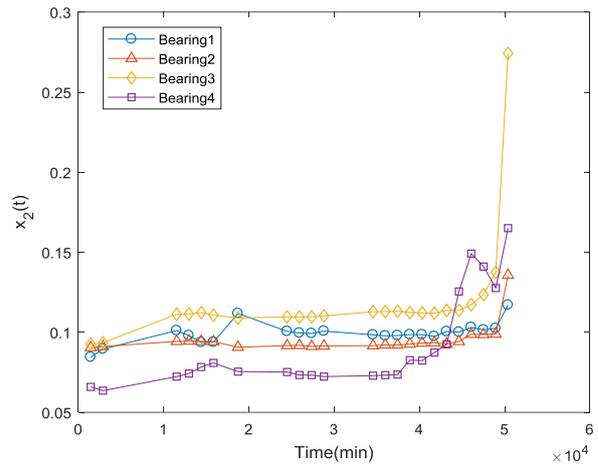
DRA of the ATWS is carried out following the procedures in Figure 6, where the real data set from [45] is used as historical training data. In the offline step, feature extraction needs to be conducted first. Three features are extracted from the vibration signals using the time domain method:

$$\begin{cases} x_1(t_i) = \frac{1}{(t_i - t_{i-1}) \cdot f} \sum_{j \in (t_{i-1}, t_i)} c_j^2 \\ x_2(t_i) = \sqrt{\frac{1}{(t_i - t_{i-1}) \cdot f} \sum_{j \in (t_{i-1}, t_i)} (c_j - \bar{c})^2} \\ x_3(t_i) = \frac{1}{(t_i - t_{i-1}) \cdot f} \sum_{j \in (t_{i-1}, t_i)} c_j \end{cases} \quad (31)$$

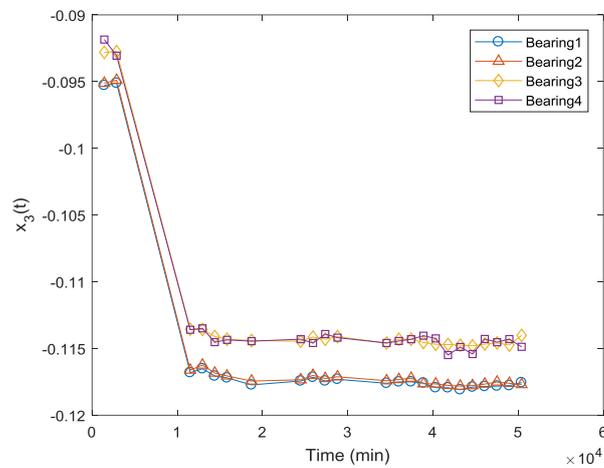
where x_1 is the average power of vibration, x_2 is the root mean square, x_3 is the mean value of vibration. In (31), f is the sampling frequency, $(t_i - t_{i-1}) \cdot f$ is the number of sampling points in time interval $[t_{i-1}, t_i]$, and c_j is the vibration signal. The extracted degradation indicators are shown in Figure 13.



(a) $x_1(t)$: average power of vibration



(b) $x_2(t)$: root mean square



(c) $x_3(t)$: mean value of vibration

Figure 13 Extracted degradation indicators.

Algorithm 1 is applied to train a HM-GMM with four discrete degradation states based on the extracted degradation indicators:

$$\begin{aligned}
A &= \begin{pmatrix} 0.5 & 0.5 & 0 & 0 \\ 0 & 0.9354 & 0.0646 & 0 \\ 0 & 0 & 0.9565 & 0.0435 \\ 0 & 0 & 0 & 1 \end{pmatrix}, \pi = [1 \ 0 \ 0 \ 0], \\
\boldsymbol{\mu} &= \begin{pmatrix} 0.0412 & 0.1176 & 0.2002 & 1.0000 \\ 0.0916 & 0.1184 & 0.2634 & 1.0000 \\ 0.0579 & 0.9168 & 0.8672 & 0.8446 \end{pmatrix}, \\
\boldsymbol{\Sigma}^{(1)} &= \begin{pmatrix} 0.0108 & 0.0018 & 0.0007 \\ 0.0018 & 0.0137 & 0.0014 \\ 0.0007 & 0.0014 & 0.0121 \end{pmatrix}, \boldsymbol{\Sigma}^{(2)} = \begin{pmatrix} 0.0111 & 0.0020 & 0.0012 \\ 0.0020 & 0.0134 & 0.0019 \\ 0.0012 & 0.0019 & 0.0137 \end{pmatrix}, \\
\boldsymbol{\Sigma}^{(3)} &= \begin{pmatrix} 0.0129 & 0.0039 & 0.0002 \\ 0.0039 & 0.0153 & 0.0002 \\ 0.0002 & 0.0002 & 0.0106 \end{pmatrix}, \boldsymbol{\Sigma}^{(4)} = \begin{pmatrix} 0.01 & 0 & 0 \\ 0 & 0.01 & 0 \\ 0 & 0 & 0.01 \end{pmatrix}.
\end{aligned} \tag{32}$$

The online condition monitoring data are generated using the bootstrap sampling: 10^4 bootstrap samples are generated from the training data set. A HM-GMM $\hat{\lambda}$ is, then, trained based on these samples using Algorithm 1:

$$\begin{aligned}
A &= \begin{pmatrix} 0.5 & 0.5 & 0 & 0 \\ 0 & 0.9613 & 0.0387 & 0 \\ 0 & 0 & 0.7150 & 0.2849 \\ 0 & 0 & 0 & 1 \end{pmatrix}, \pi = [1 \ 0 \ 0 \ 0], \\
\boldsymbol{\mu} &= \begin{pmatrix} 0.0446 & 0.1338 & 0.2339 & 0.7809 \\ 0.0974 & 0.2004 & 0.3087 & 0.8062 \\ 0.0764 & 0.9744 & 0.8779 & 0.8584 \end{pmatrix}, \\
\boldsymbol{\Sigma}^{(1)} &= \begin{pmatrix} 0.0105 & 0.0010 & 0.0005 \\ 0.0010 & 0.0122 & 0.0009 \\ 0.0005 & 0.0009 & 0.0118 \end{pmatrix}, \boldsymbol{\Sigma}^{(2)} = \begin{pmatrix} 0.0109 & 0.0016 & 0.0007 \\ 0.0016 & 0.0128 & 0.0010 \\ 0.0007 & 0.0010 & 0.0128 \end{pmatrix}, \\
\boldsymbol{\Sigma}^{(3)} &= \begin{pmatrix} 0.0123 & 0.0030 & 0.0000 \\ 0.0030 & 0.0141 & -0.0001 \\ 0.0002 & -0.0001 & 0.0105 \end{pmatrix}, \boldsymbol{\Sigma}^{(4)} = \begin{pmatrix} 0.0111 & 0.0013 & 0.0001 \\ 0.0013 & 0.0116 & 0.0001 \\ 0.0001 & 0.0001 & 0.0100 \end{pmatrix}.
\end{aligned} \tag{33}$$

The HM-GMM λ in (33) is, then, treated as the true degradation model and used to generate the condition monitoring data for the bearing that is monitored in the online step. The generated condition monitoring data are shown in Figure 14.

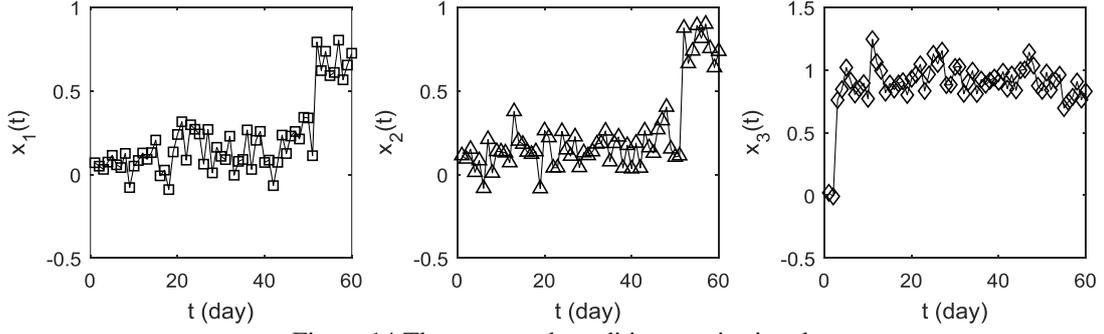


Figure 14 The generated condition monitoring data.

Inspections are conducted at three time instants, i.e., $t = 30(d)$, $t = 35(d)$ and $t = 50(d)$, respectively. The inspection data at the three time instants are given in Table 5. In Table 5, we also show the true degradation states obtained from the true degradation model in (33) and the estimated degradation states using condition monitoring data and Algorithm 2.

The estimated degradation state S_{IN} and S_{CM} are, then, integrated using (19). Note that in (17), the reliability of the inspection data is set to $R_{IN} = 0.8$. Then, the value of $P(S_{IN} | S)$ in (19) can be derived easily from (17). The values of $P(S_{CM} | S)$ are assigned by considering the distance between the neighboring degradation states: the closer the states are, the more likely a misclassification might happen. For example, the normalized distance between S_2 and S_3 is:

$$\frac{d(\boldsymbol{\mu}_2, \boldsymbol{\mu}_3)}{\sum_{i=1}^4 d(\boldsymbol{\mu}_i, \boldsymbol{\mu}_3)} = 0.4807, \quad (34)$$

and the normalized distance between S_3 and S_4 is:

$$\frac{d(\boldsymbol{\mu}_4, \boldsymbol{\mu}_3)}{\sum_{i=1}^4 d(\boldsymbol{\mu}_i, \boldsymbol{\mu}_3)} = 0.1108, \quad (35)$$

where $d(\cdot)$ is the Euclidean distance. Thus, we set $P(S_{CM} = S_2 | S = S_3) = 0.1$ and $P(S_{CM} = S_4 | S = S_3) = 0.2$. The values of the other elements in $P(S_{CM} | S)$ are determined in a similar way and reported in Table 6. Once the integrated estimation of the degradation state is obtained, risk updating and prediction can be performed by (23) and (24), respectively.

Table 5 Values of S, S_{CM} and S_{IN} at different time instants.

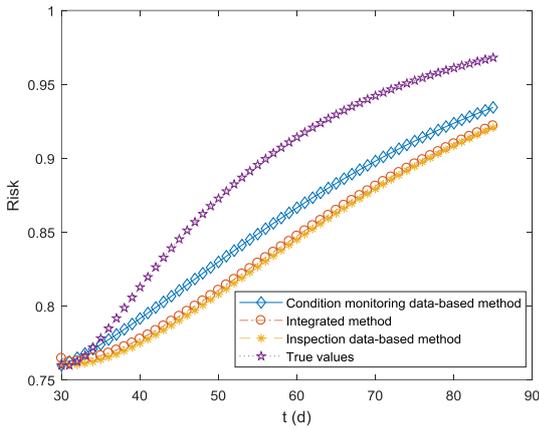
	$t = 30(d)$	$t = 35(d)$	$t = 50(d)$
S	S_2	S_3	S_3
S_{CM}	S_2	S_2	S_3
S_{IN}	S_2	S_3	S_2

Table 6 Values of $P(S_{CM} | S)$.

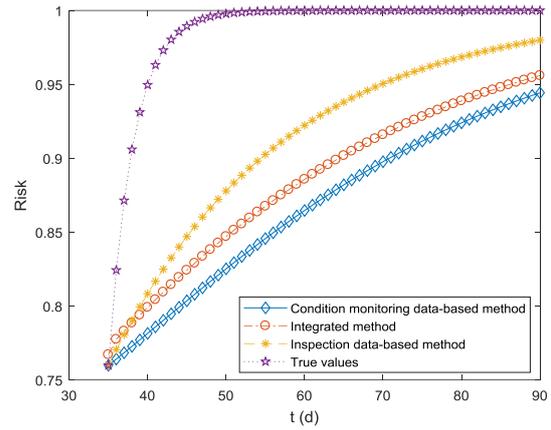
	$S = S_1$	$S = S_2$	$S = S_3$	$S = S_4$
$P(S_{CM} = S_1 S)$	0.9	0	0	0
$P(S_{CM} = S_2 S)$	0.05	0.9	0.1	0.1
$P(S_{CM} = S_3 S)$	0.05	0.1	0.9	0.1
$P(S_{CM} = S_4 S)$	0	0	0	0.8

6.3 Results and discussion

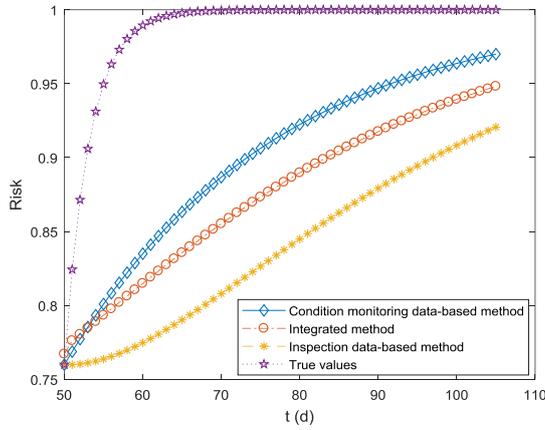
The results of risk updating and prediction at $t = 30, 35$ and $50(d)$ are given in Figure 15. In Figure 15, we also show the results from using only condition monitoring data and inspection data, for comparison.



(a) $t = 30(d)$



(b) $t = 35(d)$



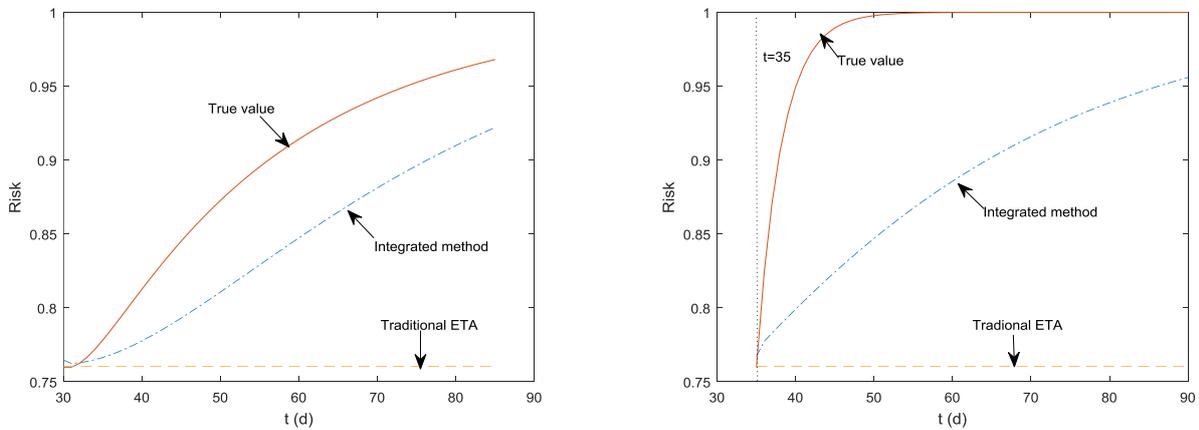
(c) $t = 50(d)$

Figure 15 The results of risk updating and prediction.

As shown in Figure 15(a), at $t = 30(d)$, the results from all the three methods are close to each other. This can be explained from Table 5: at $t = 30(d)$, both data sources correctly identify the true degradation states. However, when compared to the true risk values, the updated and predicted risks from all the three methods show relatively large discrepancies. This discrepancy is mainly due to the estimation errors in the offline step (see (32) and (33)), as we have only four samples in the training data set. A possible way to increase the accuracy of risk updating is, then, to increase the sample size of the training data in the offline step.

It can be seen from Table 5 that at $t = 35(d)$, the inspection data give correct information on the current degradation state while condition monitoring data do not. From Figure 15(b), it can be seen that the developed data-integration method improves the DRA results from the condition monitoring data-based method, as it integrates the correct information from inspection data. On the other hand, when the inspection data fail to give the correct information ($t = 50(d)$), it can be seen from Figure 15(c) that the developed data integration method can also correct the misleading results obtained from using only the inspection data. Hence, in general, applying the developed data integration method can achieve a more robust DRA result than using the two data sources individually.

In Figure 16, we compare the developed DRA method with the conventional ETA method in [2]. It can be seen from Figure 16 that the results from the developed DRA method are closer to the true risk values than those of the standard ETA. This is because through the integration of inspection and condition monitoring data, the developed method is able to capture the time-dependent behavior of the recirculation pump resulting from the degradation of the bearing. The standard ETA, however, fails to capture such time-dependencies as it assumes that the event probabilities do not change although the real system/component ages over time.



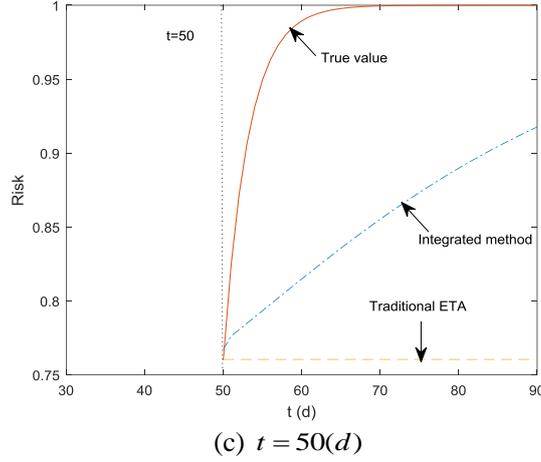
(a) $t = 30(d)$ (b) $t = 35(d)$ 

Figure 16 Comparisons of traditional ETA and DRA.

Additionally, as can be seen from Figure 16, the true risk is higher than the one estimated by the developed method. The inaccuracy of the risk estimation is caused by the imprecise estimation of the parameters in the HM-GMM (equation 32), which is primarily due to the small sample size in the offline training of the HM-GMM (see Figure 4). It can be seen from equations (32) and (33) that, since we have only four samples in the offline training phase, the estimated transition probability differs from its true value. Particularly, the probability of system remaining in S_3 given that it enters S_3 is estimated to be $a_{33} = 0.9565$, which is larger than its true value $a_{33} = 0.7150$. This indicates that the trained HM-GMM tends to overestimate the reliability of the safety barrier (S_4 is the failure state), and, hence, underestimate the risk, in this case. The inaccuracy of the estimation is caused by the fact that we have only four samples from the real dataset, for the offline training phase. In the numerical case study (Section 5), it is shown that with 10^4 training samples, the estimation accuracy is satisfactory.

A major issue with the EM algorithm (Algorithm 1) is that, when the sample size is small, there can be large uncertainty on the estimated parameter values. This uncertainty, if not properly addressed, might greatly impact the estimation accuracy of the reliability of the safety barriers, and, then, the calculated risk. One way to capture the uncertainty in the estimated risk caused by parameter estimation is to conduct a bounding risk analysis by using Bayesian inference [20, 46, 47], where posterior distributions of the parameters, rather than point estimators, are calculated to represent the parametric uncertainty. The uncertainty in the parameter estimation can be represented in terms of the credible intervals. By propagating the parametric uncertainty, a credibility interval can also be obtained for the estimated risk, which can help the decision-makers understand the confidence on the risk estimations.

7. Conclusions

In this paper, a framework has been presented to integrate condition monitoring data and inspection data for DRA. A HM-GMM has been developed to estimate the degradation states of the safety barriers based on the condition monitoring data. The estimated degradation states are integrated with the inspection data for DRA by a BN model. A numerical case study and a real-world application on a NPP accident risk assessment model (an ET) have been conducted. The results show that, as expected, integrating the two data sources into the DRA gives more accurate and robust results than using any one of the two individual data sources.

There are some challenges to be addressed when applying the developed model to real-life large-scale systems (of systems). The first one is that, to ensure the accuracy of the developed method, a sufficient number of training samples is needed. This might not be the case for real-world systems. To address this challenge, the estimation of the values of the parameters of the HM-GMM can be embedded within a Bayesian inference framework for a bounding analysis that gives due account to uncertainties. Other future developments should consider the extension of the developed model to systems with multiple degrading components and repairable components.

The current method only considers a discrete time discrete state Markov model as the degradation model. A future work is to extend the developed framework to other degradation models, e.g. the Brownian motion model [48], Gamma process model [49], etc. Moreover, in the current framework, the parameters of HM-GMM are estimated offline; in the future, online updating of the parameters can be considered in order to improve the accuracy of the DRA.

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Paper II

J. Xing, Z. Zeng, E. Zio. Dynamic business continuity assessment using condition monitoring data. *International Journal of Disaster Risk Reduction*, 2019, 41, 101334.

Dynamic business continuity assessment using condition monitoring data

Jinduo Xing ¹, Zhiguo Zeng ¹, Enrico Zio ^{2,3,4}

¹ Chair System Science and the Energy Challenge, Fondation Electricité de France (EDF), CentraleSupélec, Université Paris Saclay, Gif-sur-Yvette, France

² MINES ParisTech, PSL Research University, CRC, Sophia Antipolis, France

³ Energy Department, Politecnico di Milano, Milan, Italy

⁴ Eminent Scholar, Department of Nuclear Engineering, College of Engineering, Kyung Hee University, Republic of Korea

jinduo.xing@centralesupelec.fr, zhiguo.zeng@centralesupelec.fr, enrico.zio@polimi.it

Abstract

Concerns on the impacts of disruptive events of various nature on business operations have increased significantly during the past decades. In this respect, business continuity management (BCM) has been proposed as a comprehensive and proactive framework to prevent the disruptive events from impacting the business operations and reduce their potential damages. Most existing business continuity assessment (BCA) models that numerically quantify the business continuity are time-static, in the sense that the analysis done before operation is not updated to consider the aging and degradation of components and systems which influence their vulnerability and resistance to disruptive events. On the other hand, condition monitoring is more and more adopted in industry to maintain under control the state of components and systems. On this basis, in this work, a dynamic and quantitative method is proposed to integrate in BCA the information on the conditions of components and systems. Specifically, a particle filtering-based method is developed to integrate condition monitoring data on the safety barriers installed for system protection, to predict their reliability as their condition changes due to aging. An installment model and a stochastic price model are also employed to quantify the time-dependent revenues and tolerable losses from operating the system. A simulation model is developed to evaluate dynamic business continuity metrics originally introduced. A case study regarding a nuclear power plant (NPP) risk scenario is worked out to demonstrate the applicability of the proposed approach.

Keywords

Business continuity management (BCM), Dynamic business continuity assessment (DBCA), Condition monitoring, Prognostic and health management (PHM), Particle filtering (PF), Event tree (ET)

Acronyms

BCA	business continuity assessment
BCM	business continuity management
BCV	business continuity value
DBC	dynamic business continuity
DBCA	dynamic business continuity assessment
DRA	dynamic risk assessment
ET	event tree
MBCO	minimum business continuity objective
MTPD	maximum tolerable period of disruption
NPP	nuclear power plant
PDF	probability density function
PF	particle filtering
PRA	probabilistic risk assessment
RCS	reactor coolant system
RTO	recovery time objective
RUL	remaining useful life
SGTR	steam generator tube rupture

Notation

a	Crack size
$BCV([t, t + T])$	Business continuity value at t with reference to a time horizon T
C_o	Operation cost
C_p	Repayment cost
C_{s1}	First consequence
C_{s2}	Second consequence
D_p	Down payment
$EDBCV$	Expected value of dynamic business continuity at time t

$f(\cdot)$	State function
$f_{ET}(\cdot)$	Event tree model
$h(\cdot)$	Observation function
IN_{tol}	Total investment
L_d	Direct loss
L_{in}	Indirect loss
L_{tol}	Tolerable loss
N_s	Sample size of PF
N_p	Repayment period
$P_{BF}([t, t+T])$	Probability of business failure in $[t, t+T]$
$P_{BI}([t, t+T])$	Probability of business interruption in $[t, t+T]$
P_{ID}	Indirect loss per unit of time
q	Time length of condition monitoring
Q_0	Initial funding
t_{recv}	Recovery time
T	Time length of BC estimation
$\omega_k^{(i)}$	Weight of particle i
ψ	Indicator function
ρ	Interest rate
δ_k	Observation noise at $t = t_k$
λ_{st}	Intensity of rupture event (for static business continuity)
ΔK	Stress intensity factor
$\Delta\sigma$	Stress range

1. Introduction

Business organizations are faced with threats from various disruptive events, such as natural disasters [1, 2], intentional attacks [3] and hardware failures [4], etc. As reported in [5, 6], 43% of the companies that have suffered from severe disruptive events have been permanently closed. Among these companies, around 30% failed within two years. Being prepared for disruptive events, including prevention in pre-event phase and response in post-event phase, is, then, important for modern businesses [7]. This is the reason why business continuity management (BCM) has received increasing attention in recent years as a holistic risk management method to cope with disruptive events [8-12]. BCM is formally defined in [13] as the “holistic management process that identifies the potential threats to an organization and the potential impacts they may cause to business operations those threats, if realized, might cause, and which provides a framework for building organizational resilience with the capability of an effective response that safeguards the interest of its key stakeholders reputation, brand and value-creating activities”. Compared to conventional risk analysis, BCM not only focuses on the hazards and potential impacts, but also considers how to mitigate their consequence and quickly recover from disruptions. In this sense, it provides a framework for building organizational resilience that safeguards the interests of the business stakeholders.

Most existing works mainly discuss BCM from a management perspective [14]. For instance, the necessity and benefit of implementing BCM in a supply chain has been discussed in qualitative terms in [11]. In [15], a framework for the design, implementation and monitoring of BCM programs has been proposed. In [16], the evolution of BCM related to crisis management has been reviewed, in terms of practices and drivers of BCM. In [17], BCM has been compared with conventional risk management methods, showing that BCM considers not only the protection of the system against the disruptive event, but also the recovery process during and after the accident. The importance of reliability and simulation in BCM has been discussed in [18]. In [19], a framework for information system continuity management has been introduced. Standards concerning BCM of the Brazilian gas supply chain have been discussed in [20]. A practice on BCM in Thailand has been reviewed and a few suggestions on BCM approaches have been presented in [21]. In [22], the conceptual foundation of BCM has been presented in the context of societal safety.

For BCM effective deployment, it is necessary to define numerical indexes for the quantitative business continuity assessment (BCA). Numerical indexes have been defined in [13], e.g., maximum tolerable period of disruption (MTPD), minimum business continuity objective (MBCO) and recovery time objective (RTO). In the current practice, these numerical indexes are estimated based on expert judgements. Only a few attempts exist

concerning developing quantitative models to evaluate these numerical indexes based on objective data [22]. For example, a statistical model integrating Cox's model and Bayesian networks has been proposed to model the business continuity process [23]. In [24], a simulation model has been developed to analyze the business continuity of a company considering an outbreak of pandemic disease, where the business continuity is characterized by the operation rate and the plant-utilization rate. In [5], an integrated business continuity and disaster recovery planning framework has been presented and a multi-objective mixed integer linear programming has been used to find efficient resource allocation patterns. In [9], BCM outsourcing and insuring strategies have been compared based on the organization characteristics and the relevant data through a two-step, fuzzy cost-benefit analysis. Moreover, in [10], an enhanced risk assessment framework equipped with analytical techniques for BCM systems has been proposed. Two probabilistic programming models have been developed to determine appropriate business continuity plans, given epistemic uncertainty of input data in [25]. In [26], a new model for integrated business continuity and disaster recovery planning has been presented, considering multiple disruptive incidents that might occur simultaneously. An integrated framework has been developed in [12] for quantitative business continuity analysis, where four numerical metrics have been proposed to quantify the business continuity level based on the potential losses caused by the disruptive events.

Most quantitative BCA models mentioned above are time-static in the sense that the analysis is performed before the system of interest comes into operation, with no further consideration of the changes that occur due to aging and degradation. In particular, in practice, business continuity is influenced by the degradation of safety barriers. On the other hand, the advancing of sensor technologies and computing resources has made it possible to retrieve information on the state of components and systems, by collecting and elaborating condition monitoring data [27, 28]. For example, a condition-based fault tree has been used for dynamic risk assessment (DRA) [29], where the condition monitoring data are used to update the failure rates of specific components and predict their reliability. In [30], a Bayesian reliability updating method has been developed for dependent components by using condition monitoring data. In [4], a holistic framework that integrates the condition monitoring data and statistical data has been proposed for DRA. A sequential Bayesian approach has been developed in [31], for dynamic reliability assessment and remaining useful life prediction for dependent competing failure processes. Usually, information fusion can add value for decision support [32]. A quantitative model for information risks in supply chain has been developed where the proposed model can be updated when new data are available [33].

In this paper, we propose a framework for DBCA that integrates condition monitoring data and allows updating the business continuity analysis using information collected during system operation. The focus of this paper is on “business continuity assessment” rather than “business continuity management”, as we are concerned with developing quantitative models to evaluate the numerical business continuity indexes which are further used in the BCM process. The developed model contributes to the existing research on BCA in three aspects:

- 1) An integrated DBCA model is proposed, which can provide for BCA updating in time.
- 2) New dynamic business continuity metrics are introduced.
- 3) A simulation-based algorithm is developed to calculate the dynamic business continuity metrics.

The remainder of this paper is organized as follows. In Section 2, numerical metrics for DBCA are proposed. An integrated framework of DBCA is developed in Section 3. Section 4 describes the application of the proposed framework on a nuclear power plant (NPP) accident. Section 5 discusses applicability of the proposed DBCA method. Eventually, Section 6 concludes this work.

2. Numerical metrics for dynamic business continuity assessment

A business process is a process of producing products or supporting services by an organization. The business process of an organization can be characterized by a performance indicator, whose value reflects the degree to which the objective of the business is satisfied. For instance, for a NPP, this indicator can be monthly electricity production. As mentioned in Section 1, some numerical indexes exist for quantifying the continuity of a business process (MTPD, MBCO, RTO, etc.) [13]. These numerical indexes, however, focus only on one specific phase of the whole process at a time. For example, RTO focuses only on the post-disruption recovery phase, MBCO focuses only on the post-disruption contingency activities. In this paper, we use the numerical business continuity indexes developed in [12], which are defined in a more integrated sense to cover the whole process, from pre-disruption prevention to post-disruption contingency and recovery.

In the quantitative framework developed in [12], the business continuity is quantified based on the potential losses caused by the disruptive events. The business process is divided into four sequential stages: preventive stage, mitigation stage, emergency stage and recovery stage. Various safety measures are designed in different stages to guarantee the continuity of the business process. Business continuity value (BCV) was formally defined as [12]:

$$BCV([0, T]) = 1 - \frac{L([0, T])}{L_{\text{tot}}} \quad (1)$$

where L denotes the loss in $[0, T]$ from the disruptive event; T is the evaluation horizon for the assessment (e.g., the lifetime of the system); L_{tol} is the maximum loss that can be tolerated by an organization, which manifests system tolerance ability against disruptive events [34]. A negative value of BCV means that L is higher than L_{tol} , which is unacceptable for the targeted system. When $BCV = 0$, it implies that the loss is exactly what the system can maximally tolerate. Regarding $BCV = 1$, it means that no loss has been generated. Equation (1) measures the relative distance to a financially dangerous state by taking into account the possible losses generated by the business disruption. It should be noted that only one business process is considered in this paper, whereas in practice, an organization might be involved in multiple-businesses processes at the same time. For multiple-businesses organizations, the framework developed can be naturally extended based on the potential losses and profits generated by the different business processes.

The business continuity metrics discussed above are time-static in nature. In practice, however, various factors influencing the business continuity are time-dependent. These dynamic influencing factors can be grouped into internal factors and external factors. Internal factors are related to the safety barriers within the system of interest, such as the dynamic failure behavior of the safety barriers (e.g., corrosion [35], fatigue crack [36], and wear [37]). External factors refer to the influence from external environment. For example, variations in the price of products will affect the accumulated revenue of the organization, and, then, the tolerable loss in Equation (1). To consider these factors, the business continuity metrics are extended to the dynamic cases:

$$DBCV([t, t+T]) = 1 - \frac{L([t, T+t])}{L_{tol}(t)}, \quad (2)$$

where t is the time instant when the dynamic business continuity assessment is carried out; $DBCV([t, t+T])$ represents the business continuity value evaluated at time t , for a given evaluation horizon of T ; $L([t, t+T])$ represents the potential losses in $[t, t+T]$; $L_{tol}(t)$ denotes the maximal amount of losses that the company can tolerate at t : beyond that level of losses, it will have difficulties in recovering. It is assumed that once an organization suffer a loss beyond L_{tol} , it is unable to recover from the disruption. The physical meaning of $DBCV$ is the relative distance to a financial dangerous state at time t , by considering the possible losses in $[t, t+T]$ due to business

disruption; it measures the dynamic behavior of business continuity in a time interval of interest $[t, t+T]$. By calculating the DBCV at different t , the dynamic behavior of business continuity can be investigated.

In [12], two kinds of losses need to be considered when calculating $L([t, t+T])$: direct loss and indirect loss. Direct loss, denoted by $L_d([t, t+T])$, represents the losses that are caused directly by the disruptive event, including structural damage of the system. For example, in a NPP leakage event, $L_d([t, t+T])$ includes all equipment damage directly caused by the event. Indirect loss, denoted by $L_{in}([t, t+T])$, is the revenue loss suffered during the shutdown of the plant [38]. Hence, the total loss is calculated by:

$$L([t, T+T]) = L_d([t, t+T]) + L_{in}([t, t+T]). \quad (3)$$

In terms of other types of accident, for instance, workplace accidents, damages to the surroundings, etc. they may also affect the business continuity, but they are not included explicitly in the model developed in this paper. However, the BCA framework proposed can be naturally generalized by including more initiating events in the analysis.

The DBCV defined in (2) is a random variable. Three numerical metrics are, then, proposed for its quantification:

$$EDBCV = E[DBCVC] \quad (4)$$

$$P_{BI}([t, t+T]) = \Pr(BCV < 1, t) \quad (5)$$

$$P_{BF}([t, t+T]) = \Pr(BCV < 0, t) \quad (6)$$

$EDBCV$ is the expected value of the dynamic business continuity value. A higher $EDBCV$ indicates higher business continuity. $P_{BI}([t, t+T])$ represents the probability that at least one disruptive event causes business interruption in time interval $[t, t+T]$; $P_{BF}([t, t+T])$ is the probability that business failure occurs in $[t, t+T]$, i.e., of the event that the losses caused by the disruptive event are beyond L_{tol} . It is assumed that once an organization suffers a loss beyond L_{tol} , it is unable to recover from the disruption. In this work, both of current time t and the estimation horizon T have influences on BCV. We manage to propose a real-time BCA by considering the time-dependent variables.

3. An integrated framework for dynamic business continuity assessment

In this section, we first present an integrated modeling framework for the dynamic business continuity metrics defined in Section 2. Then, particle filtering (PF) is used to estimate the potential loss L_{tol} in real time using condition monitoring data (Section 3.2). The quantification of tolerable losses L_{tol} is, then, discussed in Section 3.3.

3.1 The integrated modeling framework

To model the dynamic business continuity, we make the following assumptions:

- (5) The evolution of the disruptive event is modeled by an event tree (ET). Depending on the states of safety barriers, different consequences can be generated from an initialing event. These consequences can be grouped into different categories based on their severities. Each consequence generates a certain amount of loss. However, it should be noted that different consequences might have the same degree of loss. According to their severities, possible consequences of a disruptive event are classified as $C_i, i = 1, 2, \dots, n$, where n is the number of severity levels. The severity and duration of the business interruption corresponds to different losses.
- (6) Some safety barriers in the ET are subject to degradation failure processes. Condition monitoring data are available for these safety barriers at predefined time instants $t_k, k = 1, 2, \dots, q$.
- (7) The other safety barriers have constant failure probabilities.
- (8) Recovery means repairing the failed component and restarting the business. The time for the recovery from consequence C_i is a random variable $t_{\text{rec},i}$, with a probability density function (PDF) $f_{\text{rec},i}$.

An integrated framework for DBCA is presented in Figure 1. The DBCA starts from collecting condition monitoring data, denoted as c_k , which is collected from sensors and can be used to characterize the degradation states of the component. The degradation of the safety barriers is estimated based on the condition monitoring data and used to update the estimated losses. Then, the potential profits are predicted and used to calculate the tolerable losses. Finally, the dynamic business continuity metrics can be calculated.

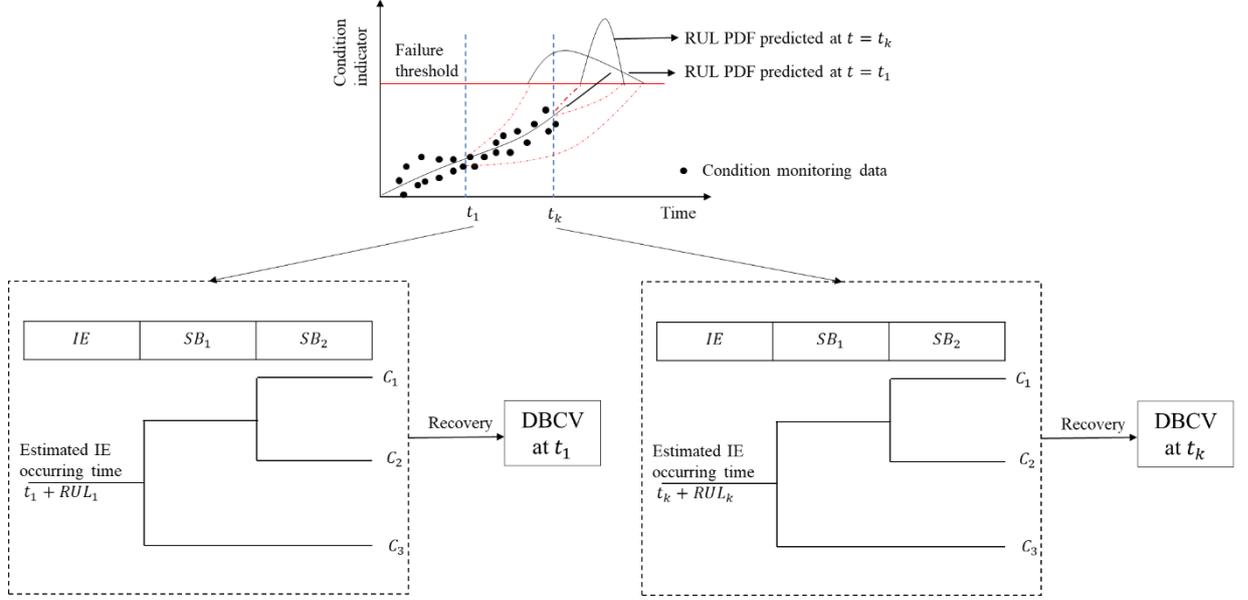


Figure 1. Integrated modeling framework for DBCA.

3.2 Loss modeling

To capture the dynamic failure behavior of a safety barrier as it ages in time, PF is employed in this work to estimate its degradation and predict its remaining useful life (RUL) based on condition monitoring data [39-41]. PF is applied because of its capability of dealing with the complex non-linear dynamics and non-Gaussian noises that are often encountered in practice [42, 43].

Suppose the degradation process of a safety barrier can be described by Equation (7), in which the current state x_k at the k -th discrete time step depends on the previous state x_{k-1} . Here, f is a non-linear function and v_k represents process noise that follows a known distribution. In practice, Equation (7) is often determined based on physics-of-failure models [39]:

$$\mathbf{x}_k = f(\mathbf{x}_{k-1}, v_k) \quad (7)$$

A sequence of condition monitoring data \mathbf{z}_k is assumed to be collected at predefined time points t_k . The sequence of measurement values is assumed to be described by an observation function:

$$\mathbf{z}_k = h(\mathbf{x}_k, \boldsymbol{\sigma}_k) \quad (8)$$

where h is the observation function (possibly nonlinear), $\boldsymbol{\sigma}_k$ is the observation noise vector sequence of known distribution. The measurement data \mathbf{z}_k are assumed to be conditionally independent given the state process \mathbf{x}_k .

Equation (8) quantifies the observation noise from the sensors.

The PF follows two steps [44]:

- 3) Filtering step, where the available condition monitoring data z_k are used to estimate the current degradation state of the system.
- 4) Prediction step, in which the RUL is predicted based on the estimated degradation state and the condition monitoring data.

In the filtering step, the posterior PDF of variable \mathbf{x}_k is approximated by the sum of weighted particles

$\{\mathbf{x}_k^{(i)}, \omega_k^{(i)}\}$:

$$p(\mathbf{x}_k | z_1, z_2, \dots, z_k) \approx \sum_{i=1}^{N_s} \omega_k^{(i)} \delta(\mathbf{x}_k - \mathbf{x}_k^{(i)}) \quad (9)$$

where $p(\mathbf{x}_k | z_1, z_2, \dots, z_k)$ is the estimated posterior PDF of \mathbf{x}_k , δ is the Dirac Delta function, $\omega_k^{(i)}$ is the weight assigned to particle $\mathbf{x}_k^{(i)}$ and is generated by sequential importance sampling [32]. When the new measurement z_k is available, the required posterior distribution of the current state x_k can be obtained by updating the prior distribution:

$$p(\mathbf{x}_k | \mathbf{z}_k) = \frac{p(z_k | \mathbf{x}_k) p(\mathbf{x}_k | \mathbf{z}_{k-1})}{\int p(z_k | \mathbf{x}_k) p(\mathbf{x}_k | \mathbf{z}_{k-1}) d\mathbf{x}_k} \quad (10)$$

where $p(z_k | \mathbf{x}_k)$ is the likelihood function that can be derived from the observation function (8). Generally, if the samples $\mathbf{x}_k^{(i)}$ are drawn from the sampling distribution $p(\mathbf{x}_k | \mathbf{z}_k)$, then, the particle weight can be updated with a new observation z_k , as follows [32]:

$$\omega_k^{(i)} = \omega_{k-1}^{(i)} \frac{p(z_k | \mathbf{x}_k^{(i)}) p(\mathbf{x}_k^{(i)} | \mathbf{x}_{k-1}^{(i)})}{p(\mathbf{x}_k^i | \mathbf{x}_{0:k-1}^i, \mathbf{z}_k)}. \quad (11)$$

Note that the weights are normalized as $\sum_{i=1}^{N_s} \omega_k^{(i)} = 1$.

Algorithm 1 summarizes the major steps of PF [45].

Algorithm 1: Procedures of PF.

Inputs: $\{\mathbf{x}_{k-1}^{(i)}, \omega_{k-1}^{(i)}, \mathbf{z}_k\}$

Outputs: $\{\mathbf{x}_k^{(i)}, \omega_k^{(i)}\}_{i=1}^{N_s}$

For $i = 1$ to N_s do

$$\mathbf{x}_k^{(i)} \sim p(\mathbf{x}_k | \xi_{k-1}^{(i)}) \text{ using (7),}$$

$$\omega_k^{(i)} \sim p(z_k | \mathbf{x}_k^{(i)}, \theta_k^{(i)}) \text{ using (11),}$$

End for

For $i = 1$ to N_s do

$$\omega_k^{(i)} \leftarrow \omega_k^{(i)} / \sum_{i=1}^{N_s} \omega_k^{(i)}$$

End for

$$N_{eff} \leftarrow \left(\sum_{i=1}^{N_s} (\omega_k^{(i)})^2 \right)^{-1}$$

If $N_{eff} < N_s$ then

$$\{\mathbf{x}_k^{(i)}, \omega_k^{(i)}\}_{i=1}^{N_s} \leftarrow \text{resample} \left(\{\mathbf{x}_k^{(i)}, \omega_k^{(i)}\}_{i=1}^{N_s} \right)$$

End if

Return $\{\mathbf{x}_k^{(i)}, \omega_k^{(i)}\}_{i=1}^{N_s}$

Then, in the prediction step, the RUL associated to the i -th particle at $t = t_k$ can be estimated through state function (7) by simulating the evolution trajectory of the particles until they reach the failure threshold z_{th} :

$$RUL_k^{(i)} = \left\{ (T_{th}^{(i)} - 1 - k) \mid x_{T_{th}^{(i)}-1} < z_{th}, x_{T_{th}^{(i)}} \geq z_{th} \right\}, \quad (12)$$

where $T_{th}^{(i)}$ is the first time the particle reaches the threshold z_{th} . Thus, the PDF of the RUL can be generated by:

$$p(RUL | \mathbf{z}_k, z_{th}) \approx \sum_{i=1}^{N_s} \omega_k^{(i)} \delta(RUL - RUL_k^{(i)}). \quad (13)$$

The predicted $RUL_k^{(i)}, i = 1, 2, \dots, N_s$ can, then, be used in a simulation process to generate samples of the total loss L , according to Equation (3). The procedures are summarized in Algorithm 2, where P_{ID} is the indirect loss per unit of time.

Algorithm 2: Generating samples for the losses

Input: $\{RUL_k^{(i)}, \omega_k^{(i)}\}_{i=1}^{N_s}, T$

Output: $L_k^{(i)}$

Initial value $L_k^{(i)} = 0, t = 0, t_1 = 0, T = t_k + T, t_2 = 0;$

$RUL_{pseudo,k} \leftarrow$ randomly select one element from $\{RUL_k^{(i)}\}_{k=1}^{N_p}$, where $RUL_k^{(i)}$ is selected with probability $\omega_k^{(i)}$;

Calculate $T_k^{(i)} = t_k + RUL_{pseudo,k}$

```

► While  $t < T$ 
   $t_1 = t; t_1 = t_1 + TTF_k^{(i)}$ ;
  ► if  $t_1 > T$ 
     $L_k^{(i)} = L_k^{(i)}$ 
  else
    Using the event tree determine the consequence;
    Using the  $f_{recv,i}$  generate the  $t_{recv}$ ;
     $t_2 = t_1 + t_{recv}$ ;
    ► If  $t_2 > T$ 
       $L_k^{(i)} = L_k^{(i)} + L_d + (T - t_2) \cdot P_{ID}$ 
    else  $t = t_2$ 
       $L_k^{(i)} = L_k^{(i)} + L_d + t_{recv} \cdot P_{ID}$ 
    end if
  end if
end if
end while

```

3.3 Tolerable losses modeling

Budget limitations are the primary driver of resilience-enhancing investments [46], which influence protection, prevention, and recovery capabilities of system. Tolerable losses L_{tol} depend on the cash flow of the company and also the risk attitude of the decision maker [13]. In this paper, we assume that at t_k , the organization can tolerate up to α (in percentage) of its cash flow $Q(t_k)$ at t_k :

$$L_{tol}(t_k) = Q(t_k) \cdot \alpha \quad (14)$$

For example, $\alpha = 0.1$ (as assumed in this paper) means that 10% of the current cash flow can be used to withstand potential losses caused by a disruptive event. In practice, the value of α should be determined by the decision maker and reflects his/her risk attitude.

We make the following assumptions to model the dynamic behavior of cash flows:

- (3) At $t = 0$, there is an initial capital of Q_0 .
- (4) Installment is used for the company to purchase the asset, where an equal repayment of C_p is payed each month for N_p months.

It is noteworthy that the cash flow $Q(t)$ depends on the profit earned by the normal operation of the asset:

$$Q(t_k) = Q_0 + I(t_k) - C_o(t_k) - \sum_{i=1}^k (\Psi \cdot C_p(t_i)), \quad (15)$$

where Q_0 is the initial capital, $I(t_k)$ is the accumulated revenues of the organizations up to t_k by selling the product of the asset. For example, in a NPP, $I(t_k)$ is determined by the electricity price ; in oil exploitation, $I(t_k)$ depends on the petroleum price [47]. $C_o(t_k)$ is the operational cost in $[0, t_k]$, which is assumed to be not changing over time. $C_p(t_i)$ is the amount of repayment of the installment in $[t_{i-1}, t_i]$, which can be modeled by (see [48] for details):

$$C_p = \frac{(IN_{\text{tol}} - D_p)}{N_p} (1 + \rho)^{N_p}, \quad (16)$$

where IN_{tol} denotes the total investment and equals the whole value of the system, D_p represents the down payment, ρ is the interest rate, Ψ is an indicator function:

$$\Psi = \begin{cases} 1, & \text{if } t \leq N_p \\ 0, & \text{otherwise} \end{cases}, \quad (17)$$

where N_p is the repayment period.

4. Application

In this section, we consider the development of DBCA in a case study regarding a disruptive initialing event for a NPP [49]. The business continuity of the NPP is evaluated at different ages $t = 1, 2, \dots, 40$ (year) and different evaluation horizons $T = 1, 2, \dots, 60$ (year). The evaluation is made with reference to a specific risk scenario, with the initialing event being the steam generator tube rupture (SGTR).

The targeted system is briefly introduced in Section 4.1. Subsequently, in Section 4.2, the RUL prediction for a SGTR and the modeling of the potential losses are conducted. The time-dependent L_{tol} is calculated in Section 4.3. The results of the DBCA are presented and discussed in Section 4.4.

4.1 System description

For illustrative purposes, it is assumed that the NPP has one reactor with a capacity of 550 MW. It is also assumed that the NPP is subject to the threat of only one disruptive event, the SGTR. The whole value of the NPP is 10^9 € and the operator purchases the NPP using an installment, where the down payment is $5 \cdot 10^8$ € and the repayment period is 10 years with an interest rate of 2%.

SGTR is a potential accident that is induced by the degradation of the tubes in the steam generator, which can lead to tube cracking and rupture [50]. Steam generator tubes transfer the heat from the reactor core to the cooling water that is transformed into steam to drive turbines and produce electricity [49]. The steam generator tube is often manufactured with alloy material to attain high structural integrity and prevent leakage of radioactive materials. An ET has been developed for the probabilistic risk assessment (PRA) of the SGTR for a NPP in South Korea, as shown in Figure 2. In Figure 2, eight safety barriers ($SB_1 \sim SB_8$) are designed to control the accident and mitigate its impact (Table 1). Depending on the states of the safety barriers, 28 sequences are generated ($S_1 \sim S_{28}$). Based on the degree of their severities, the consequence of the sequences can be categorized into two groups. The first group,

$$C_{S1} = \{SE_1, SE_2, SE_4, SE_6, SE_7, SE_9, SE_{11}, SE_{12}, SE_{14}, SE_{16}, SE_{20}, SE_{24}\} \quad (18)$$

represents the event sequences in which a SGTR occurs but the consequence is contained by the safety barriers without causing severe damages. The remaining event sequences form the second group C_{S2} and represent severe consequences of core damage. Regarding C_{S1} , albeit no severe losses are caused, normal production of the NPP is disturbed because the ruptured tube has to be repaired. For C_{S2} , it is assumed that the NPP has to be shut down permanently and the losses incurred are denoted by C_{CD} .

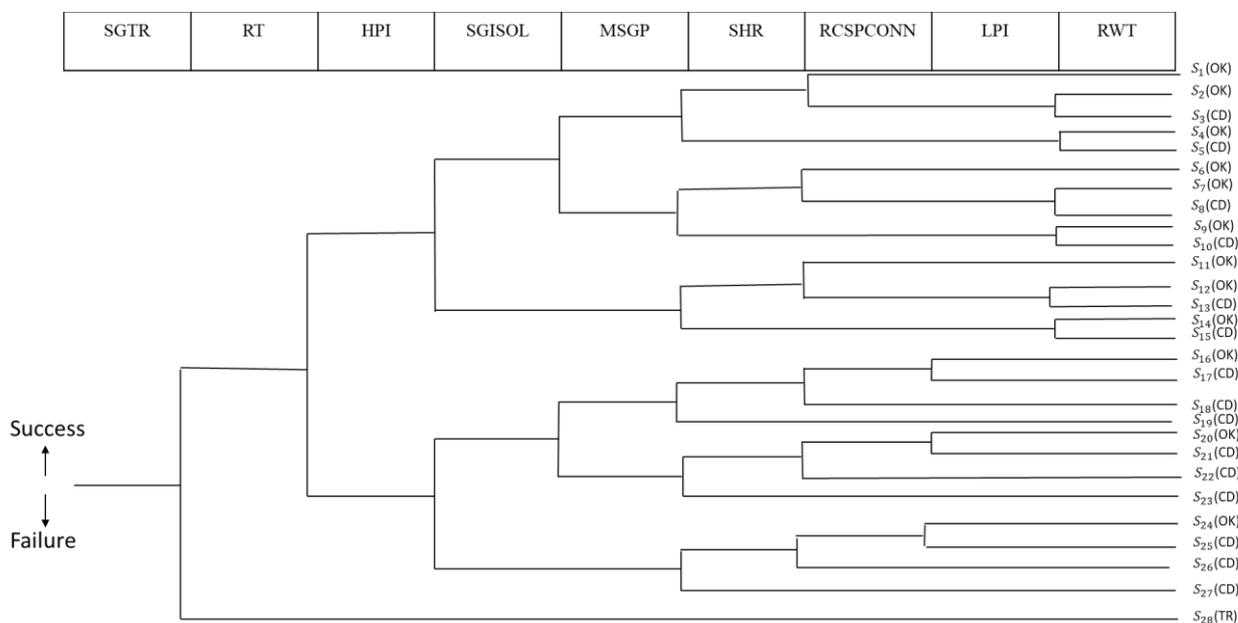


Figure 2. ET for SGTR accident initialing event [49].

Table 1. Safety barriers in the target system [51, 52].

Safety barrier	Failure probability	Description
Reactor trip (RT)	$P_{RT} = 1.8 \times 10^{-4}$	When there is off-normal condition, the protection system automatically inserts control rods into the reactor core to shut down the nuclear reaction.
High pressure safety injection (HPI)	$P_{HPI} = 4.6 \times 10^{-4}$	Inject cool water (at a pressure of about 13.79 MPa) into the reactor coolant system (RCS) to cool the reactor core and provide RCS inventory make-up.
Main steam isolation valve (SGISOL)	$P_{SGI} = 1.0 \times 10^{-4}$	A valve used to isolate the affected steam generator (SG).
Maintain the affected SG pressure (MSGP)	$P_M = 1.5 \times 10^{-4}$	Maintain the affected SG pressure through the pressurizer.
Secondary heat removal (SHR)	$P_{SHR} = 3.4 \times 10^{-5}$	Heat removal by unaffected SG.
Reactor coolant system pressure control (RCSPCON)	$P_{RCSM} = 1.0 \times 10^{-2}$	Open the turbine bypass valve to control the secondary side pressure.
Low pressure safety injection (LPI)	$P_{LPI} = 4.6 \times 10^{-4}$	Inject cool water (at a pressure of about 1.03MPa) to cool down the RCS and provide RCS inventory make-up.
Refill RWT (RWT)	$P_{RWT} = 2.4 \times 10^{-8}$	Refill water storage tank.

The crack growth process that leads to SGTR can be monitored through non-destructive inspection (e.g., ultrasonic testing [53], eddy current testing [54]). In practice, this is done during planned shutdowns of the NPP, often during the refueling stage. The condition monitoring data collected from these inspections are, then, used for the dynamic business continuity assessment.

4.2 Particle filtering and loss modeling

The first step is to update the occurrence probability of the initiating event, based on the condition monitoring data. Note that, due to the lack of real data, the condition monitoring data employed in the case study is generated from a known physical model. For illustrative purposes, the evolution of the tube crack growth process is assumed to follow the Paris-Erdogan model, which has been applied to model SGTR in [52, 55],

$$\frac{da}{dt} = C(\Delta K)^m, \Delta K = \Delta\sigma\sqrt{\pi a}, \quad (19)$$

where a is the crack length, C and m are constant parameters related to the component material properties, ΔK is the stress intensity factor, $\Delta\sigma$ is the stress range. The model can be rewritten in the form of a state transition function [56]:

$$a_k = C_k (\Delta\sigma\sqrt{\pi a_k})^{m_k} dt + a_{k-1} \quad (20)$$

The crack size a_k at $t=t_k$ is obtained from non-destructive inspection, such as ultrasonic testing; the corresponding observation z_k is:

$$z_k = a_k + \delta_k, \quad (21)$$

where δ_k is the observation noise with $\delta_k \sim N(0, \delta_o^2)$.

Due to environment and measurement noises, the measured crack lengths are different from the true values. In this paper, we generate the true values of the crack in Figure 3 using a theoretical model with known parameters and generate the observation data by adding a random noise. The purpose of using PF is to estimate the true crack length from the noised observation data and predict the RUL. The number of particles simulated is $N_s = 5000$. It should be noted that for the tube degradation process, the state vector \mathbf{x} includes the crack size a and the model parameter variables C , m . The initial values for these variables are drawn uniformly from the intervals of values listed in Table 2:

$$\begin{cases} C_k = C_{k-1} + N(0, \sigma_c^2) \\ m_k = m_{k-1} + N(0, \sigma_m^2) \end{cases} \quad (22)$$

Table 2. Initial intervals for the parameters.

Parameters	Initial interval
C	[0.1, 0.2]
m	[1.1, 1.3]
σ_c	$[0.9 \times 10^{-3}, 0.2 \times 10^{-2}]$
σ_m	$[0.9 \times 10^{-3}, 0.2 \times 10^{-2}]$
σ_o	[0.65, 0.85]

The results of PF are shown in Figure 4, where we find that the RUL prediction results become more accurate when more condition monitoring data are available.

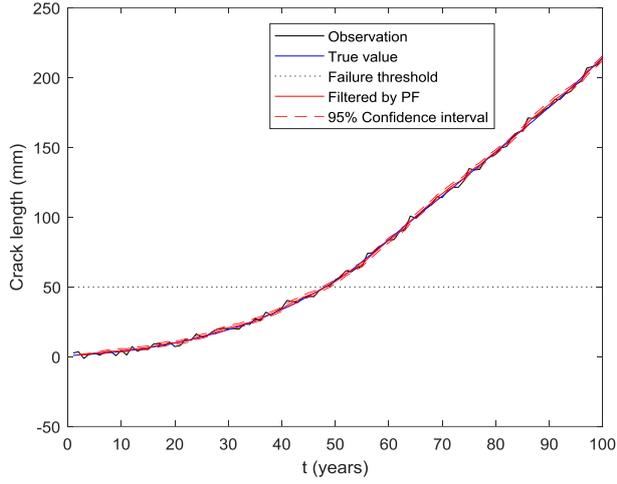


Figure 23. Crack growth process.

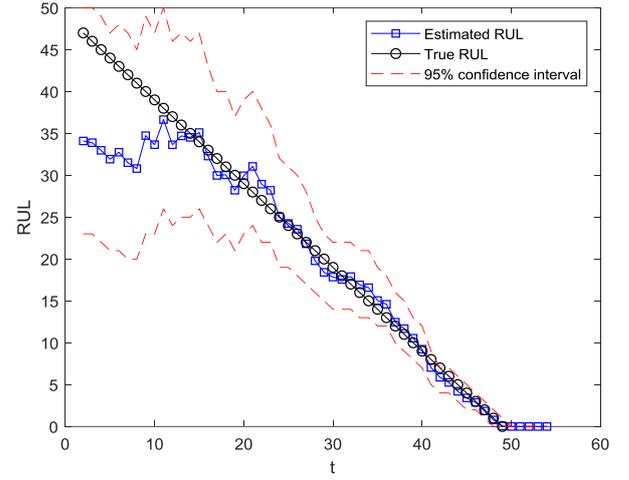


Figure 24. RUL prediction results.

Afterwards, the loss $L([t, t + T])$ in Equation (2) can be calculated. The losses caused by a SGTR event, include the direct losses and indirect losses. In this case study, the direct losses, denoted by L_d , equal to the value of the damaged equipment. For the consequence C_{S1} , L_d is identical to the value of the ruptured tube. For the consequence C_{S2} , L equals the value of the NPP production since the NPP has to be shutdown. In this paper, we assume that if C_{S2} occurs, we have $L = 5 \cdot 10^9 \text{ €}$ [57].

The indirect losses L_{in} are calculated considering the revenue losses during the recovery process, which depends on the recovery time and electricity price. Due to the common use of lognormal distribution for modeling the repair process [58-60], we also assume that the recovery time follows a lognormal distribution with the parameters summarized in Table 3, where ε and β are parameters of the lognormal distribution, whose PDF is

$$f(t_{recv}) = \begin{cases} \frac{1}{\sqrt{2\pi}\beta t_{recv}} e^{-\frac{(\ln(t_{recv})-\varepsilon)^2}{2\beta^2}}, & t_{recv} > 0 \\ 0, & t_{recv} \leq 0. \end{cases} \quad (23)$$

Then, the value of L_{in} is calculated by Monte Carlo simulation.

Table 3. Values of the recovery model parameters.

Parameter	Description	Value
ε	The mean value of the lognormal distribution.	1 year
β	The variance value of the lognormal distribution.	0.1 year ²

4.3 Tolerable loss modeling

We assume that the decision-maker of the NPP determines that the organization can tolerate losses up to 10% of the cash flow. Therefore, we have $\alpha = 0.1$. For the NPP, $I(t_k)$ depends on the electricity price, which often exhibits large variabilities. In this paper, we use the following model, as much as possible incorporating the features of electricity price (such as seasonal volatility, time-varying mean reversion and seasonally occurring price spikes) to simulate the stochastic behavior of the electricity price [61]:

$$dx_t = \theta\tau(t)(\mu_p - x_t)dt + \sigma\sqrt{\tau(t)}dW_t + dZ_t \quad (24)$$

where x_t is the electricity price at t , $\theta > 0$ and μ_p is the mean value of the price, W_t is a standard Brownian motion and Z_t is a compound Poisson process with levy measure $\nu(dx) = \lambda g(x)dx$, λ is the jump intensity and g is the density of the jump size distribution, $\tau(t)$ is a positive stochastic process which satisfies:

$$\tau(t) = s(t) + \nu(t) \quad (25)$$

where $s(t)$ is a deterministic, time-dependent and positive seasonal component, which is often modeled by a trigonometric function:

$$S_1(t) = a_1 \sin\left(\frac{a_2 + 2\pi t}{5}\right) + a_3 \left(\frac{a_4 + 2\pi t}{251}\right) + a_5. \quad (26)$$

The value of the seasonal component parameters are shown in Table 4.

Table 4. Values of the seasonal component parameters of the spot prices.

Parameter	Value
a_1	0.41
a_2	1.90
a_3	0.40
a_4	43.11
a_5	0.29

$\nu(t)$ is a stochastic process, representing the stochastic part of the time change. The Cox-Ingersoll-Ross process [62] is used to model $\nu(t)$,

$$d\nu(t) = \kappa(\eta - \nu(t))dt + \sqrt{\nu(t)\sigma_2}dW_2(t). \quad (27)$$

By using Itô's lemma [61], Equation (24) can be solved and we can derive the following form:

$$x(t) = x(0) + \int_0^t \theta(\mu - x(t))dt + \int_0^t \sigma\sqrt{\tau(t)}dB(t) + \int_0^t dZ(t). \quad (28)$$

The parameters of the stochastic electricity model are tabulated in Table 5, which is estimated from the German EEX² (a market platform for energy and commodity products), from 12.03.2009 until 31.12.2013. The interested readers may refer to details and derivations in [61].

Table 5. Parameters in the stochastic electricity model [61].

Parameter	Value
x_0	40
θ	0.22
μ	50
σ	5.98
dt	1
λ	0.12
μ_1	1.02
σ_1	1.35

Eventually, the generated stochastic electricity price trajectory is shown in Figure 5.

² <https://www.eex.com>, accessed 2019-9-12

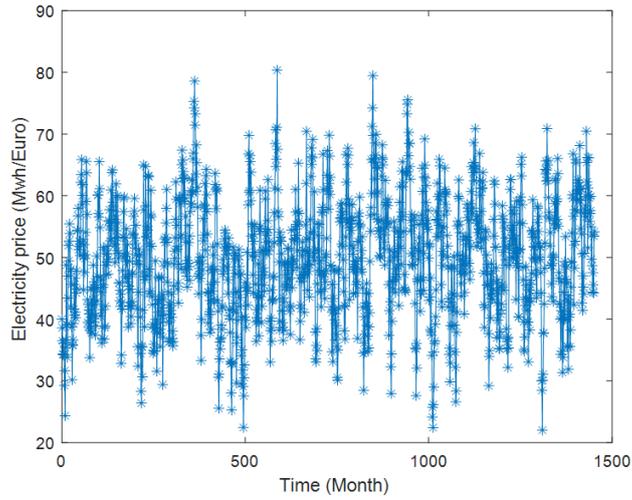


Figure 5. Simulated time-varying electricity price trajectory for 1500 months.

The operation cost $C_o(t_k)$ in Equation (15) is set as constant 20€/MWh, which includes the cost of uranium fuel and the cost of disposing used fuel and wastes [63]. Finally, the cash flow at different time points is shown in Figure 6. We can see that the accumulated profit is small at the beginning. This is because this period is still under the repayment period and a large amount of the revenue is used for repaying the installment. After $t = 10$ years, the repayment is paid off and, thus, the profit increases significantly.

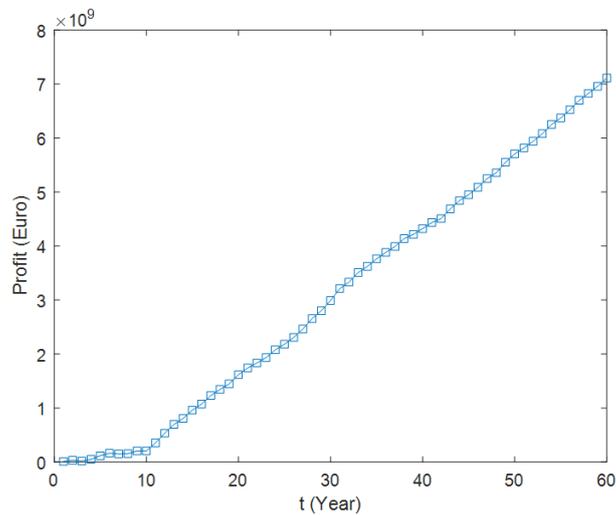


Figure 6. Profit trajectory at different estimation points.

4.4 Results

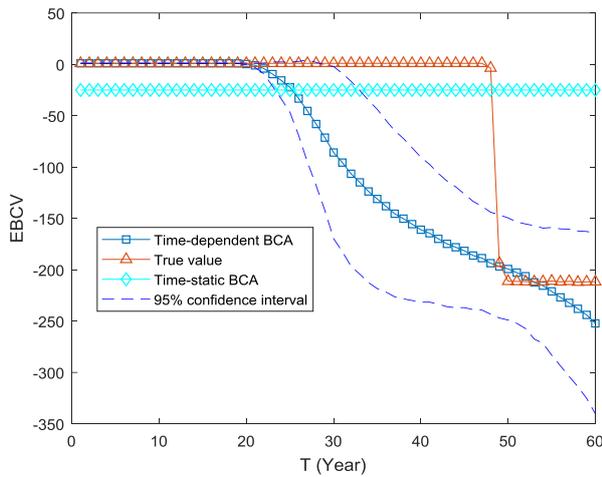
A DBCA is conducted using Algorithm 2. The analyses investigate the dynamic business continuity behavior for the plant at different ages $t = 1, 2, \dots, 40$ (years) and under different evaluation horizons $T = 1, 2, \dots, 60$ (years),

as shown in Figures 7~9. To show the difference between DBCA and (time-static) BCA, a comparison is also carried out. For the BCA, the occurrence of SGTR is assumed to follow a Poisson process, where $\lambda_{st} = 7.0 \times 10^{-3}$ per year [49]. The estimated time horizon is chosen to be the lifetime of the NPP, $T = 60$ years. The time-static business index is defined as:

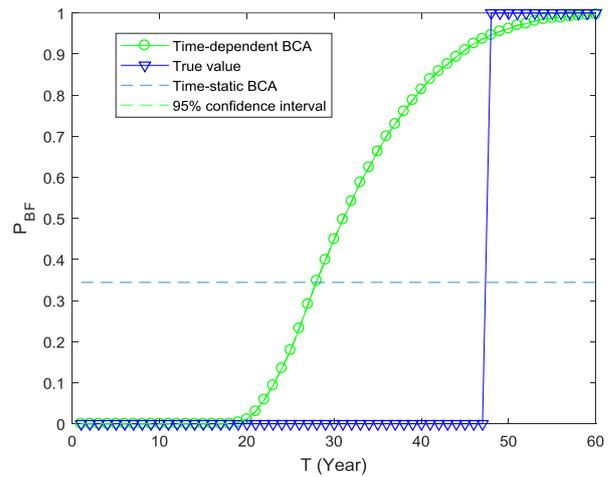
$$BCV(0,T) = 1 - \frac{L(0,T)}{L_{tol}} \quad (29)$$

where BCV is the business continuity value; L_{tol} is the tolerable losses and is assumed to be a constant value, which equals Q_0 (i.e., the initial capital). The recovery time model for the BCA is identical to the one employed in DBCA.

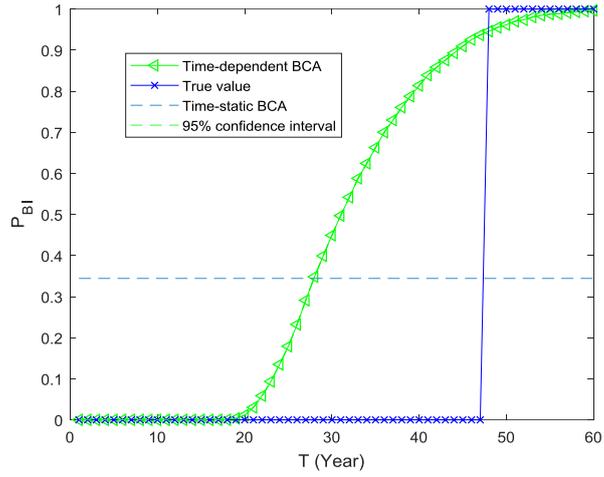
The results from the time-static and time-dependent BCA are compared in Figure 7~9, where the true values are generated based on a theoretical model with known parameters. The abscissa axis shows the estimation horizon T , and the vertical axis stands for the different BCV indexes. Then, the Figures represent the trend of business continuity of NPPs at different age (t), if it is operated for different durations (T).



(a) EDBCV

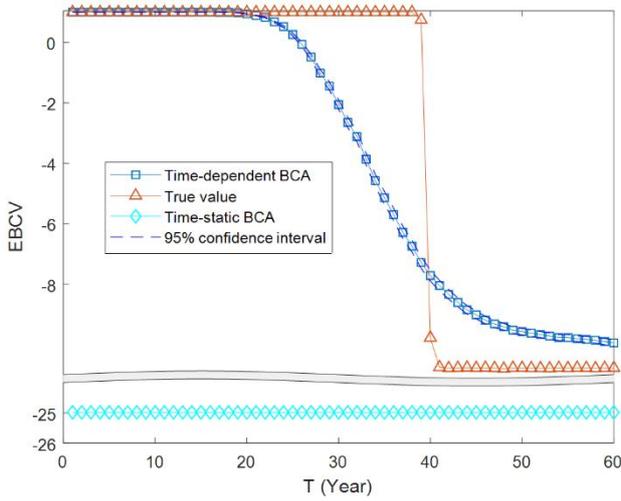


(b) P_{BF}

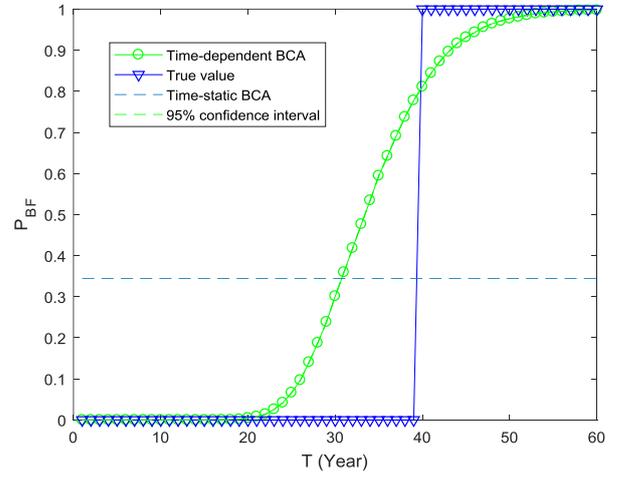


(c) P_{BI}

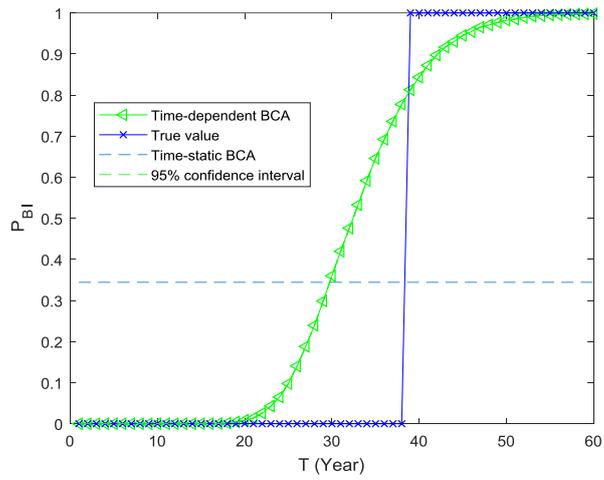
Figure 7. Business continuity metrics at $t=1$ year.



(a) EBCV



(b) P_{BF}



(c) P_{BI}

Figure 8. Business continuity metrics at $t=10$ years.

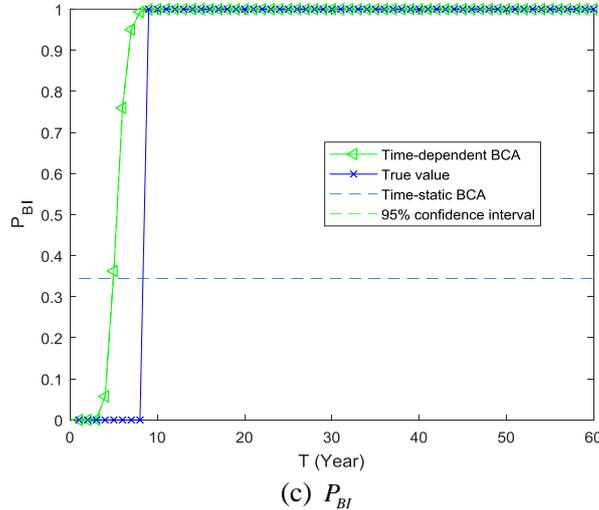
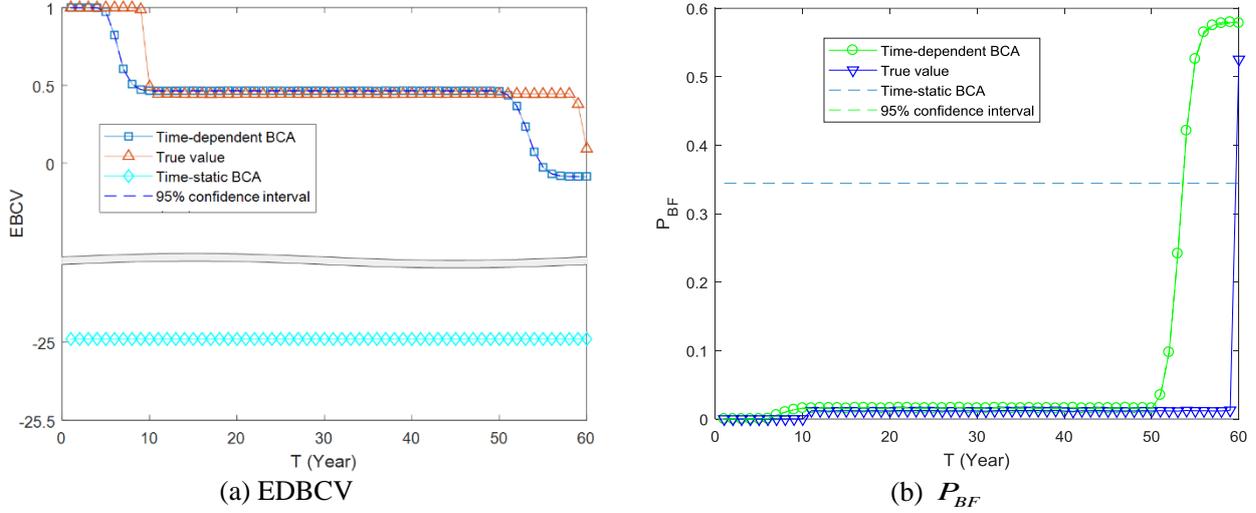


Figure 9. Business continuity metrics at $t=40$ years.

- At each t , with the increase of the estimation horizon T , the DBCV decreases. This means that regardless of the age t of the NPP, the longer the NPP is operated, the worse its business continuity: this is logical, as it is primarily caused by the tube's degradation process. No rupture is supposed to occur at the beginning of system operation. Subsequently, as the crack grows, rupture will occur eventually and lead to system failure. In addition, the dynamic business continuity (DBC) indexes curves drop (Figure 7 (a), Figure 8 (a), Figure 9 (a)) or rise (Figure 7 (b, c), Figure 8 (b, c), Figure 9 (b, c)) significantly after a certain value of T . In practice, intervention measures like overhauls need to be taken before this T , in order to prevent serious losses from occurring failures and ensure the business continuity.

- 2) For the same estimation horizon T , as the NPP age t increases, the EDBC V shifts left, which means that the financial safety margin is shrinking with t . This is because the steam generator tube is getting closer to a dangerous state with age.
- 3) When T is beyond a certain value, the business continuity metrics becomes invariant. This is mainly because when T is sufficiently long, the rupture event will surely happen and after that no loss occurs any more.
- 4) There are plateau sections in the curves of EBC V (Figure 7 (a), Figure 8(a), Figure 9 (a)); the height of these plateaus increases with time t , which makes sense because the system potential profits increase over time t .
- 5) The results comparison between DBCA and time-static BCA shows that the time-static BCA grossly underestimates the damage of SGTR on system business and, thus, underestimates the NPP's business loss. Moreover, the results from the DBCA using condition-monitoring data are closer to the true BC V than those of the time-static BCA. This is because the DBCA using condition monitoring data incorporates the time-dependent behavior of SGTR degradation.
- 6) the confidence intervals quantitatively express the level of confidence that the BC V metrics values are contained in the interval. From Figures 7~9, we can see that with more data available, the widths of the confidence intervals reduce. This is because with more condition monitoring data, the component state estimation becomes more accurate and the uncertainty in the BCA results reduces.

5 Discussion

The method developed in this work is applied on a case study regarding NPP operation, but it can also be applied to a wide variety of other scenarios. For systems with the following characteristics: (1) business continuity is related to financial losses; (2) system behavior and/or profit are potentially time-dependent; (3) condition monitoring data are available to inform on the time-dependent system behavior. For instance, in the example of oil storage tanks in [4], the profit of the oil storage tank depends on the price of the oil and is, therefore, time-dependent; lithium batteries are used to drive some critical safety barriers and are subject to degradation, so that the performance of the safety barriers is also time-dependent. Besides, condition monitoring data are available from the mounted sensors and can be used for online updating the failure probability of the safety barriers. For IT services, the profits also exhibit time-

dependent behaviors, the failure behavior of the hardware in the IT infrastructure is also time-dependent due to various degradation mechanisms, and if condition monitoring data are available to monitor the state of the hardware, the developed DBCA method can be applied.

Compared to the original time-static BCA method, the developed model captures the time-dependent features of both profits and system failure behaviors. Therefore, the proposed method can more precisely quantify the business continuity that exhibits time-dependent behaviors. However, the price one needs to pay is that the model is more complex in both development and analysis. In practice, there is the need to choose the most appropriate method based on a tradeoff between the complexity of the modelling and the accuracy of the results, and this depends on the characteristics of the problem and on the knowledge, information and data available for its description [64]. For example, for systems whose failure behavior is not time-dependent or not significant for business continuity, the traditional time-static BCA method might be sufficient. However, for safety critical systems that have significant time-dependency, the developed method is preferred due to its potential to provide a more accurate assessment.

It should be noted that in this work we assume that the operation costs (including the inspection and maintenance costs) do not change with time (as seen in Equation (15)). This assumption is reasonable for NPPs, because they are usually designed with sufficient margins so that even when they reach the design life, their performance is not degraded severely. However, these costs might be time-dependent, and typically increasing with time for other systems: this should be considered in the modelling, then.

Moreover, to illustrate the proposed DBCA model, we use a stochastic electricity model to predict the electricity price, considering a variety of factors contributing to electricity price variations (such as seasonal volatility, time-varying mean reversion and seasonally occurring price spikes). The predicted electricity price is shown in Figure 5. It should be noted that the predicted values are here used to illustrate the developed method only. There are various factors that have a potential influence on the electricity price (such as new energy source and new consumption patterns), which make the predicted results inevitably subject to uncertainty, especially in a long-time span of prediction. Therefore, when the developed method is applied in practice, up-to-date electricity information should be used, instead of the predicted value, in order to reduce the uncertainty and assessment errors.

It is noteworthy that this work considers as disruptive events only those that are caused by safety-related hazards. In practice, however, the problem of business continuity might arise for disruptive events generated by hazards other

than safety-related ones, e.g., natural hazards: the method developed can be extended to capture also these disruptive events.

6. Conclusions

In this paper, a DBCA method that integrates condition monitoring data is proposed. Two factors that influence the dynamic behavior of business continuity are considered explicitly. The first one is the dynamics of the degradation-to-failure process affecting the safety barriers. Condition monitoring data are used to update and predict the time-dependent failure behavior by PF. The second factor is the time-dependent profit and tolerable losses. This is quantified by applying a stochastic price model and an installment model. A simulation-based framework is developed to calculate the time-dependent business continuity metrics originally introduced. A case study regarding the analysis of an accident initiated by SGTR in a NPP shows that the proposed framework allows capturing the dynamic character of business continuity.

The outcomes of such dynamic analysis can provide insights to stakeholders and decision-makers, that can help them to identify when best to take actions for preventing serious losses and ensuring business continuity.

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Paper III

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Joint optimization of safety barriers against steam generator tube rupture to enhance business continuity of nuclear power plants

Jinduo Xing¹, Zhiguo Zeng¹, Enrico Zio^{2,3,4}

¹ Chair System Science and the Energy Challenge, Laboratoire Génie Industriel (LGI), CentraleSupélec, Université Paris Saclay, Gif-sur-Yvette, France

² MINES ParisTech, PSL Research University, CRC, Sophia Antipolis, France

³ Energy Department, Politecnico di Milano, Milan, Italy

⁴ Eminent Scholar, Department of Nuclear Engineering, College of Engineering, Kyung Hee University, Republic of Korea

jinduo.xing@centralesupelec.fr, zhiguo.zeng@centralesupelec.fr, enrico.zio@polimi.it

Abstract

In nuclear power plants (NPPs), different types of safety barriers are designed to ensure the safe and continuous operation of the NPP against disruptive events. These safety barriers, although designed to operate in different phases of the accidents evolution, are often optimized separately, without considering their collective effects on preventing disruptions and quickly recovering from the disruptions. This paper develops a joint optimization model for synthetically optimizing safety barriers of different natures, including prevention, mitigation, emergency and recovery barriers to enhance the business continuity of the NPP, considering the threat of steam generator tube rupture (SGTR) accidents. The joint optimization is guided by a business continuity metric called expected business continuity value (EBCV). A physics-of-failure model is developed to describe the crack growth process of the steam generator tube and to model the effect of the prevention barriers, i.e., periodical inspection of the crack length. An event tree model is developed to describe the evolution of the SGTR-initiated accident and to model the effect of the mitigation and emergency barriers. Recovery measures are also considered via a widely-used logarithmic function model. A mixed-integer genetic algorithm (MIGA) is used to obtain optimal solutions of the joint optimization model. The results show that the developed joint optimization model can achieve better performance in terms of business continuity, compared to the conventional methods that optimize the safety barriers separately.

Keywords

Business continuity management (BCM), Safety barrier, Joint optimization, Event tree (ET), Mixed-integer genetic algorithm (MIGA), Nuclear power plant (NPP), Steam generator tube rupture (SGTR).

Acronyms

BCM	business continuity management
EBCV	expected business continuity value
ET	event tree
MIGA	mixed-integer genetic algorithm
NPP	nuclear power plant
PDF	probability density function
PSA	probabilistic safety assessment
RDS	reactor depressurization system
RTS	reactor trip system
RWST	refueling water storage tank
SG	steam generator
SGTR	steam generator tube rupture

Notation

a	crack length
c_e	cost-effective parameter for II – th consequence
C	production capacity of NPP
C_i	consequence with i – th severity
C_p	cost in preventive phase
C_M	cost in mitigation phase
C_R	cost in recovery phase
C_{plug}	unit price for plugging one tube
C_{th}	total budget
$\frac{da}{dt}$	crack growth rate
ΔK	stress intensity factor
$L([0, T])$	potential loss in $[0, T]$

$L_d([0, T])$	direct loss in $[0, T]$
$L_{in}([0, T])$	indirect loss in $[0, T]$
L_{tol}	maximum tolerable loss
n_{tube}	number of tubes
PF	vector representing the baseline value for the failure probabilities of the mitigation measures
$P_{rup, tube}$	probability of one tube rupture
T	estimation horizon for business continuity assessment
T_{recv, C_i}	recovery time for the i – th consequence
μ	relationship between cost-effectiveness parameters of C_{II} and C_{III}

1. Introduction

Steam generator (SG) is a passive heat-exchanging system that transfers heat from the primary loop to the secondary loop in a pressurized water reactor (PWR) to produce steam to drive the turbines [1]. One relevant safety issue in PWR is the rupture of SG tubes, known as steam generator tube rupture (SGTR), which can be an accident initiating event induced by a crack growth process in the SG tube [2]. SGTR accidents can cause severe consequences related to the leakage of radioactive materials [3]. The safe operation of nuclear power plants (NPPs) is ensured by a suite of safety barriers designed to prevent undesired events or accidents, and contain or mitigate their consequences when they occur [4, 5].

Prevention barriers are designed to work in the pre-accident or pre-failure phases and aim at reducing the probability of occurrence of accidents [6, 7]. In the case of SGTR accidents, one commonly adopted prevention barriers is to make periodical inspection and timely preventive plugging of the defective tubes [8]. Mitigation and emergency barriers intervene after the accident initiating event occurs and aim at containing the evolution of the accident so that its consequences can be minimized. Examples of mitigation and emergency barriers for SGTR accidents include the reactor trip system (RTS), reactor depressurization system (RDS), refueling water storage tank (RWST), reactor cooling system (RCS), etc [9]. Recovery barriers aim at restoring to the normal operation system functionality timely after the accident [10]. For example, in the event of a SGTR, the recovery measures could include replacing the ruptured tube, cleaning up the contaminated area (if any), etc [11, 12].

In practice, the different safety barriers are designed separately for optimal performance, considering the constraints of limited resources [13]. For instance, [14] has proposed the optimal design of risk-based inspections in power and process plants. In [15], an enhanced preventive maintenance optimization model based on a three-stage failure process has been proposed for NPP components. In [16], condition-based maintenance optimization for deteriorating systems has been investigated by considering inspection intervals and preventive maintenance thresholds as decision variables. Only prevention activities were taken into account in these works. There are also a number of works considering the optimal design of emergency and mitigation barriers. For example, a simplified probabilistic safety assessment (PSA) and a reliability allocation model have been developed to improve the safety level of PWR by the optimal allocation of redundancies of the emergency and mitigation barriers [17]. In [18, 19], a redundancy allocation model for series-parallel systems has been used to improve the reliability of mitigation measures. In [20], an optimization problem is formulated and solved for the minimum average total cost of nuclear

fail-safe systems, where the average total system cost was subject to a restricted type I design error. The optimization of recovery measures has also been considered in the literature. For example, in [21], economically optimal strategies for recovering from a NPP accident have been discussed. In [22], joint restoration processes for multiple systems have been modeled and the effectiveness of five different restoration strategies with respect to resilience has been compared. A resilience-based optimization methodology has been proposed in [23].

Most existing research works, as reviewed above, optimize the safety barriers separately. In practical problems, however, the safety barriers at different phases need to work jointly to ensure that the NPP can be operated continuously and safely. In this paper, we propose a joint optimization model to ensure that the different safety barriers can achieve holistic optimal performances. The performance of the safety barriers system is quantified through the concept of business continuity. Defined as “the holistic management process that identifies potential threats to an organization and the impacts to business operations those threats, if realized, might cause, and which provides a framework for building organizational resilience with the capability of an effective response that safeguards the interest of its key stakeholders, reputation, brand and value-creating activities” [24], business continuity management (BCM) has received more and more attention in recent years, as a holistic requirement on the overall performance of a system [25, 26]. A loss-based business continuity metric was defined in [5] for business continuity assessment of an oil tank farm. In this work, we adopt the quantitative business continuity metric in [5] to guide the joint optimization of the pre- and post-disruption barriers, and use a mixed-integer genetic algorithm (MIGA) [27] to derive the optimal solution. The contributions of this work can be summarized as follows:

- (1) A joint optimization model is proposed to enhance the business continuity of NPP.
- (2) MIGA is implemented to solve the joint optimization problem.

The remaining of the paper is structured as follows. Section 2 formally defines the problem. Section 3 illustrates the SGTR event and the corresponding safety barriers at different stages. Section 4 provides a joint optimization model and a solution method for enhancing business continuity of the NPP. Section 5 performs a sensitivity analysis of the parameters in the four stages. Finally, Section 6 concludes this work.

2 Problem description

One SG of a PWR is considered, which is equipped with a bundle of 3592 inverted U tubes. Each U tube has an outside diameter Gaussian distributed with a mean of 22.23 mm and a standard deviation 0.1667mm; and a thickness Gaussian distributed with a mean of 1.27 mm and a standard deviation of 0.0592 mm. The tubes are subject to

different degradation mechanisms like stress corrosion cracking (SCC), fatigue, pitting corrosion and fretting wear.

A list of the NPP parameter values is presented in Table 1 [8].

Table 1 Parameters of the NPP.

Parameter	Value
Generation capacity of NPP (C)	1100 Mwh
Number of tubes (n_{tube})	3592
Tube outer diameter (d)	$N(22.23, 0.1667)$ mm
Tube thickness (b)	$N(1.27, 0.0592)$ mm

Different safety barriers are presented for preventing SGTR, containing its consequences and recovering from the possible disruptions caused. Table 2 summarizes the safety measures considered in this paper and highlights the category they belong to: prevention, mitigation, emergency and recovery.

Table 2 Safety barriers considered in this study.

Safety barrier	Category	Function description
Periodic inspection and maintenance	Prevention	Periodically inspect the tubes and timely plug those defective tubes whose crack length is beyond the maintenance threshold.
Reactor trip system (RTS)	Mitigation	When the reactor power exceeds a given safety operating limit, the RTS automatically shut down the reactor, in order to prevent core damage.
Reactor depressurization system (RDS)	Mitigation	When a loss-of-coolant event is caused by the tube rupture, the RTS will work to prevent over-pressurization of the reactor vessel.
Refueling water storage tank (RWST)	Mitigation	Store cooling water for emergency cooling of the reactor core.
Reactor cooling system (RCS)	Mitigation	If the reactor fails to scram, RCS will pump water into the reactor for emergency cooling.
Repair of the damages caused by the SGTR	Recovery	Replace the ruptured tubes and restore the plant to normal operation.

3 Modelling the individual safety barriers

In this section, we present the model of the performance of each safety barrier and the associated costs. Prevention safety barriers are discussed in Section 3.1, followed by mitigation and emergency barriers in Section 3.2, and recovery measures in Section 3.3.

3.1 Modeling prevention safety barriers

SGTR is the break of one or more SG tubes, which can be caused by different degradation mechanisms, e.g., stress corrosion cracking (SCC), fatigue, pitting corrosion, fretting wear [28]. As reported in [3], a fraction of 60% ~ 80% of SGTR events is caused by SCC. For this reason, without loss of generality, in this work, the SGTR is considered to be due only to SCC. The main prevention safety measure is to inspect the tube periodically and plug the dangerous tubes when necessary.

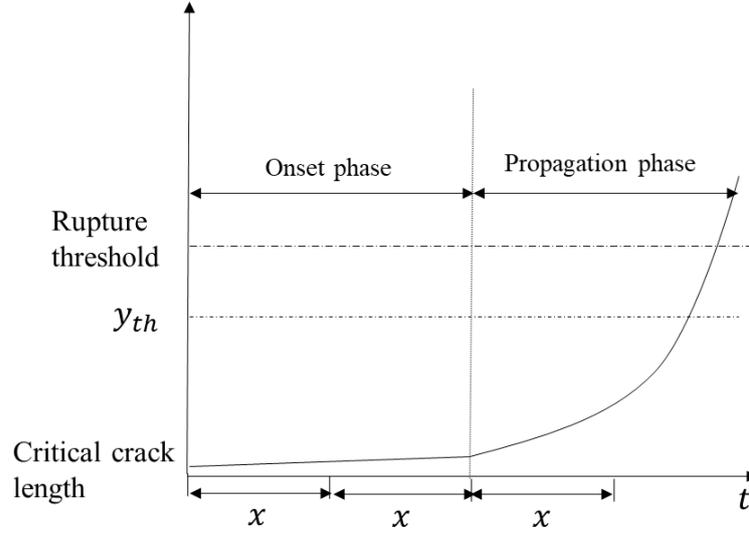


Figure 1 Tube crack growth process.

The tube SCC growth process can be divided into two stages: onset and propagation of cracks inside the tube wall [29], as shown in Figure 1. During the onset phase, the crack grows slowly; then, when the crack size reaches a critical point, it begins to grow more rapidly in the propagation phase. The SCC process is often modelled by a two-stage physics-based crack growth model. In the two-stage model, it is assumed that the critical crack length beyond which the process enters the propagation phase (see Figure 1) is 0.1 mm. The duration of the onset phase is assumed to follow a lognormal distribution with parameters μ_c and σ_c , where μ_c denotes the mean value and σ_c is the standard deviation [8]. Let $t_{critical}$ represent the time needed for the crack to reach the length of critical size 0.1 mm. Then, we have:

$$t_{critical} \sim \log normal(\mu_c, \sigma_c^2), \quad (1)$$

In this paper, we take the values of the parameters from [8]: $\mu_c = 9.3$ years, $\sigma_c = 3.162$ years.

The Scott model is often used to model the propagation phase of the crack growth process (see Figure 1), in which the crack growth rate is empirically modelled as a function of stress [30]:

$$\frac{da}{dt} = \alpha(K - K_{th})^m, \quad (2)$$

$$K = F\sigma\sqrt{\frac{\pi a}{2}}, \quad (3)$$

$$\sigma = \frac{\Delta P \cdot d}{2b} \quad (4)$$

where $\frac{da}{dt}$ is the crack growth rate, a is the crack length, α , K_{th} and m are constant parameters related to the component material properties, σ is the stress at the crack tip, ΔP denotes the pressure difference. In this paper, the material of steam generator tubes is assumed to be Alloy 600. Based on the material properties, the values for the parameters in Equations (2)-(3) can be determined, as shown in Table 3.

Table 3 Parameter values of crack growth model.

Parameter	Value
a_0	0.1
F	0.93
m	1.16
K_{th}	9
ΔP	$N(8.3,0.33)$ Mpa

Based on the physical model of the crack growth process, the effectiveness of the prevention safety measures can be further modeled. It is assumed that periodical inspections are conducted every x months and during each inspection, the crack size can be measured through techniques like eddy current testing [31], ultrasonic testing [32]), etc. If the measured crack length is beyond a given preventive maintenance threshold, denoted by y_{th} , the corresponding tube is plugged to prevent further damages. The probability of SGTR, denoted by $p_{rupt,tube}$, is used to represent the performance of the prevention safety barrier: the higher the value of $p_{rupt,tube}$, the worse the performance of the prevention barrier. In our model, the probability of SGTR is formulated as a function of x and y_{th} :

$$p_{rupt,tube} = f(x, y_{th}). \quad (5)$$

Monte Carlo simulation is used to evaluate the value of $p_{rupt,tube}$, as shown in Algorithm 1.

Algorithm 1: $p_{rupt,tube}$ calculation based on Monte Carlo simulation.

Inputs: $x, d, \Delta P, b, a_0, F, m, K_{th}, a_{rup,th}, T, N_s$;

Outputs: $p_{rupt,tube}$;

$t_{in} = [x, 2x, 3x, \dots, kx], k = 1, 2, \dots, n, nx \leq T$;

$Index = 0$;

```

For  $i = 1: N_s$ 
  Use equation (1) to generate  $t_{critical}(i)$ ;
  If  $t_{critical}(i) < T$ ;
     $t = t_{critical}$ ;
    Based on  $x, T, t_{critical}, a_{th}$ , calculate the rupture time  $t_{rup}(i)$ ;
    For  $k = 1: n$ 
      If  $t_{rup}(i) > T$ 
         $Index = Index$ ;
        else if  $a_{th} > a_{in}(k) > y_{th} > a_{in}(k-1)$ 
           $Index = Index$ ;
        else if  $a_{in}(k) > a_{th} > a_{in}(k-1)$ , and  $a_{in}(k-1) < y_{th}$ ;
           $Index = Index + 1$ ;
        End
      End
    End
  End
End
End

```

$p_{rupt,tube} = Index / N_s$.

where T is the calculation time horizon; a_{th} represents the rupture threshold; t_{rup} is the rupture time; a_{in} implies the crack size measured at inspection.

The cost spent in the prevention phase, represented by C_p , is a function of inspection interval x and plugging rate p_{plug} :

$$\begin{aligned}
 C_p &= C_{insp} \cdot n_{insp} + n_{tube} \cdot p_{plug} \cdot C_{plug} \\
 &= \frac{T}{x} \cdot C_{insp} + n_{tube} \cdot p_{plug} \cdot C_{plug},
 \end{aligned} \tag{6}$$

where C_{insp} denotes the cost of a single inspection of the tube; C_{plug} is the unit price for plugging one tube; and p_{plug} is the plugging rate of the tube, i.e., the fraction of tubes being plugged. The value of p_{plug} depends on the values of x and y_{th} , as shown in Algorithm 1.

3.2 Modeling emergency and mitigation barriers

As shown in Table 2, the emergency and mitigation barriers include the RTS, RDS, RWST and RCS. Their effects on containing the consequence of SGTR can be modelled using an event tree (ET), as shown in Figure 2.

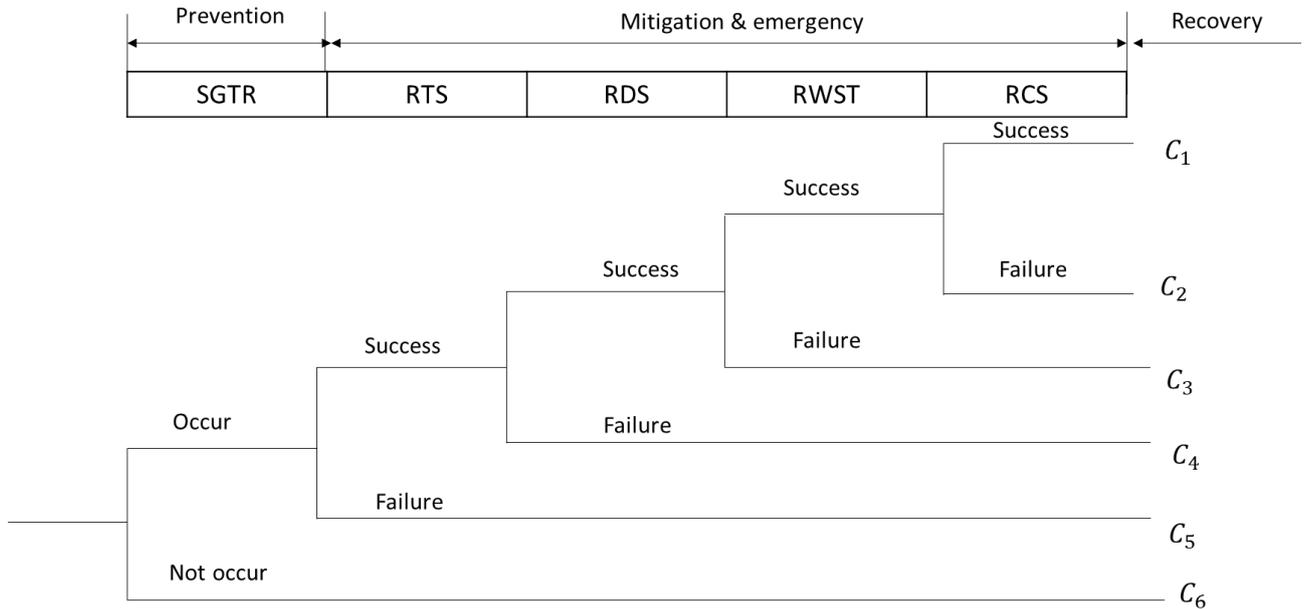


Figure 2 Schematic ET model of SGTR accident.

Depending on the performance of different barriers, 6 consequences i.e., C_1, C_2, \dots, C_6 can be caused by the SGTR. These consequences can be grouped into three categories in Table 5 based on their severity. In this regard, different consequences are caused by the variant performance of safety barrier. Based on the ET in Figure 2, the occurrence probabilities for $C_1 \sim C_6$ can be quantified as a function of the event probabilities. Conceptually, we denote these by:

$$p_{C_i} = f_{ET}(p_{SGTR}, p_1, p_2, p_3, p_4) \quad (6)$$

where p_{SGTR} is the occurrence probability of SGTR and can be calculated based on the models of the prevention safety barrier, p_1, p_2, p_3, p_4 represent the failure probability of RTS, RDS, RWST and RCS, respectively.

Table 4 Classification of consequences.

Consequences	Group	Meaning
C_6	C_I	No SGTR occurs, the NPP is operating normally.
C_1	C_{II}	SGTR occurs, but the consequence is successfully controlled by the mitigation and emergency barriers. The power generation business is temporarily terminated.

$C_2 \sim C_5$	C_{III}	Core damage is caused by SGTR, the power generation business is terminated for a long time.
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The performance of the mitigation & emergency barriers is primarily determined by their failure probabilities. One common way to reduce the failure probabilities of these barriers is to add redundancies. If one safety barrier collapses, the redundant system will substitute it. Redundancy allocation models are often used for designing the redundant safety systems within constraints on costs and resources [30, 31].

In this paper, we assume that parallel redundancy applying the same type of equipment is used for the four mitigation and emergency safety barriers. It is easy to show that the failure probability of the i -th measure is:

$$p_i = (p_{i,b})^{n_i+1}, \quad (7)$$

where $p_{i,b}$ is the failure probability of the i -th safety barrier system and n_i is the number of redundant system added to the original system. The total cost associated with the redundancy design is:

$$C_{SB_M} = \sum_{i=1}^4 C_{R,i} \cdot n_i, \quad (8)$$

where $C_{R,i}$ is the price for adding one i -th redundancy equipment.

3.3 Modeling recovery measures

Recovery ability refers to the ability of a system to be repaired and quickly restored its normal operation after failures or disruptions [32]. The repair ability depends on a variety of factors including the training and preparedness of repair groups, the readiness of repair materials and resources, etc. A range of models have been proposed in the literature for the post-disruptive event recovery process [33-35]. According to these models, the performance of the recovery process directly depends on the resources spent on the recovery processes, e.g., the investment on training repair crews, preparing equipment and spare parts used for repairing the failed items [36]. In general, the more budgets or resources planned for recovery process, the better recovery performances.

In this work, we use the following model for the recovery process [37]:

$$T_{recv,i} = \frac{T_{recv,C_i}}{1 + \ln(1 + c_{e,i} \cdot \mu \cdot C_{S_{BR}})} \quad (9)$$

where $T_{recv,i}$ is a random variable that represents the time needed to recover from the i -th consequence; T_{recv,C_i} denotes the basic recovery time for consequence C_i , which is dependent on the basic requirement on recovery time, C_{SB_R} denotes the resources invested on the recovery process and $c_{e,i}$ is the effective parameters of resources on i -th consequence; μ denotes the relationship between different cost-effectiveness parameters for different severity consequence; Its value should be set by decision makers based on the capability of the organization.

As most literatures (e.g., [5, 38, 39]) applied, we also assume that T_{recv,C_i} follows a lognormal distribution:

$$f(T_{recv,C_i}) = \begin{cases} \frac{1}{\sqrt{2\pi}\beta_i T_{recv,C_i}} e^{-\frac{(\ln(T_{recv,C_i})-\varepsilon_i)^2}{2\beta_i^2}}, & T_{recv,C_i} > 0 \\ 0, & T_{recv,C_i} \leq 0 \end{cases} \quad (10)$$

where $f(\cdot)$ is the probability density function (PDF) of T_{recv,C_i} ; ε_i and β_i are the mean value and variance value of the lognormal distribution, respectively. The values of ε_i and β_i are depends on the recovery ability of the target organization. It is noting that the more serious consequence i , the smaller of $c_{e,i}$.

4 Joint optimization model based on business continuity

In this Section, we develop a joint optimization model to ensure global optimal performances of the safety barriers (prevention, mitigation, emergency and recovery). The joint optimization model is based on the objective of maximizing the business continuity of the plant. In Section 4.1, we start from a review of the business continuity model and numerical metrics we used to guide the optimization. The joint optimization model is presented in Section 4.2. Section 4.3 discusses how to solve this joint optimization by using a MIGA. The results of the application on the case study of Section 2 is presented and discussed in Section 4.4.

4.1 Basics of business continuity modeling and assessment

As explained in Section 1, business continuity management (BCM) is a comprehensive method that integrates pre-event and post-event management together to ensure the resilience and continuous operation of system business. Compared to conventional risk analysis method, BCM not only focuses on the potential hazards and their impacts, but also considers how to mitigate the consequence and quickly recover from the disruption. A quantitative index, i.e., expected business continuity value (EBCV) has been developed in [5] for business continuity modeling and assessment:

$$EBCV = 1 - \frac{E[L([0,T])]}{L_{tol}} \quad (11)$$

where T is the evaluation time horizon for the business continuity assessment; $L([0,T])$ is a random variable that describes potential losses in $[0,T]$ caused by the disruptive event; L_{tol} denotes the maximum tolerable losses that an organization can tolerate: beyond that level of loss, it will have difficulty to recover the corresponding business. As can be seen from this definition, EBCV measures expected system financial risk level. A higher EBCV indicates higher business continuity.

Two kinds of losses are considered when calculating $L([0,T])$: direct loss $L_d([0,T])$ and indirect loss $L_{in}([0,T])$. The former one represents the losses that are caused directly by the disruptive event, including structural damage of the system. The latter is the revenue loss suffered during the shutdown of the plant, quantified by equipment damage and the other direct loss. For example, in the case study of SGTR in Section 2, an example of $L_d([0,T])$ is the direct financial due to damages caused to the assets. $L_{in}([0,T])$ might be the downtime costs of the NPP due to the maintenance and recovery process. The total loss is calculated as:

$$L([0,T]) = L_{in}([0,T]) + L_d([0,T]) \quad (12)$$

The direct losses $L_d([0,T])$ are mainly determined by the performance of prevention, mitigation and emergency measures, while the indirect losses $L_{in}([0,T])$ are more related to the performance of the recovery process. Therefore, the EBCV can be viewed as a global performance measures that integrates the performance of prevention, mitigation, emergency and recovery measures.

In Figure 3, we describe a general process for business continuity modeling and assessment. The first step is to identify the potential disruptive events. Because different disruptive events might lead to different losses and, then result in different business continuity. Subsequently, we analyze the performance of the safety barriers and develop models to support their evaluation. Then, the potential losses caused by the disruptive events should be modelled and estimated through models like ET and semi-Markov process [5, 40]. Finally, the value of business continuity metrics can be calculated based on the estimated loss.

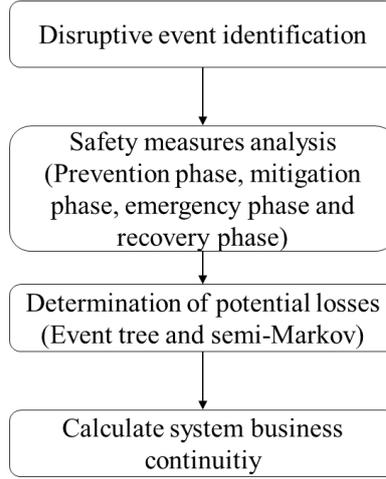


Figure 3 A general procedures for business continuity modelling and assessment.

4.2 The joint optimization model

In this section, we present a joint optimization model of the safety barriers in different phases, with an objective of maximizing the EBCV [9]. As shown in Equation (11), the EBCV is determined by the direct loss L_d and indirect losses L_{in} caused by the SGTR. The different possible consequences can be modelled by the ET model in Figure 2. Based on their severity, these consequences were grouped into three categories in Table 5. It is assumed that for each category, the direct losses caused by SGTR are equipment damage, such as steam generator tube, the whole NPP, etc.

If the consequences C_{II} and C_{III} happen, the NPP will become temporality unavailable for producing electricity, until the repair crew successfully handle the incident/accident and restore the normal operation of NPP. The indirect losses caused in the recovery process can, then, be modelled by:

$$L_{in,C_i} = P_e \cdot C \cdot T_{recv,i}, \quad (13)$$

where L_{in,C_i} represents the indirect losses in the recovery process for consequence $C_i, i = I, II, III$ ($L_{in,C_i} = 0$ as there is no disruption for this event sequence); P_e is the unit electricity price; C is the generating capacity of the NPP; $T_{recv,i}$ denotes the recovery time for the i -th consequence and is calculated by Equation (9). Then, the *EBCV* can be formulated by:

$$\begin{aligned}
EBCV &= \frac{\sum_{i=1}^3 E(L_{C_i}([0, T])) \cdot p_{C_i}}{L_{tol}} \\
&= \frac{\sum_{i=1}^3 E(L_{d,C_i} + L_{in,C_i}) \cdot p_{C_i}}{L_{tol}} \\
&= \left[p_{C_{II}} \cdot (L_{d,C_{II}} + P_e \cdot C \cdot E(T_{recv,II})) + p_{C_{III}} \cdot (L_{d,C_{III}} + P_e \cdot C \cdot E(T_{recv,III})) \right] / L_{tol}.
\end{aligned} \tag{14}$$

In Equation (14), $p_{C_{II}}$ and $p_{C_{III}}$ represent the occurrence probabilities of consequences II and III, respectively. These probabilities are calculated using the ET model in Figure 2, where the event probabilities in the ET are further determined based on the model developed in Section 3.1 and Section 3.2. The distributions of $T_{recv,II}$ and $T_{recv,III}$ are determined based on Equations (9) and (10). As shown in Equation (14), the EBCV can be conceptually as a function of $x, y_{th}, n_1, n_2, n_3, n_4, C_{SB_R}$, whose meanings are listed in the notation list.

Notation list:

x periodic inspection time;

y_{th} preventive maintenance threshold;

n_1, n_2, n_3, n_4 the redundancy of RTS, RDS, RWST, RCS, respectively;

C_{SB_R} The resource/budget allocated on recovery phase.

Hence, a joint optimization model can be set up, with maximizing EBCV as objective function:

$$\max EBCV = f(x, y_{th}, n_1, n_2, n_3, n_4, C_{SB_R}) \tag{15}$$

$$s.t. C_{SB_p} + C_{SB_M} + C_{SB_R} \leq C_{Th}, \tag{16}$$

$$C_{SB_p}, C_{SB_M}, C_{SB_R} \geq 0, \tag{17}$$

$$x \in [6, 7, 8, \dots, 17, 18], \tag{18}$$

$$P_{plug} \leq P_{th}, \tag{19}$$

$$\begin{aligned}
n_{L,1} &\leq n_1 \leq n_{U,1}, n_1 \in N \\
n_{L,2} &\leq n_2 \leq n_{U,2}, n_2 \in N \\
n_{L,3} &\leq n_3 \leq n_{U,3}, n_3 \in N \\
n_{L,4} &\leq n_4 \leq n_{U,4}, n_4 \in N.
\end{aligned} \tag{20}$$

The objective function is only represented conceptually here. In practice, Algorithm 1 is often used to evaluate the EBCV through Monte Carlo simulation. The first constraint in Equation (16) regards the total budget on all safety measures: the total costs on the different safety measures cannot exceed a limited value C_{th} . In Equation (16), the costs $C_{SB_P}, C_{SB_M}, C_{SB_R}$ are further calculated by Equations (5), (8) and (9). The constraint in Equation (18) defines the possible value of inspection intervals (in years). In this work, it is assumed that the inspection can reveal the exact degradation state of the tube. The constraint in Equation (19) means that the total number of plugged tubes cannot exceed a maximum value. The value of p_{th} is determined based on the power generation efficiency requirement of the NPP. According to the nuclear regulations, a steam generator of the type employed in Zion PWR NPP can tolerate up to 30% plugged tubes before a significant reduction in efficiency occurs [41]. Therefore, here, we see $p_{th} = 0.3$. The last constraint in Equation (20) describes the minimal and maximal number of redundant system for the mitigation measures. The parameter values in this paper are tabulated in Table 6.

Table 6 Parameter values used in the case study.

Parameter	Meaning	Value	Source
L_{tot}	Tolerable losses	5×10^6 (k€)	Assumed
C_{insp}	Cost for one inspection	500 (k€)	Assumed
C_1	Cost for adding one redundant RTS (Equation (8))	20 (k€)	Assumed
C_2	Cost for adding one redundant RDS	4 (k€)	Assumed
C_4	Cost for adding one redundant RWST	11 (k€)	Assumed
C_4	Cost for adding one redundant RCS	5 (k€)	Assumed
C_{plug}	Cost for plugging one tube	5 (k€)	[42]
$c_{II,e}$	Cost effectiveness parameter for consequence C_{II}	0.001	Assumed
μ	The relationship of cost-effectiveness parameter between C_{II} and C_{III}	0.5	Assumed
$T_{recv,1}$	Basic recovery time for consequence C_B	Lognormal (3.9828, 0.4724) (days)	Assumed
$T_{recv,2}$	Basic recovery time for consequence C_C	Lognormal (6.5922, 0.4724) (days)	Assumed

P_e	Unit price for electricity	50 (k€/MWh)	[43]
$n_{L,1}, n_{L,2}, n_{L,3}, n_{L,4}$	Lower bound of mitigation measures' number	0	Assumed
$n_{U,1}, n_{U,2}, n_{U,3}, n_{U,4}$	Upper bound of mitigation measures' number	4	Assumed
C_{th}	Total budget (Equation (16))	8000 (k€)	Assumed

4.3 Mixed integer genetic algorithm

The joint optimization model in Equation (17) is a mixed integer programming problem, as some decision variables (n_1, n_2, n_3, n_4) are restricted to take integer values whereas the others can take also non-integer values. There are a lot of methods for solving the mixed-integer programming problem, e.g., branch and bound technique [45], Lagrange multiplier [46]. In this paper, we choose the MIGA to solve the joint optimization model for its powerful capability to handle highly complex, nonlinear numerical models and its successful application in related areas like optimization of critical infrastructure resilience [22, 40]. A flowchart of implementing the MIGA is shown in Figure 3. In this paper, we use the MIGA toolbox in MATLAB 2017b to solve this joint optimization model.

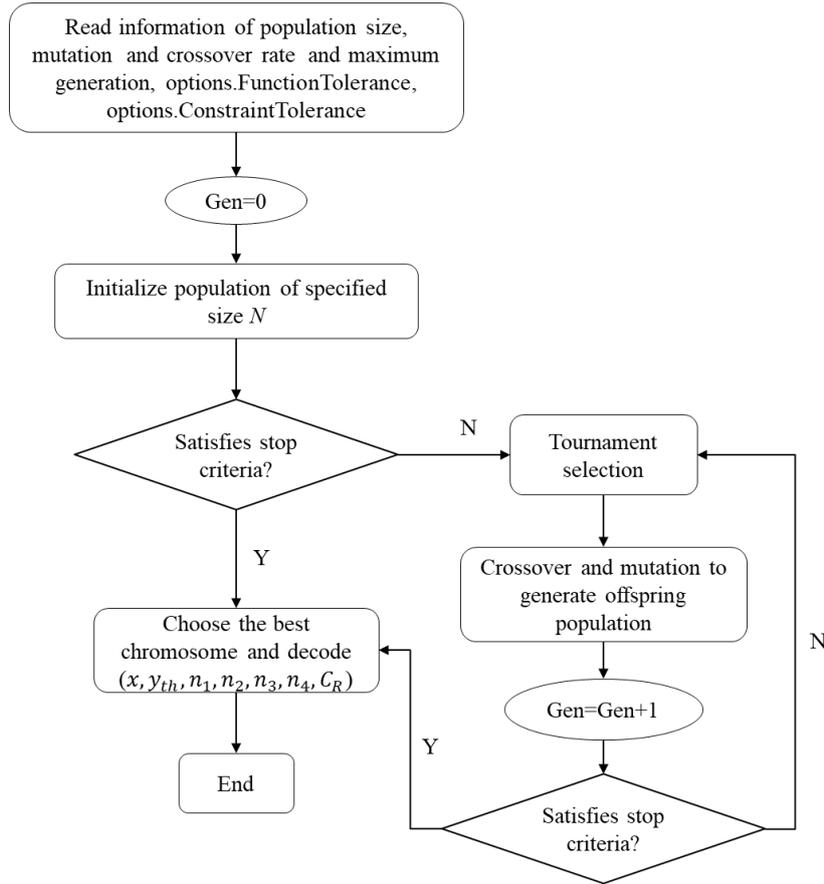


Figure 3 Flowchart of MIGA.

The parameter values of MIGA used in this work are shown in Table 7.

Table 7 Parameter values of MIGA.

Parameter	Values
Population size	40
Crossover rate	0.9
Mutation rate	0.5
Maximum generation	300

4.4 Results and discussions

The optimization model in Equation (17) was solved numerically using the MIGA described in Section 4.3. The MIGA was run 10 times since the MIGA tends to converge to local minimum. The optimal EBCV value for each run is shown in Figure 4. As a comparison, the EBCV values from the individual optimization models are obtained by Equation (17), assuming that all the budget is invested on prevention safety barriers. It can be seen that, in general, the joint optimization works better in terms of achieving higher EBCV than the individual optimization model. More specifically, the joint optimization method can reduce the potential total losses and achieve a higher business

continuity against SGTR events than only investing all the budget on the prevention phase. The same conclusions are revealed considering the cost of investing only in the mitigation and emergency phase or in the recovery phase.

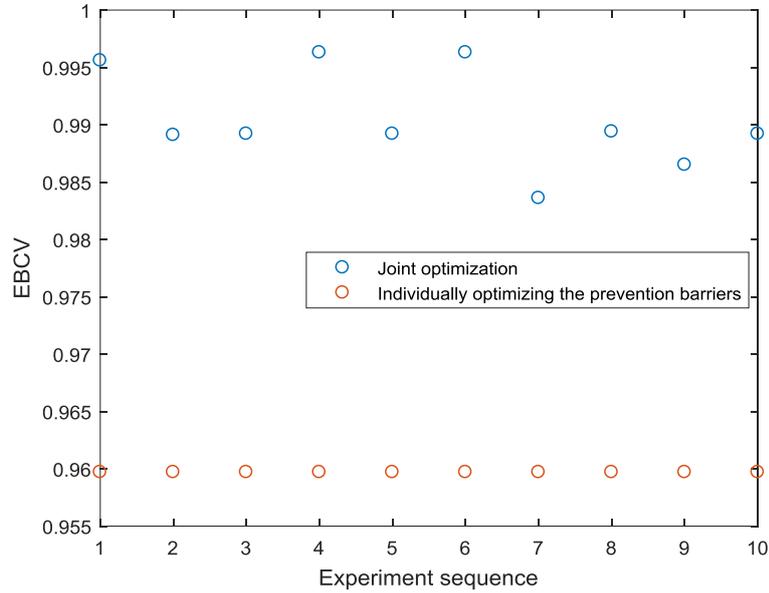


Figure 4 Comparison between the proposed joint optimization and individually optimizing the prevention barriers (

$$C_{total} = 8000\text{k€}).$$

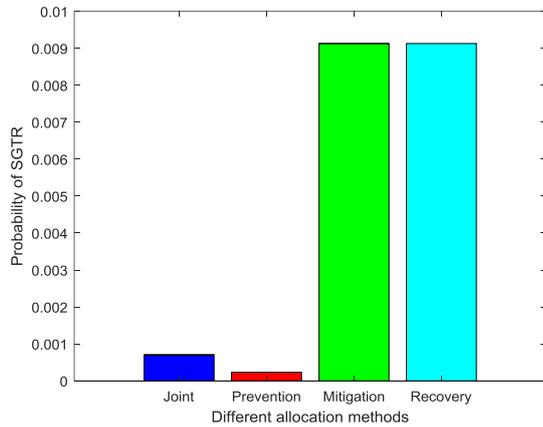
To further examine the differences between the different optimization strategies, we compare the optimal values of the decision variables from different optimization models in Table 8. Note that for the joint optimization model, the best solution among the 10 times is selected as the optimal solution to the model: it requires a periodical inspection of the steam generation tube every 15 months and a tube should be plugged whenever its crack length exceed 7.8780 mm. The number of redundant components in mitigation and emergency barriers are $n_1 = 1, n_2 = 4, n_3 = 4, n_4 = 4$, where 1~4 correspond to the redundant components in the RTS, RDS, RWST, RCS, respectively. Another $C_R = 108.2384$ (k€) is allocated for improving the performance of the recovery process.

Table 8 Comparison results for business continuity under different strategies.

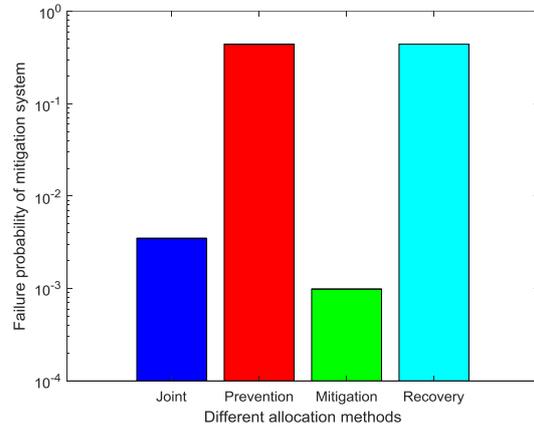
Variable	Joint optimization	Prevention measures only	Mitigation and emergency	Recovery (only)
x	1.25 (year)	1 (year)	/	/
y_{th}	7.8780 (mm)	14.1347(mm)	/	/
n_1	1	/	4	0

n_2	4	/	4	0
n_3	4	/	4	0
n_4	4	/	4	0
C_P	7792.2 (k€)	7990.05(k€)	0	0
C_M	90 (k€)	0	120 (k€)	0
C_R	108.2384 (k€)	0	0	8000 (k€)
C_{th}	7990.4 (k€)	7999.05 (k€)	120 (k€)	8000 (k€)
EBCV	0.9963	0.9597	-20.06	-44.63

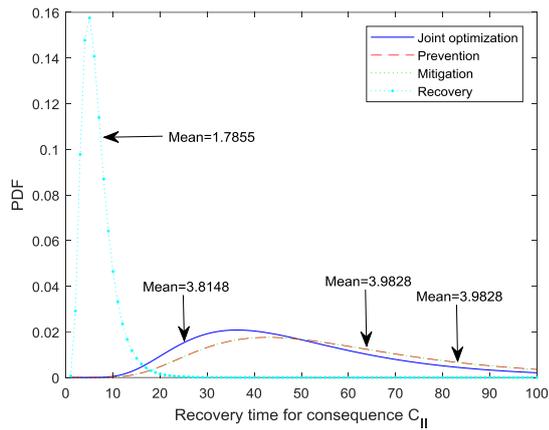
In Figure 5. we compare the performances of the different safety barriers under different optimization model. In particular, the probability of SGTR is used to represent the performance of prevention safety barriers: a higher value of SGTR probability indicates worse performance of the prevention barriers. The failure probability of mitigation indicates the mitigation system performance, where the higher the failure probability the worse the performance of the mitigation system. Moreover, the mean time to restore normal operation and the PDF of the recovery times show the recovery ability, where longer recovery times indicate poorer recovery ability.



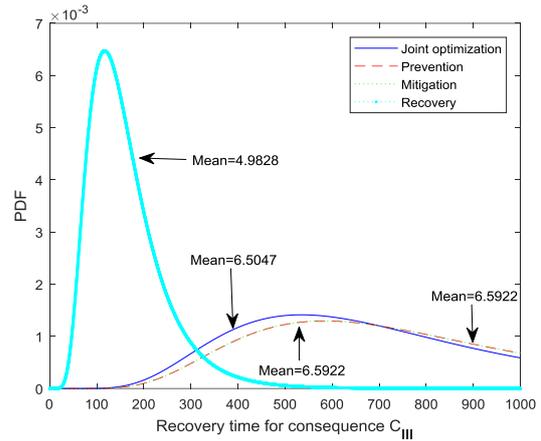
(a) Probability of SGTR.



(b) Failure probability of mitigation system.



(c) PDF of recovery time for consequence C_{II} .



(d) PDF of recovery time for consequence C_{III} .

Figure 5 Performance indexes in prevention, mitigation & emergency, and recovery phases.

It can be seen from Figure 5(a) that the SGTR occurrence probability takes the lowest value 6.9998×10^{-4} when all the budget is spent on the preventive measures. However, the corresponding EBCV is lower than the joint optimization model (as shown in Table 8), indicating that the joint optimization model can achieve a better performance globally. This is because considering the preventive measures individually might sometimes affect the performance of the mitigation and emergency and recovery processes, as less resources can be invested on these measures. Similar results can be found in Figure 5 (b), (c) and (d): although the solution obtained from the joint optimization model is not optimal with respect to each safety barrier, it can achieve an overall optimal performance with respect to business continuity.

5. Sensitivity analysis

A sensitivity analysis is conducted to investigate how does the optimal design solution changes with total budget, basic failure probability of mitigation barriers ($p_{i,b}$ in Equation (9)) and cost-effectiveness parameter (c_e in Equation (11)) of improving recovery barriers. The sensitivity analysis is gradually done by changing the parameter of interest while fixing the values of the other parameters.

Figure 6 shows how does the optimal EBCV change under different values of total budget C_{th} . As expected, as the total budget grows, the optimal EBCV correspondingly increases. The error bar in Figure 6 shows the mean and standard deviation of EBCV using the proposed joint optimization model for 10 experiments. The main reason for this is that as the C_{th} increases, more available resources can be used to enhance system business continuity.

However, it should be noted that when C_{th} exceeds 8000k€, the marginal effects on EBCV become very small. This result shows that although increasing C_{th} can improve the business continuity, increasing the budget further when it already reaches a threshold value (8000k€ in this case) might become a waste of resources, as it cannot further improve the business continuity. In practice, the most cost-benefit way of setting C_{th} is to set the budget around this threshold value. A dramatic change of the EBCV value can also be observed in Figure 6. This is mainly because when the total budget is larger than 7888.2k€, the occurrence probability of SGTR will dramatically decreased.

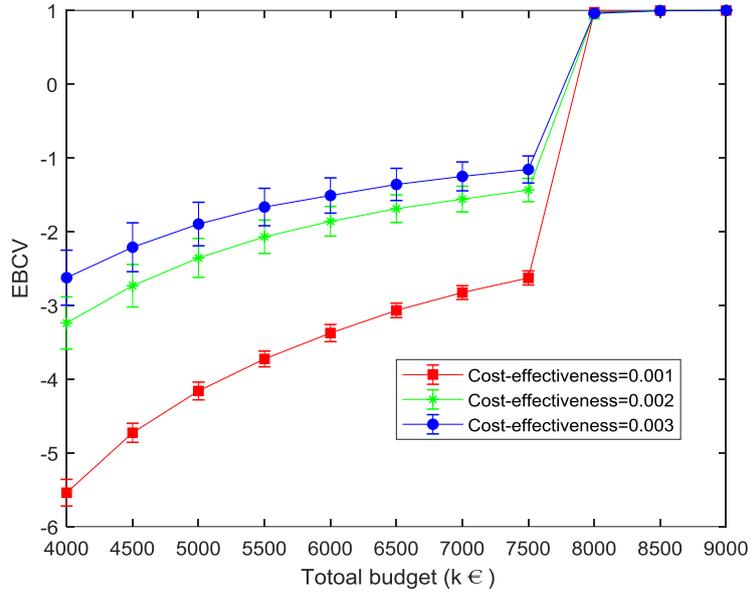


Figure 6 Sensitivity of the optimal EBCV against the C_{th} .

To investigate the influence of the cost-effective parameter c_e on the optimal solution, the optimal values of EBCV under different values of c_e are shown in Figure 7. Figure 7 shows the joint optimization results where the C_{th} is 7500k€ and 8000k€, respectively. As can be seen from Figure 7, the optimal EBCV value is very sensitive to the increase of c_e when the total budget is small (7500k€), whereas when the C_{th} is large enough, the optimal EBCV value almost does not change with the changing of c_e . This is because the partial differentiation of $T_{recv,i}$ with respect to C_R is:

$$\frac{\partial T_{recv,i}}{\partial C_R} = \frac{-c_e}{c_e \cdot C_R + 1} \quad (23)$$

We can see that a smaller value of C_R leads to a larger value of $\frac{\partial T_{recv,i}}{\partial C_R}$: therefore, the smaller the total budget

(corresponding to the smaller C_R), the more sensitive is EBCV to the value of c_e .

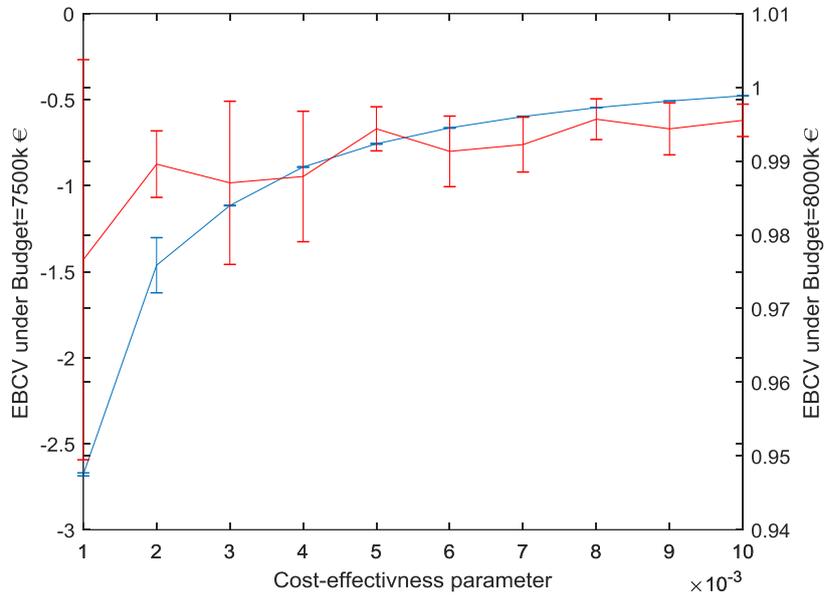


Figure 7 Comparison EBCV with different cost-budgets.

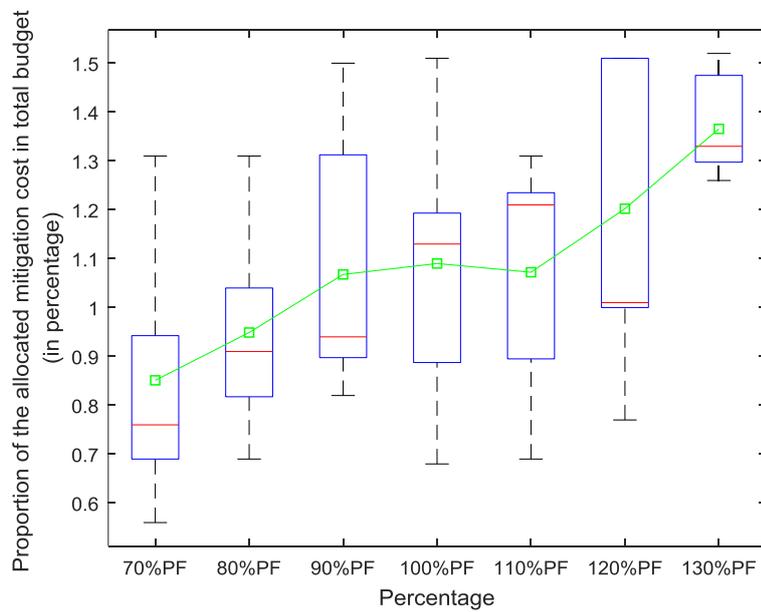


Figure 8 Schematic of changing failure probability of mitigation measures (70%-130%) under budget

$$C_{total} = 8000k\text{€}.$$

To investigate the influence of the basic failure probabilities of mitigation barriers on the budget allocated to improve them, we do a sensitivity analysis by changing the mitigation failure probability from 70% to 130% of the baseline values used in Table 2 for the failure probability of the mitigation barriers. As shown in Figure 8 (where the black lines are the minimum and maximum values among the different runs of the MIGA, the red line denote the median values, the blue box shows the upper 75% quartile and lower 25% quartile, respectively, when the failure probabilities of the mitigation barriers increase, the proportion of the total budget that is allocated for improving the mitigation barriers increases accordingly, This is a straightforward conclusion: more redundant safety systems are needed if the mitigation safety system has higher failure probability.

6. Conclusions

In this work, a joint optimization model is formulated to optimally allocate the limited resource among the different safety barriers to enhance business continuity of NPPs against SGTR. The model allows integrated design of prevention measure, mitigation, emergency and recovery measure. The prevention measures are modeled by a periodical inspection model based on a physics-of-failure model. The mitigation and emergency measures are modelled through a parallel redundancy model. The recovery measures are described by a logarithm recovery function. An application on an NPP demonstrates the feasibility of the developed model and can provide a globally optimal solution than optimizing the safety barriers separately. Through sensitivity analysis, we can infer that: (1) larger resource budget can result in higher business continuity; the change of EBCV is marginally decreasing with the increase of the budget; there is an optimal budget for the given NPP; (2) higher failure probability of the safety measures in mitigation phase, less redundancy is needed; (3) the smaller the budget, the more sensitive of cost-effectiveness parameters on business continuity. The optimization method can jointly provide a better scheme than separative optimization strategies for decision makers under limited budget or resources.

This work focuses on initial trail on joint optimization based on business continuity. Our further efforts will be devoted to utilizing the joint optimization approach proposed in this research by considering dependency of safety barriers, using advanced method, such as Bayesian network.

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Titre : Un cadre quantitatif pour l'évaluation et l'optimisation dynamique de la continuité d'activité des systèmes énergétique

Mots clés: Evaluation dynamique des risques, Evaluation dynamique de la continuité des opérations, Données de surveillance des conditions

Résumé: La gestion de la continuité des opérations est un cadre complet visant à éviter que les événements perturbateurs n'affectent les opérations commerciales, à rétablir rapidement les activités et à réduire les dommages potentiels correspondants pour les systèmes énergétiques, tels que les centrales nucléaires. Cette thèse propose des discussions sur les aspects suivants: développement de méthodes appropriées d'évaluation des risques afin d'intégrer les données de surveillance de l'état et les données d'inspection pour une mise à jour et des pronostics robustes et en temps réel du profil de risque. Pour tenir compte de l'incertitude des données de surveillance de l'état, un modèle de mélange gaussien de Markov caché est développé pour modéliser les données de surveillance de l'état. Un réseau bayésien est appliqué pour intégrer les deux sources de données. Pour améliorer l'applicabilité de la continuité des opérations dans la pratique, les variables variant dans le temps considèrent l'indice de continuité des opérations, par ex. la dégradation des composants, les revenus en fonction du temps, etc. sont pris en compte dans le processus de modélisation de la continuité des activités. Sur la base de l'indice de continuité d'activité proposé, une méthode d'optimisation conjointe prenant en compte toutes les mesures de sécurité dans le processus d'évolution des événements, y compris les étapes de prévention, d'atténuation, d'urgence et de récupération, est développée pour améliorer la continuité des opérations du système avec des ressources limitées. Les méthodologies proposées sont appliquées aux centrales nucléaires contre les événements perturbateurs.

Title: Business continuity of energy systems: a quantitative framework for dynamic assessment and optimization

Keywords: Dynamic risk assessment, Dynamic business continuity assessment, Condition monitoring data

Business continuity management is a comprehensive framework to prevent the disruptive events from impacting the business operations, quickly recovering business and reducing the corresponding potential damages for energy system, such as nuclear power plants (NPPs). This dissertation provides discussions on the following aspects: developing appropriate risk assessment methods in order to integrate condition monitoring data and inspection data for a robust and real-time risk profile updating and prognostics. To account for the uncertainty of condition monitoring data, a hidden Markov gaussian mixture model is developed to model the condition monitoring data. A Bayesian network is applied to integrate the two data sources. For improving applicability of business continuity in practice, time-variant variables regard business continuity index, e.g. component degradation, time-dependent revenue, etc are taken into consideration in the business continuity modelling process. Based on the proposed business continuity index, a joint optimization method considering all the safety measures in event evolvment process including prevention stage, mitigation stage, emergency stage and recovery stage is developed to enhance system business continuity under limited resources. The proposed methodologies are applied to NPP against disruptive event.